

APPENDIX TO A NONPARAMETRIC R^2 TEST FOR THE PRESENCE OF RELEVANT VARIABLES

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Appendix

Below we list three lemmas that are used frequently and then we provide the proof of Theorems 1-5 and sketch of proof of Remarks 1 and 2. Throughout the proof, c will represent an inconsequential and arbitrary constant that may take different values in different contexts.

Lemma 1 For second order U -statistic $U_n = \frac{1}{n^2} \sum_{t=1}^n \sum_{i=1}^n \phi_n(Z_t, Z_i)$, where $\phi_n(Z_t, Z_i)$ is a symmetric function of Z_t and Z_i that could depend on n , and $\{Z_t\}_{t=1}^n$ is a sequence of IID random variables. Define $\hat{U}_n = \frac{1}{n} \sum_{t=1}^n \int \phi_n(Z_t, Z_i) dP(Z_i) - \frac{1}{2} E\phi_n(Z_t, Z_i)$, where $P(Z_i)$ is the distribution function of Z_i . Then we have $U_n - \hat{U}_n = O_p(n^{-1}(E\phi_n^2(Z_t, Z_i))^{\frac{1}{2}})$.

Note: from Lemma 1 we obtain the following convenient expression.

$$\begin{aligned} \frac{1}{n^2} \sum_{t=1}^n \sum_{i=1}^n \phi_n(Z_t, Z_i) &= \frac{1}{n^2} \sum_{t=1}^n \sum_{\substack{i=1 \\ t \neq i}}^n \phi_n(Z_t, Z_i) + \frac{1}{n^2} \sum_{t=1}^n \sum_{\substack{i=1 \\ t < i}}^n \phi_n(Z_t, Z_i) = \frac{2}{n^2} \sum_{t=1}^n \sum_{\substack{i=1 \\ t < i}}^n \phi_n(Z_t, Z_i) \\ &= \frac{2}{n} \sum_{t=1}^n \int \phi_n(Z_t, Z_i) dP(Z_i) - E\phi_n(Z_t, Z_i) + O_p(n^{-1}(E\phi_n^2(Z_t, Z_i))^{\frac{1}{2}}) \end{aligned}$$

Lemma 1 is similar in spirit to Lemma 8.4 in Newey and McFadden (1994). Thus the proof is omitted.

Lemma 2 Define $S_{n,j}(z_0) = \frac{1}{n} \sum_{i=1}^n K_h(Z_i - z_0) \left(\frac{Z_i - z_0}{h}\right)^j g(U_i) w(Z_i - z_0; z_0)$, $|j| = 0, \dots, J$, where Z_i, U_i are iid, $Z_i \in \mathbb{R}^{l_c}$, $K_h(\cdot) = \frac{1}{h^{l_c}} K(\frac{\cdot}{h})$, and $K(\cdot)$ is a kernel function defined on $\mathbb{R}^{l_c}$. Here $j = (j_1, j_2, \dots, j_{l_c})$, $Z_i = (Z_{i,1}, Z_{i,2}, \dots, Z_{i,l_c})$, $z_0 = (z_{0,1}, z_{0,2}, \dots, z_{0,l_c})$, and $\left(\frac{Z_i - z_0}{h}\right)^j = \left(\frac{Z_{i,1} - z_{0,1}}{h}\right)^{j_1} \times \dots \times \left(\frac{Z_{i,l_c} - z_{0,l_c}}{h}\right)^{j_{l_c}}$. Assume L_1 . $K(\cdot)$ is bounded with compact support and for Euclidean norm $\|\cdot\|$,

$$|u^j K(u) - v^j K(v)| \leq c_K \|u - v\|, \text{ for } 0 \leq |j| \leq J.$$

L_2 . $g(U_i)$ is a measurable function of U_i and $E|g(U_i)|^s < \infty$ for $s > 2$.

L_3 . For G a compact subset of $\mathbb{R}^{l_c}$, define the joint density of Z_i and U_i at (z_0, u) as $f(z_0, u)$, conditional density of Z_i and U_i given U_i at $Z_i = z_0$ and $U_i = u$ as $f_{z|u}(z_0)$. Assume

$\sup_{z_0 \in G} \int |g(u)|^s f_{z,u}(z_0, u) du < \infty$, $f_{z|u}(z_0) < \infty$, and $f_{z,u}(z_0, u)$ is continuous around z_0 .

L_4 . $w(Z_i - z_0; z_0)$ is a function of $Z_i - z_0$ and z_0 . $|w(Z_i - z_0; z_0)| < c < \infty$, $|w(Z_i - z_0; z_0) - w(Z_i - z_k; z_k)| \leq c \|z_0 - z_k\|$ almost everywhere.

L_5 . $\frac{nh^{l_c}}{\ln(n)} \rightarrow \infty$.

$$\text{Then for } z_0 \in G, \sup_{z_0 \in G} |S_{n,j}(z_0) - E(S_{n,j}(z_0))| = O_{a.s.} \left(\left(\frac{nh^{l_c}}{\ln(n)} \right)^{-\frac{1}{2}} \right).$$

Lemma 2 is similar in spirit to Lemma 1 in Yao and Zhang (2010). Thus the proof is omitted.

Comment: If $ES_{n,j}(z) = 0$, then L_3 could be replaced with

$L3'$: For G a compact subset of $\mathbb{R}^{l_c}$, define the joint density of Z_i and U_i at (z_0, u) as $f(z_0, u)$, conditional density of Z_i and U_i given U_i at $Z_i = z_0$ and $U_i = u$ as $f_{z|u}(z_0)$. Assume $f_{z|u}(z_0) < \infty$.

Lemma 3 For $\{Z_t\}_{t=1}^n$ an IID sequence of random variables, we define a third order U -statistic by

$$U_n = \binom{n}{3}^{-1} \sum_{(n,3)} \psi_n(Z_{i_1}, Z_{i_2}, Z_{i_3}),$$

where the sum $\sum_{(n,3)}$ is taken over all subsets $1 \leq i_1 < i_2 < i_3 \leq n$ of $\{1, 2, \dots, n\}$ and the function $\psi_n(\cdot)$

is symmetric in Z_{i_1}, Z_{i_2} , and Z_{i_3} . Define the conditional expectations $\psi_{1n}(z_1) = E(\psi_n(Z_1, Z_2, Z_3) | Z_1 = z_1)$, $\psi_{2n}(z_1, z_2) = E(\psi_n(Z_1, Z_2, Z_3) | Z_1 = z_1, Z_2 = z_2)$ and $\psi_{3n}(z_1, z_2, z_3) = \psi_n(z_1, z_2, z_3)$. $\sigma_{1n}^2 = \text{Var}(\psi_{1n}(Z_1))$, $\sigma_{2n}^2 = \text{Var}(\psi_{2n}(Z_1, Z_2))$, $\sigma_{3n}^2 = \text{Var}(\psi_{3n}(Z_1, Z_2, Z_3))$. Let $\theta_n = E\psi_n(Z_1, Z_2, Z_3)$. Then we have the following H -decomposition

$$\begin{aligned}
U_n &= \theta_n + \sum_{j=1}^3 \binom{3}{j} H_n^{(j)} \\
&= \theta_n + \binom{n}{3}^{-1} \binom{n-1}{2} \left[\sum_{t=1}^n \psi_{1n}(Z_t) - n\theta_n \right] \\
&\quad + \binom{n}{3}^{-1} \binom{n-2}{1} \left[\sum_{\substack{t=1 \\ t < i}}^n \sum_{i=1}^n \psi_{2n}(Z_t, Z_i) - (n-1) \sum_{t=1}^n \psi_{1n}(Z_t) + \frac{n(n-1)}{2} \theta_n \right] \\
&\quad + \binom{n}{3}^{-1} \left\{ \sum_{(n,3)} \psi_n(Z_{i_1}, Z_{i_2}, Z_{i_3}) - \frac{(n-1)(n-2)}{2} \sum_{t=1}^n [\psi_{1n}(Z_t) - \theta_n] \right. \\
&\quad \left. - (n-2) \sum_{\substack{t=1 \\ t < i}}^n \sum_{i=1}^n [\psi_{2n}(Z_t, Z_i) - \psi_{1n}(Z_t) - \psi_{1n}(Z_i) + \theta_n] - \binom{n}{3} \theta_n \right\}
\end{aligned}$$

Furthermore, we have $V(H_n^{(3)}) = \binom{n}{3}^{-1} \sum_{c=1}^3 (-1)^{3-c} \binom{3}{c} \sigma_{cn}^2 = O(n^{-3}(\sigma_{1n}^2 + \sigma_{2n}^2 + \sigma_{3n}^2)) = O(n^{-3}\sigma_{3n}^2)$.
 $V(H_n^{(2)}) = O(n^{-2}\sigma_{2n}^2)$ and $V(H_n^{(1)}) = O(n^{-1}\sigma_{1n}^2)$.

Lemma 3: Proof. The lemma follows from Theorem 1-3 in section 1.6 of Lee (1990).

We find it easy to provide proof of Theorem 3 first as we base the proof of other Theorems on it.

Theorem 3: Proof.

Under the alternative hypothesis, $m(x) \neq \mu$. From the definition in equation (2), we define

$$\tilde{R}^2 = 1 - \frac{\frac{1}{n} \sum_{t=1}^n (y_t - \hat{m}(x_t))^2}{\frac{1}{n} \sum_{t=1}^n (y_t - \bar{y})^2}, \text{ so } \hat{R}^2 = \tilde{R}^2 I\left(\frac{1}{n} \sum_{t=1}^n (y_t - \bar{y})^2 \geq \frac{1}{n} \sum_{t=1}^n (y_t - \hat{m}(x_t))^2\right).$$

Since $\frac{\hat{a}}{\hat{b}} - \frac{a}{b} = \frac{1}{b} \left[(\hat{a} - a) - (\hat{b} - b) \frac{a}{b} \right]$, we have

$$\begin{aligned}
\tilde{R}^2 - R^2 &= -\frac{\frac{1}{n} \sum_{t=1}^n (y_t - \hat{m}(x_t))^2}{\frac{1}{n} \sum_{t=1}^n (y_t - \bar{y})^2} + \frac{E(y_t - m(x_t))^2}{V(y_t)} \\
&= -\frac{1}{\frac{1}{n} \sum_{t=1}^n (y_t - \bar{y})^2} \left[\underbrace{\frac{1}{n} \sum_{t=1}^n (y_t - \hat{m}(x_t))^2 - E(y_t - m(x_t))^2}_{B_1} - \underbrace{\left(\frac{1}{n} \sum_{t=1}^n (y_t - \bar{y})^2 - V(y_t) \right) \frac{E(y_t - m(x_t))^2}{V(y_t)}}_{B_2} \right].
\end{aligned}$$

(1) We show $B_2 \frac{E(y_t - m(x_t))^2}{V(y_t)} = \frac{1}{n} \sum_{t=1}^n \left\{ \frac{E\sigma^2(x_t)}{V(y_t)} (y_t^2 - Ey_t^2) - 2(y_t - Em(x_t)) \frac{E\sigma^2(x_t)}{V(y_t)} Em(x_t) + O_p(n^{-1}) \right\}$.

$$\begin{aligned}
\text{Note } B_2 \frac{E(y_t - m(x_t))^2}{V(y_t)} &= \left(\frac{1}{n} \sum_{t=1}^n (y_t - \bar{y})^2 - V(y_t) \right) \frac{E\sigma^2(x_t)}{V(y_t)} = \left(\frac{1}{n} \sum_{t=1}^n y_t^2 - \bar{y}^2 - V(y_t) \right) \frac{E\sigma^2(x_t)}{V(y_t)} \\
&= \frac{1}{n} \sum_{t=1}^n \left\{ (y_t^2 - Ey_t^2) \frac{E\sigma^2(x_t)}{V(y_t)} - (\bar{y}^2 - (Ey_t)^2) \frac{E\sigma^2(x_t)}{V(y_t)} \right\}.
\end{aligned}$$

So if $\bar{y}^2 = \frac{2}{n} \sum_{t=1}^n [y_t(Em(x_t)) - \frac{1}{2}(Em(x_t))^2] + O_p(n^{-1})$ then

$$\bar{y}^2 - (Ey_t)^2 = \frac{2}{n} \sum_{t=1}^n y_t Em(x_t) - 2(Em(x_t))^2 + O_p(n^{-1}) = \frac{2Em(x_t)}{n} \sum_{t=1}^n (y_t - Em(x_t)) + O_p(n^{-1}), \text{ then}$$

we have the claim in (1).

So we only need to show $\bar{y}^2 = \frac{2}{n} \sum_{t=1}^n [y_t Em(x_t) - \frac{1}{2}(Em(x_t))^2] + O_p(n^{-1})$.

$$\bar{y}^2 = \left[\frac{1}{n} \sum_{t=1}^n y_t \right]^2 = \frac{1}{n^2} \sum_{t=1}^n y_t^2 + \frac{1}{n^2} \sum_{t=1}^n \sum_{\substack{i=1 \\ t \neq i}}^n \underbrace{y_t y_i}_{\phi(y_t, y_i)} = B_{21} + B_{22}.$$

First, we note $B_{21} = O_p(n^{-1})$ since $\frac{1}{n} \sum_{t=1}^n y_t^2 \xrightarrow{P} E y_t^2 < \infty$ by A2.

Second, we note in B_{22} , $\phi(y_t, y_i)$ is symmetric. We apply Lemma 1 on B_{22} . $E \phi^2(y_t, y_i) = (E y_t^2)^2 < \infty$ since $t \neq i$ and by A2. $\int \phi(y_t, y_i) dP(y_i) = y_t E(y_i)$, so

$$\begin{aligned} B_{22} &= \frac{2}{n} \sum_{t=1}^n \int \phi(y_t, y_i) dP(y_i) - E \phi(y_t, y_i) + O_p(n^{-1}) = \frac{2}{n} \sum_{t=1}^n y_t E(y_i) - (E y_t)^2 + O_p(n^{-1}) \\ &= \frac{2}{n} \sum_{t=1}^n [y_t (E m(x_t)) - \frac{1}{2} (E m(x_t))^2] + O_p(n^{-1}). \end{aligned}$$

From B_{21} and B_{22} we have the claim that $\bar{y}^2 = \frac{2}{n} \sum_{t=1}^n [y_t E m(x_t) - \frac{1}{2} (E m(x_t))^2] + O_p(n^{-1})$.

$$(2) \frac{1}{n} \sum_{t=1}^n (y_t - \hat{m}(x_t))^2 = \frac{1}{n} \sum_{t=1}^n [\epsilon_t^2 + 2(m(x_t) - \hat{m}(x_t))\epsilon_t + (m(x_t) - \hat{m}(x_t))^2] = I_1 + I_2 + I_3.$$

Note: $\hat{m}(x_t) - m(x_t) = \frac{1}{n h_n f(x_t)} \sum_{i=1}^n K(\frac{x_i - x_t}{h_n}) \epsilon_i + \frac{1}{2 n h_n f(x_t)} \sum_{i=1}^n K(\frac{x_i - x_t}{h_n}) m^{(2)}(x_{it})(x_i - x_t)^2 + w_n(x_t)$, where

$x_{it} = \lambda x_t + (1 - \lambda)x_i$ for some $\lambda \in (0, 1)$, and $w_n(x_t) = \hat{m}(x_t) - m(x_t) - \frac{1}{n h_n f(x_t)} \sum_{i=1}^n K(\frac{x_i - x_t}{h_n}) y_{i*}$, where $y_{i*} = y_i - m(x_t) - m^{(1)}(x_t)(x_i - x_t) = \epsilon_i + \frac{1}{2} m^{(2)}(x_{it})(x_i - x_t)^2$.

(i) We show that $w_n(x_t) = O_p(R_{n,2}(x_t))$ uniformly for $x_t \in G$, where

$$R_{n,2}(x_t) = \left| \frac{1}{n} \sum_{i=1}^n K(\frac{x_i - x_t}{h_n}) y_{i*} \right| + \left| \frac{1}{n} \sum_{i=1}^n K(\frac{x_i - x_t}{h_n}) (\frac{x_i - x_t}{h_n}) y_{i*} \right|.$$

Define $S_n(x_t) = \begin{pmatrix} S_{0n}(x_t) & S_{1n}(x_t) \\ S_{1n}(x_t) & S_{2n}(x_t) \end{pmatrix}$ with $S_{jn}(x_t) = \frac{1}{n h_n} \sum_{i=1}^n K(\frac{x_i - x_t}{h_n}) (\frac{x_i - x_t}{h_n})^j$, and

$$S(x_t) = \begin{pmatrix} f(x_t) & 0 \\ 0 & f(x_t) \sigma_K^2 \end{pmatrix} = \begin{pmatrix} S_0(x_t) & S_1(x_t) \\ S_1(x_t) & S_2(x_t) \end{pmatrix}.$$

Then $\hat{m}(x_t) - m(x_t) = \frac{1}{n h_n} \sum_{i=1}^n W_n(\frac{x_i - x_t}{h_n}, x_t) y_{i*}$, where $W_n(z, x) = (1, 0) S_n^{-1}(x) (1, z)' K(z)$. So

$$\begin{aligned} |w_n(x_t)| &= \frac{1}{n h_n} \left| \sum_{i=1}^n (W_n(\frac{x_i - x_t}{h_n}, x_t) - \frac{1}{f(x_t)} K(\frac{x_i - x_t}{h_n})) y_{i*} \right| \\ &= \frac{1}{n h_n} \left| (1, 0) (S_n^{-1}(x_t) - S^{-1}(x_t)) \begin{pmatrix} \sum_{t=1}^n K(\frac{x_i - x_t}{h_n}) y_{i*} \\ \sum_{t=1}^n K(\frac{x_i - x_t}{h_n}) (\frac{x_i - x_t}{h_n}) y_{i*} \end{pmatrix} \right| \\ &\leq \frac{1}{h_n} ((1, 0) (S_n^{-1}(x_t) - S^{-1}(x_t))^2 (1, 0)')^{\frac{1}{2}} \frac{1}{n} \left(\left| \sum_{t=1}^n K(\frac{x_i - x_t}{h_n}) y_{i*} \right| + \left| \sum_{t=1}^n K(\frac{x_i - x_t}{h_n}) (\frac{x_i - x_t}{h_n}) y_{i*} \right| \right). \end{aligned}$$

We use Lemma 2 with A4, and follow Lemma 2 of Martins-Filho and Yao (2007) to obtain $w_n(x_t) = O_{a.s.}(R_{n,2}(x_t))$ uniformly for $x_t \in G$. With assumption A6-A8, we apply Lemma 2 to obtain

$$\begin{aligned} \frac{1}{n} \sum_{i=1}^n K(\frac{x_i - x_t}{h_n}) y_{i*} &= \frac{1}{n} \sum_{i=1}^n K(\frac{x_i - x_t}{h_n}) \epsilon_i + \frac{1}{2n} \sum_{i=1}^n K(\frac{x_i - x_t}{h_n}) (x_i - x_t)^2 m^{(2)}(x_{it}) = O_{a.s.}(h_n (\frac{n h_n}{\ln n})^{-\frac{1}{2}}) + \\ &O_{a.s.}(h_n^3) \text{ uniformly in } x_t \in G. \text{ In a similar fashion we obtain } \frac{1}{n} \sum_{i=1}^n K(\frac{x_i - x_t}{h_n}) (\frac{x_i - x_t}{h_n}) y_{i*} = O_{a.s.}(h_n (\frac{n h_n}{\ln n})^{-\frac{1}{2}}) + \\ &O_{a.s.}(h_n^3) \text{ uniformly in } x_t \in G. \text{ Thus, } \sup_{x_t \in G} |w_n(x_t)| = O_{a.s.}(h_n (\frac{n h_n}{\ln n})^{-\frac{1}{2}}) + O_{a.s.}(h_n^3). \end{aligned}$$

(ii) Since $f(x_t) \geq \underline{B}_f$ in A4(1), results in (i) give us

$$\sup_{x_t \in G} |\hat{m}(x_t) - m(x_t)| = O_{a.s.}((\frac{n h_n}{\ln n})^{-\frac{1}{2}}) + O_{a.s.}(h_n^2).$$

$$\sup_{x_t \in G} |\hat{m}(x_t) - m(x_t)|^2 = O_{a.s.}((\frac{n h_n}{\ln n})^{-1}) + O_{a.s.}(h_n^4) + O_{a.s.}(h_n^2 (\frac{n h_n}{\ln n})^{-\frac{1}{2}}).$$

So $\sqrt{n} \sup_{x_t \in G} |\hat{m}(x_t) - m(x_t)|^2 = O_{a.s.}(\frac{\ln n}{\sqrt{n} h_n^2}) + O_{a.s.}(\sqrt{n} h_n^8) + O_{a.s.}((n h_n^8)^{\frac{1}{4}} (\frac{\ln n}{\sqrt{n} h_n^2})^{\frac{1}{2}}) = o_{a.s.}(1)$ by A6.

We conclude $I_3 = o_p(n^{-\frac{1}{2}})$.

$$\begin{aligned}
\text{(iii)} \quad I_2 &= -\frac{2}{n} \sum_{t=1}^n (\hat{m}(x_t) - m(x_t)) \epsilon_t \\
&= -\frac{2}{n^2 h_n} \sum_{t=1}^n \sum_{i=1}^n \frac{1}{f(x_t)} K\left(\frac{x_i - x_t}{h_n}\right) \epsilon_i \epsilon_t - \frac{1}{n^2 h_n} \sum_{t=1}^n \sum_{i=1}^n \frac{1}{f(x_t)} K\left(\frac{x_i - x_t}{h_n}\right) (x_i - x_t)^2 m^{(2)}(x_{it}) \epsilon_t - \frac{2}{n} \sum_{t=1}^n w_n(x_t) \epsilon_t \\
&= I_{21} + I_{22} + I_{23} \\
I_{21} &= -\frac{1}{n^2} \sum_{t=1}^n \sum_{i=1}^n \underbrace{\left[\frac{1}{h_n f(x_t)} K\left(\frac{x_i - x_t}{h_n}\right) \epsilon_i \epsilon_t + \frac{1}{h_n f(x_i)} K\left(\frac{x_t - x_i}{h_n}\right) \epsilon_t \epsilon_i \right]}_{\phi_n(Z_t, Z_i) \text{ where } Z_t = (x_t, \epsilon_t)} \\
&= -\frac{2}{n^2 h_n} \sum_{t=1}^n \frac{K(0) \epsilon_t^2}{f(x_t)} - \frac{1}{n^2} \sum_{t=1}^n \sum_{\substack{i=1 \\ t \neq i}}^n \phi_n(Z_t, Z_i) = -I_{211} - I_{212}.
\end{aligned}$$

First note: $nh_n E|I_{211}| = 2EK(0) \frac{\sigma^2(x_t)}{f(x_t)}$, so $I_{211} = O_p((nh_n)^{-1}) = o_p(n^{-\frac{1}{2}})$ with A6. Second, we apply Lemma 1 to obtain

$$I_{212} - \left[\frac{2}{n} \sum_{t=1}^n \int \phi_n(Z_t, Z_i) dP(Z_i) - E\phi_n(Z_t, Z_i) \right] = O_p(n^{-1} (E\phi_n^2(Z_t, Z_i))^{\frac{1}{2}}).$$

We observe $\int \phi_n(Z_t, Z_i) dP(Z_i) = E\phi_n(Z_t, Z_i) = 0$ since $E(\epsilon_t | x_t) = 0$ in A3 and $E\phi_n^2(Z_t, Z_i) \leq c \left[\frac{1}{h_n^2} E \frac{1}{f^2(x_t)} K^2\left(\frac{x_i - x_t}{h_n}\right) \epsilon_i^2 \epsilon_t^2 + \frac{1}{h_n^2} E \frac{1}{f^2(x_i)} K^2\left(\frac{x_t - x_i}{h_n}\right) \epsilon_t^2 \epsilon_i^2 \right]$.

Since $\frac{1}{h_n} E \frac{1}{f^2(x_t)} K^2\left(\frac{x_i - x_t}{h_n}\right) \epsilon_i^2 \epsilon_t^2 = \int \frac{1}{f^2(x_t)} K^2(\psi) \sigma^2(x_t + h_n \psi) \sigma^2(x_t) f(x_t) f(x_t + h\psi) dx_t d\psi$
 $\rightarrow \int K^2(\psi) d\psi \int (\sigma^2(x_t))^2 dx_t$, we have $(E\phi_n^2(Z_t, Z_i))^{\frac{1}{2}} = O(h_n^{-\frac{1}{2}})$.
So we conclude $I_{212} = O_p((n^2 h_n)^{-\frac{1}{2}}) = o_p(n^{-\frac{1}{2}})$ and $I_{21} = o_p(n^{-\frac{1}{2}})$.

$$\begin{aligned}
I_{22} &= \frac{1}{n^2} \sum_{t=1}^n \sum_{\substack{i=1 \\ t \neq i}}^n \left[\frac{-1}{2h_n f(x_t)} K\left(\frac{x_i - x_t}{h_n}\right) (x_i - x_t)^2 m^{(2)}(x_{it}) \epsilon_t - \frac{1}{2h_n f(x_i)} K\left(\frac{x_t - x_i}{h_n}\right) (x_t - x_i)^2 m^{(2)}(x_{ti}) \epsilon_i \right] \\
&= \frac{1}{n^2} \sum_{t=1}^n \sum_{\substack{i=1 \\ t \neq i}}^n [\psi_n(Z_t, Z_i) + \psi_n(Z_i, Z_t)] = \frac{1}{n^2} \sum_{t=1}^n \sum_{\substack{i=1 \\ t \neq i}}^n \phi_n(Z_t, Z_i).
\end{aligned}$$

Since $\phi_n(Z_t, Z_i)$ is symmetric, we apply Lemma 1 again to obtain

$$I_{22} - \left[\frac{2}{n} \sum_{t=1}^n \int \phi_n(Z_t, Z_i) dP(Z_i) - E\phi_n(Z_t, Z_i) \right] = O_p(n^{-1} (E\phi_n^2(Z_t, Z_i))^{\frac{1}{2}}).$$

Note $E\phi_n(Z_t, Z_i) = 0$ by A3.

$$\frac{2}{n} \sum_{t=1}^n \int \phi_n(Z_t, Z_i) dP(Z_i) = \frac{-1}{n} \sum_{t=1}^n \frac{\epsilon_t}{h_n f(x_t)} E\left(K\left(\frac{x_i - x_t}{h_n}\right) (x_i - x_t)^2 m^{(2)}(x_{it}) | Z_t\right).$$

Since $\frac{1}{h_n} E\left(K\left(\frac{x_i - x_t}{h_n}\right) \left(\frac{x_i - x_t}{h_n}\right)^2 m^{(2)}(x_{it}) | Z_t\right) = \int K(\psi) \psi^2 m^{(2)}(x_t + \lambda h\psi) f(x_t + h\psi) d\psi$
 $\rightarrow \sigma_K^2 m^{(2)}(x_t) f(x_t)$ for $\lambda \in (0, 1)$ by A4 and A5,

$$E \frac{\sqrt{n}}{h_n^2} \frac{2}{n} \sum_{t=1}^n \int \phi_n(Z_t, Z_i) dP(Z_i) = 0, \text{ and}$$

$$V\left(\frac{\sqrt{n}}{h_n^2} \frac{2}{n} \sum_{t=1}^n \int \phi_n(Z_t, Z_i) dP(Z_i)\right) = E \frac{\epsilon_t^2}{f^2(x_t)} \left[\frac{1}{h_n} E\left(K\left(\frac{x_i - x_t}{h_n}\right) \left(\frac{x_i - x_t}{h_n}\right)^2 m^{(2)}(x_{it}) | Z_t\right) \right]^2$$

$$\rightarrow E(\sigma_K^2 m^{(2)}(x_t))^2 \sigma^2(x_t) < \infty, \text{ so } \frac{2}{n} \sum_{t=1}^n \int \phi_n(Z_t, Z_i) dP(Z_i) = O_p(n^{-\frac{1}{2}} h_n^2) = o_p(n^{-\frac{1}{2}}).$$

On the other hand, $E\phi_n^2(Z_t, Z_i) \leq c[E\psi_n^2(Z_t, Z_i) + E\psi_n^2(Z_i, Z_t)]$.
 $\frac{1}{h_n^3} E\psi_n^2(Z_t, Z_i) = \frac{1}{h_n} E \frac{1}{4f^2(x_t)} K^2\left(\frac{x_i - x_t}{h_n}\right) \left(\frac{x_i - x_t}{h_n}\right)^4 (m^{(2)}(x_{it}))^2 \sigma^2(x_t)$
 $\rightarrow \frac{1}{4} \int K^2(\psi) \psi^4 d\psi \int (m^{(2)}(x_t))^2 \sigma^2(x_t) dx_t < \infty$. So $n^{-1} (E\phi_n^2(Z_t, Z_i))^{\frac{1}{2}} = O_p(n^{-1} h_n^{\frac{3}{2}}) = o_p(n^{-\frac{1}{2}})$.

With the results above we have $I_{22} = o_p(n^{-\frac{1}{2}})$.

$$|I_{23}| \leq \sup_{x_t \in G} |w_n(x_t)| \frac{2}{n} \sum_{t=1}^n |\epsilon_t| = o_p(n^{-\frac{1}{2}}), \text{ since } \sup_{x_t \in G} |w_n(x_t)| = O_p(h_n \left(\frac{nh_n}{l n n}\right)^{-\frac{1}{2}}) + O_p(h_n^3) = o_p(n^{-\frac{1}{2}})$$

by A6 and $E|\epsilon_t| < \infty$ by A3. So in all we have $I_2 = o_p(n^{-\frac{1}{2}})$.

With (i)-(iii), we conclude that $\frac{1}{n} \sum_{t=1}^n (y_t - \hat{m}(x_t))^2 = \frac{1}{n} \sum_{t=1}^n \epsilon_t^2 + o_p(n^{-\frac{1}{2}})$.

So $B_1 = \frac{1}{n} \sum_{t=1}^n (y_t - \hat{m}(x_t))^2 - E(y_t - m(x_t))^2 = \frac{1}{n} \sum_{t=1}^n \epsilon_t^2 - E\sigma^2(x_t) + o_p(n^{-\frac{1}{2}})$.

(3) It is easy to see that $\frac{1}{n} \sum_{t=1}^n (y_t - \bar{y})^2 \xrightarrow{p} V(y_t) > 0$, so with results in (1) and (2),

$$\begin{aligned} & \tilde{R}^2 - R^2 \\ &= -\left\{ \left[\frac{1}{n} \sum_{t=1}^n (y_t - \bar{y})^2 \right]^{-1} - (V(y_t))^{-1} + (V(y_t))^{-1} \right\} \left\{ \frac{1}{n} \sum_{t=1}^n \epsilon_t^2 - E\sigma^2(x_t) + o_p(n^{-\frac{1}{2}}) \right\} \\ & \quad - \frac{1}{n} \sum_{t=1}^n \left[\frac{E\sigma^2(x_t)}{V(y_t)} (y_t^2 - Ey_t^2) - 2(y_t - Em(x_t)) \frac{E\sigma^2(x_t)}{V(y_t)} Em(x_t) \right] + O_p(n^{-1}) \} \\ &= -\left[(V(y_t))^{-1} + o_p(1) \right] \left\{ \frac{1}{n} \sum_{t=1}^n [\epsilon_t^2 - E\sigma^2(x_t) - \frac{E\sigma^2(x_t)}{V(y_t)} (y_t^2 - Ey_t^2)] \right. \\ & \quad \left. + 2(y_t - Em(x_t)) \frac{E\sigma^2(x_t)}{V(y_t)} Em(x_t) \right\} + o_p(n^{-\frac{1}{2}}) \} \\ &= -\left[(V(y_t))^{-1} \right] \left\{ \frac{1}{n} \sum_{t=1}^n [\epsilon_t^2 - \frac{E\sigma^2(x_t)}{V(y_t)} (y_t - E(y_t))^2] \right\} + o_p(n^{-\frac{1}{2}}) \end{aligned}$$

Let $W_t = \epsilon_t^2 - E\sigma^2(x) - \frac{E\sigma^2(x)}{V(y)} (y_t^2 - Ey_t^2) + 2\frac{E\sigma^2(x)}{V(y)} Em(x)(y_t - Em(x)) = \epsilon_t^2 - \frac{E\sigma^2(x_t)}{V(y_t)} (y_t - E(y_t))^2$. Since W_t is IID, with zero mean and variance EW_t^2 , we apply the Central Limit Theorem to obtain

$$\sqrt{n} \frac{1}{n} \sum_{t=1}^n W_t \xrightarrow{d} N(0, EW_t^2), \text{ and thus } \sqrt{n}(\tilde{R}^2 - R^2) \xrightarrow{d} N(0, (V(y_t))^{-2} EW_t^2).$$

(4) We show that $I(\frac{1}{n} \sum_{t=1}^n (y_t - \hat{m}(x_t))^2) \leq \frac{1}{n} \sum_{t=1}^n (y_t - \bar{y})^2 - 1 = o_{a.s.}(1)$.

Note by Kolmogorov's Theorem and with assumptions A1 and A2, we have

$$\frac{1}{n} \sum_{t=1}^n (y_t - \bar{y})^2 = \frac{1}{n} \sum_{t=1}^n y_t^2 - \bar{y}^2 \xrightarrow{a.s.} V(y_t), \text{ since } \frac{1}{n} \sum_{t=1}^n y_t^2 \xrightarrow{a.s.} Ey_t^2 \text{ and } \bar{y}^2 \xrightarrow{a.s.} (Ey_t)^2.$$

From result (2), $\frac{1}{n} \sum_{t=1}^n (y_t - \hat{m}(x_t))^2 = \frac{1}{n} \sum_{t=1}^n [\epsilon_t^2 + 2(m(x_t) - \hat{m}(x_t))\epsilon_t + (m(x_t) - \hat{m}(x_t))^2]$.

By Kolmogorov's Theorem again, we have $\frac{1}{n} \sum_{t=1}^n \epsilon_t^2 \xrightarrow{a.s.} E\epsilon_t^2 = E\sigma^2(x_t)$. As shown in result (2)(ii) above,

$\sup_{x_t \in G} |\hat{m}(x_t) - m(x_t)| = o_{a.s.}(1)$. So

$$\left| \frac{1}{n} \sum_{t=1}^n 2(m(x_t) - \hat{m}(x_t))\epsilon_t \right| \leq o_{a.s.}(1) \frac{1}{n} \sum_{t=1}^n |\epsilon_t| = o_{a.s.}(1), \text{ since } \frac{1}{n} \sum_{t=1}^n |\epsilon_t| \xrightarrow{a.s.} E|\epsilon_t| < (E\epsilon_t^2)^{\frac{1}{2}} < \infty.$$

Similarly, we obtain $\frac{1}{n} \sum_{t=1}^n (m(x_t) - \hat{m}(x_t))^2 = o_{a.s.}(1)$. So we have $\frac{1}{n} \sum_{t=1}^n (y_t - \hat{m}(x_t))^2 \xrightarrow{a.s.} E\sigma^2(x_t)$.

Combine results above, we have $\frac{1}{n} \sum_{t=1}^n (y_t - \bar{y})^2 - \frac{1}{n} \sum_{t=1}^n (y_t - \hat{m}(x_t))^2 \xrightarrow{a.s.} V(y_t) - E\sigma^2(x_t) = V(m(x_t)) > 0$

under H_1 . So we obtain the claimed result that $I(\frac{1}{n} \sum_{t=1}^n (y_t - \hat{m}(x_t))^2) \leq \frac{1}{n} \sum_{t=1}^n (y_t - \bar{y})^2 - 1 = o_{a.s.}(1)$.

(5) Claim: $\sqrt{n}(I(\cdot) - 1)\tilde{R}^2 = o_p(1)$, where $I(\cdot) = I(\frac{1}{n} \sum_{t=1}^n (y_t - \hat{m}(x_t))^2) \leq \frac{1}{n} \sum_{t=1}^n (y_t - \bar{y})^2$.

Since $\tilde{R}^2 \xrightarrow{p} R^2 > 0$, we only need to show $\sqrt{n}(I(\cdot) - 1) = o_p(1)$.

By definition $I(\frac{1}{n} \sum_{t=1}^n (y_t - \hat{m}(x_t))^2) \leq \frac{1}{n} \sum_{t=1}^n (y_t - \bar{y})^2$

$$= I(\frac{1}{n} \sum_{t=1}^n (y_t - \hat{m}(x_t))^2 - E\sigma^2(x) + E\sigma^2(x) \leq \frac{1}{n} \sum_{t=1}^n (y_t - \bar{y})^2 - V(y) + V(y)).$$

Define the event $D = \{w : |I(\cdot) - 1| \neq 0\}$, $A = \{w : |\frac{1}{n} \sum_{t=1}^n (y_t - \hat{m}(x_t))^2 - E\sigma^2(x)| > \delta_1\}$, and $B =$

$\{w : |\frac{1}{n} \sum_{t=1}^n (y_t - \bar{y})^2 - V(y)| > \delta_2\}$ for some $\delta_1, \delta_2 > 0$. Since $V(y) > E\sigma^2(x)$ under the alternative,

we have $D \subset A \cup B$. From (1) and (2) above, we have $\frac{1}{n} \sum_{t=1}^n (y_t - \hat{m}(x_t))^2 - E\sigma^2(x) = O_p(n^{-1/2})$, and

$\frac{1}{n} \sum_{t=1}^n (y_t - \bar{y})^2 - V(y) = O_p(n^{-1/2})$. So when $I(\cdot) \neq 1$,

$$|\sqrt{n}(I(\cdot) - 1)| = \sqrt{n}I(D) \leq \sqrt{n}I(D)(I(A) + I(B)).$$

With event A , $\sqrt{n}I(A) = \sqrt{n}1 < \frac{1}{\delta_1}|\sqrt{n}(\frac{1}{n}\sum_{t=1}^n(y_t - \hat{m}(x_t))^2 - E\sigma^2(x))| = O_p(1)$.

With event B , $\sqrt{n}I(B) = \sqrt{n}1 < \frac{1}{\delta_2}|\sqrt{n}(\frac{1}{n}\sum_{t=1}^n(y_t - \bar{y})^2 - V(y))| = O_p(1)$.

Since $I(D) = o_p(1)$, we have the claimed result that $\sqrt{n}(I(\cdot) - 1) = o_p(1)$.

By results (3)-(5) above, we have

$$\sqrt{n}(\hat{R}^2 - R^2) = \sqrt{n}(\tilde{R}^2 I(\cdot) - \tilde{R}^2 + \tilde{R}^2 - R^2) \xrightarrow{d} N(0, (V(y_t))^{-2} E W_t^2).$$

Theorem 1: *Proof.*

(a) Under H_0 , $P(E(y_t|x_t) = \mu) = 1$, so we have $m(x_t) = \mu$, and thus $y_t = \mu + \epsilon_t$, $V(y_t) = E\sigma^2(x_t)$

and $R^2 = 0$. Recall the definition of $\tilde{R}^2 = \frac{\frac{1}{n}\sum_{t=1}^n(y_t - \bar{y})^2 - \frac{1}{n}\sum_{t=1}^n(y_t - \hat{m}(x_t))^2}{\frac{1}{n}\sum_{t=1}^n(y_t - \bar{y})^2}$. We have from proof of Theorem

3 that $\frac{1}{n}\sum_{t=1}^n(y_t - \bar{y})^2 \xrightarrow{p} V(y_t) = E\sigma^2(x_t)$ under H_1 .

(1) As in result (1) of Theorem 3's proof,

$$\bar{y}^2 = \frac{2}{n}\sum_{t=1}^n y_t E m(x_t) - (E m(x_t))^2 + O_p(n^{-1}) = \mu^2 + \frac{2}{n}\sum_{t=1}^n \mu \epsilon_t + O_p(n^{-1}), \text{ so}$$

$$\begin{aligned} \frac{1}{n}\sum_{t=1}^n(y_t - \bar{y})^2 &= \frac{1}{n}\sum_{t=1}^n y_t^2 - \bar{y}^2 = \frac{1}{n}\sum_{t=1}^n(\mu + \epsilon_t)^2 - \bar{y}^2 \\ &= \mu^2 + \frac{2}{n}\sum_{t=1}^n \mu \epsilon_t + \frac{1}{n}\sum_{t=1}^n \epsilon_t^2 - (\mu^2 + \frac{2}{n}\sum_{t=1}^n \mu \epsilon_t + O_p(n^{-1})) = \frac{1}{n}\sum_{t=1}^n \epsilon_t^2 + O_p(n^{-1}). \end{aligned}$$

$$(2) \frac{1}{n}\sum_{t=1}^n(y_t - \hat{m}(x_t))^2 = \frac{1}{n}\sum_{t=1}^n \epsilon_t^2 + \frac{2}{n}\sum_{t=1}^n(\mu - \hat{m}(x_t))\epsilon_t + \frac{1}{n}\sum_{t=1}^n(\mu - \hat{m}(x_t))^2.$$

Define $e' = (1, 0)$, and we follow result (2)(i) in Theorem 3 to write

$$\hat{m}(x_t) - \mu = \frac{1}{nh_n}\sum_{i=1}^n W_n\left(\frac{x_i - x_t}{h_n}, x_t\right)\epsilon_i = e' S_n^{-1}(x_t) \begin{bmatrix} \frac{1}{nh_n}\sum_{i=1}^n K\left(\frac{x_i - x_t}{h_n}\right)\epsilon_i \\ \frac{1}{nh_n}\sum_{i=1}^n K\left(\frac{x_i - x_t}{h_n}\right)\frac{x_i - x_t}{h_n}\epsilon_i \end{bmatrix}$$

$$(i) \text{ So } \frac{2}{n}\sum_{t=1}^n(\mu - \hat{m}(x_t))\epsilon_t = -\frac{2}{n^2 h_n}\sum_{t=1}^n \sum_{i=1}^n e' S_n^{-1}(x_t) K\left(\frac{x_i - x_t}{h_n}\right)\epsilon_i \epsilon_t \begin{bmatrix} 1 \\ \frac{x_i - x_t}{h_n} \end{bmatrix}.$$

Recall the result in (2)(i) in Theorem 3 that $S_n(x_t) = \begin{bmatrix} S_{0n}(x_t) & S_{1n}(x_t) \\ S_{1n}(x_t) & S_{2n}(x_t) \end{bmatrix}$. Define

$$S_{jn}(x_t) = \frac{1}{nh_n}\sum_{i=1}^n K\left(\frac{x_i - x_t}{h_n}\right)\left(\frac{x_i - x_t}{h_n}\right)^j, j = 0, 1, 2. \text{ We apply Lemma 2 to obtain } \sup_{x_t \in G} |S_{jn}(x_t) - ES_{jn}(x_t)| =$$

$O_{a.s.}((\frac{nh_n}{\ln n})^{-\frac{1}{2}})$ with assumption A6 and A7. Furthermore, $ES_{jn}(x_t) = \int K(\psi)\psi^j f(x_t + h_n\psi)d\psi \rightarrow f(x_t) \int K(\psi)\psi^j d\psi$ uniformly over $x_t \in G$ with assumption A4. So $\sup_{x_t \in G} |S_{jn}(x_t) - S_j(x_t)| = o_{a.s.}(1)$.

With similar arguments, $\sup_{x_t \in G} |S_{jn}^{-1}(x_t) - S_j^{-1}(x_t)| = o_{a.s.}(1)$. So we obtain

$$\begin{aligned} \frac{2}{n}\sum_{t=1}^n(\mu - \hat{m}(x_t))\epsilon_t &= -\frac{2}{n^2 h_n}\sum_{t=1}^n \sum_{i=1}^n e' [S_n^{-1}(x_t) - S^{-1}(x_t) + S^{-1}(x_t)] K\left(\frac{x_i - x_t}{h_n}\right)\epsilon_i \epsilon_t \begin{bmatrix} 1 \\ \frac{x_i - x_t}{h_n} \end{bmatrix} \\ &= -\frac{2}{n^2 h_n}\sum_{t=1}^n \sum_{i=1}^n e' S^{-1}(x_t) K\left(\frac{x_i - x_t}{h_n}\right)\epsilon_i \epsilon_t \begin{bmatrix} 1 \\ \frac{x_i - x_t}{h_n} \end{bmatrix} (1 + o_{a.s.}(1)) \\ &= -\frac{2}{n^2 h_n}\sum_{t=1}^n \sum_{i=1}^n \frac{1}{f(x_t)} K\left(\frac{x_i - x_t}{h_n}\right)\epsilon_i \epsilon_t (1 + o_{a.s.}(1)) \\ &= \left[-\frac{2}{n^2 h_n}\sum_{t=1}^n \frac{1}{f(x_t)} K(0)\epsilon_t^2 - \frac{2}{n^2 h_n}\sum_{t=1}^n \sum_{\substack{i=1 \\ t \neq i}}^n \frac{1}{f(x_t)} K\left(\frac{x_i - x_t}{h_n}\right)\epsilon_i \epsilon_t\right] (1 + o_{a.s.}(1)) \\ &= [A_{1n} + A_{2n}](1 + o_{a.s.}(1)) \end{aligned}$$

where $o_{a.s.}(1)$ above indicates the term of smaller magnitude almost surely.

$$\begin{aligned}
\text{(ii)} \quad & \frac{1}{n} \sum_{t=1}^n (\mu - \hat{m}(x_t))^2 = \frac{1}{n} \sum_{t=1}^n \left[e' S_n^{-1}(x_t) \frac{1}{nh_n} \sum_{i=1}^n K\left(\frac{x_i - x_t}{h_n}\right) \epsilon_i \left[\frac{1}{h_n} \right] \right]^2 \\
& \text{since } \sup_{x_t \in G} |S_{jn}^{-1}(x_t) - S_j^{-1}(x_t)| = o_{a.s.}(1) \\
= & \underbrace{\frac{1}{n^3 h_n^2} \sum_{t=1}^n \sum_{i=1}^n \sum_{j=1}^n \frac{1}{f^2(x_t)} K\left(\frac{x_i - x_t}{h_n}\right) K\left(\frac{x_j - x_t}{h_n}\right) \epsilon_i \epsilon_j (1 + o_{a.s.}(1))}_I
\end{aligned}$$

(a) When $t = i = j$, $I = \frac{1}{n^3 h_n^2} \sum_{t=1}^n \frac{1}{f^2(x_t)} K^2(0) \epsilon_t^2$. Since $|E \frac{1}{n} \sum_{t=1}^n \frac{1}{f^2(x_t)} \epsilon_t^2| = E \frac{\sigma^2(x_t)}{f^2(x_t)} < \infty$ with assumptions A3 and A4, $I = O_p((nh_n)^{-2}) = o_p(n^{-1})$.

(b) When $t = i$ (or $t = j$), $I = \frac{2}{n^3 h_n^2} \sum_{t=1}^n \sum_{j=1, j \neq t}^n \frac{1}{2\sqrt{h_n} f^2(x_t)} K(0) K\left(\frac{x_j - x_t}{h_n}\right) \epsilon_t \epsilon_j = \frac{2}{n^3 h_n^2} \sum_{t=1}^n \sum_{j=1, j \neq t}^n \psi_n(Z_t, Z_j)$

where $Z_t = (x_t, \epsilon_t)$. We apply Lemma 1 on

$$nh_n^{\frac{3}{2}} I = \frac{1}{n^2} \sum_{t=1}^n \sum_{j=1, j \neq t}^n (\psi_n(Z_t, Z_j) + \psi_n(Z_j, Z_t)) = \frac{1}{n^2} \sum_{t=1}^n \sum_{j=1, j \neq t}^n \phi_n(Z_t, Z_j).$$

Note $\int \phi_n(Z_t, Z_j) dP(Z_j) = 0$ and $E\phi_n(Z_t, Z_j) = 0$. So $nh_n^{\frac{3}{2}} I = O_p(n^{-1} (E\phi_n^2(Z_t, Z_j))^{\frac{1}{2}})$.
 $E\phi_n^2(Z_t, Z_j) \leq c E\psi_n^2(Z_t, Z_j) = \frac{c}{4h_n} K^2(0) E \frac{1}{f^4(x_t)} K^2\left(\frac{x_j - x_t}{h_n}\right) \epsilon_t^2 \epsilon_j^2$
 $\rightarrow \frac{c}{4} K^2(0) E \frac{1}{f^3(x_t)} \sigma^4(x_t) \int K^2(\psi) d\psi < \infty$ with assumption A3 and A9.

So we conclude $nh_n^{\frac{3}{2}} I = O_p(n^{-1})$ and $I = o_p(n^{-1})$.

(c) When $i = j$ but $t \neq i$, $I = \frac{1}{n^3 h_n^2} \sum_{t=1}^n \sum_{i=1, i \neq t}^n \frac{1}{f^2(x_t)} K^2\left(\frac{x_i - x_t}{h_n}\right) \epsilon_i^2 = A_n$.

(d) When indices t, i and j are distinct, we let $\psi_n(Z_t, Z_i, Z_j) = \frac{1}{h_n^2} \frac{1}{f^2(x_t)} K\left(\frac{x_i - x_t}{h_n}\right) K\left(\frac{x_j - x_t}{h_n}\right) \epsilon_i \epsilon_j$,
 $I = \frac{1}{n^3} \sum_{t=1}^n \sum_{i=1, i \neq t}^n \sum_{j=1, j \neq t, j \neq i}^n \psi_n(Z_t, Z_i, Z_j) = \frac{1}{3n^3} \sum_{t=1}^n \sum_{i=1, i \neq t}^n \sum_{j=1, j \neq t, j \neq i}^n (\underbrace{\psi_n(Z_t, Z_i, Z_j) + \psi_n(Z_i, Z_t, Z_j) + \psi_n(Z_j, Z_i, Z_t)}_{\phi_n(Z_t, Z_i, Z_j)})$
 $= \frac{1}{n^3} \sum_{1=t < i < j=n} \sum \sum \sum 2\phi_n(Z_t, Z_i, Z_j)$, $\phi_n(Z_t, Z_i, Z_j)$ is symmetric.
 $= \frac{1}{3} \left[\frac{6}{n^3} - \binom{n}{3}^{-1} + \binom{n}{3}^{-1} \right] \sum_{1=t < i < j=n} \sum \sum \sum \phi_n(Z_t, Z_i, Z_j)$

Since $\frac{6}{n^3} - \binom{n}{3}^{-1} = O(n^{-4})$ and we note $U_n = \binom{n}{3}^{-1} \sum_{1=t < i < j=n} \sum \sum \sum \phi_n(Z_t, Z_i, Z_j)$ is a third order

U-statistic. We use the H-decomposition in Lemma 3. $E\phi_n(Z_t, Z_i, Z_j) = 0$, $E(\phi_n(Z_t, Z_i, Z_j)|Z_t) = 0$, and

$$\begin{aligned}
& E(\phi_n(Z_t, Z_i, Z_j)|Z_t, Z_i) = E(\psi_n(Z_t, Z_i, Z_j) + \psi_n(Z_i, Z_t, Z_j) + \psi_n(Z_j, Z_i, Z_t)|Z_t, Z_i) \\
& = \frac{\epsilon_i \epsilon_t}{h_n} E\left(\frac{1}{h_n f^2(x_j)} K\left(\frac{x_j - x_t}{h_n}\right) K\left(\frac{x_j - x_i}{h_n}\right) | Z_t, Z_i\right) = \phi_{2n}(Z_t, Z_i),
\end{aligned}$$

So $U_n = \frac{6}{n(n-1)} \sum_{t=1}^n \sum_{i=1, i \neq t}^n \phi_{2n}(Z_t, Z_i) + O_p(H_n^{(3)})$. $Var(H_n^{(3)}) = O(n^{-3} \sigma_{3n}^2)$. $\sigma_{3n}^2 \leq 3c E\psi_n^2(Z_t, Z_i, Z_j)$.

$h_n^2 E\psi_n^2(Z_t, Z_i, Z_j) = \frac{1}{h_n^2} E \frac{1}{f^4(x_t)} K^2\left(\frac{x_i - x_t}{h_n}\right) K^2\left(\frac{x_j - x_t}{h_n}\right) \sigma^2(x_i) \sigma^2(x_j) \rightarrow (\int K^2(\psi) d\psi)^2 E \frac{\sigma^4(x_t)}{f^2(x_t)} < \infty$ with assumptions A4 and A9. So $\sigma_{3n}^2 = O(h_n^{-2})$. $Var(H_n^{(3)}) = O(n^{-3} h_n^{-2}) = o(n^{-2})$, and thus

$$U_n = \frac{6}{n(n-1)} \sum_{t=1}^n \sum_{i=1, i \neq t}^n \phi_{2n}(Z_t, Z_i) + o(n^{-1}) \text{ and } I = \frac{1}{n(n-1)} \sum_{t=1}^n \sum_{i=1, i \neq t}^n [\phi_{2n}(Z_t, Z_i) + \phi_{2n}(Z_i, Z_t)] + o(n^{-1}).$$

We conclude that

$$\begin{aligned}
& \frac{1}{n} \sum_{t=1}^n (\mu - \hat{m}(x_t))^2 - A_n(1 + o_{a.s.}(1)) = \frac{1}{n(n-1)} \sum_{t=1}^n \sum_{\substack{i=1 \\ t < i}}^n [\phi_{2n}(Z_t, Z_i) + \phi_{2n}(Z_i, Z_t)] + o(n^{-1})(1 + o_{a.s.}(1)) \\
&= \frac{1}{n(n-1)} \sum_{t=1}^n \sum_{\substack{i=1 \\ t < i}}^n \left[\frac{\epsilon_i \epsilon_t}{h_n} E\left(\frac{1}{h_n f^2(x_j)} K\left(\frac{x_i - x_j}{h_n}\right) K\left(\frac{x_t - x_j}{h_n}\right) \middle| Z_t, Z_i\right) \right. \\
&\quad \left. + \frac{\epsilon_t \epsilon_i}{h_n} E\left(\frac{1}{h_n f^2(x_j)} K\left(\frac{x_t - x_j}{h_n}\right) K\left(\frac{x_i - x_j}{h_n}\right) \middle| Z_i, Z_t\right) \right] (1 + o_{a.s.}(1)) + o_p(n^{-1}) \\
&= A_{3n}(1 + o_{a.s.}(1)) + o_p(n^{-1})
\end{aligned}$$

We summarize the results obtained in (2) as

$$\frac{1}{n} \sum_{t=1}^n (y_t - \hat{m}(x_t))^2 - (A_n + A_{1n})(1 + o_{a.s.}(1)) = \frac{1}{n} \sum_{t=1}^n \epsilon_t^2 + (A_{2n} + A_{3n})(1 + o_{a.s.}(1)) + o_p(n^{-1}).$$

A_n and A_{1n} are the bias terms. We focus on A_{2n} and A_{3n} to determine the asymptotic distribution.

$$\begin{aligned}
(3) \quad A_{2n} &= \frac{1}{n^2 h_n} \sum_{t=1}^n \sum_{\substack{i=1 \\ t < i}}^n \left[(-2) \frac{1}{f(x_t)} K\left(\frac{x_i - x_t}{h_n}\right) \epsilon_i \epsilon_t - 2 \frac{1}{f(x_i)} K\left(\frac{x_t - x_i}{h_n}\right) \epsilon_t \epsilon_i \right]. \\
A_{3n} &= \left[\frac{1}{n(n-1)} - \frac{1}{n^2} + \frac{1}{n^2} \right] \sum_{t=1}^n \sum_{\substack{i=1 \\ t < i}}^n [\phi_{2n}(Z_t, Z_i) + \phi_{2n}(Z_i, Z_t)] \\
&= [O(n^{-3}) + \frac{1}{n^2}] \sum_{t=1}^n \sum_{\substack{i=1 \\ t < i}}^n [\phi_{2n}(Z_t, Z_i) + \phi_{2n}(Z_i, Z_t)] = \frac{1}{n^2} \sum_{t=1}^n \sum_{\substack{i=1 \\ t < i}}^n [\phi_{2n}(Z_t, Z_i) + \phi_{2n}(Z_i, Z_t)] + o_p(n^{-1}),
\end{aligned}$$

since $A_{31n} = \frac{1}{n^2} \sum_{t=1}^n \sum_{\substack{i=1 \\ t < i}}^n [\phi_{2n}(Z_t, Z_i) + \phi_{2n}(Z_i, Z_t)]$ is a second order U-statistic in Lemma 1, we easily obtain

$A_{31n} = O_p(n^{-1} (E\phi_{2n}^2(Z_t, Z_i))^{\frac{1}{2}})$ as ϵ_t has conditional mean zero as in Assumption A3.

Furthermore, $h_n E\phi_{2n}^2(Z_t, Z_i) = \frac{1}{h_n} E\sigma^2(x_i)\sigma^2(x_t) [E \frac{1}{h_n f^2(x_j)} K(\frac{x_i - x_j}{h_n}) K(\frac{x_t - x_j}{h_n}) | Z_t, Z_i]^2$
 $= \int \sigma^2(x_t)\sigma^2(x_t - h_n\psi_1) [\int K(\psi)K(\psi_1 + \psi) \frac{1}{f(x_t - h(\psi + \psi_1))} d\psi]^2 f(x_t)f(x_t - h_n\psi_1) d\psi_1 dx_t$
 $\rightarrow \int \kappa^2(\psi_1) d\psi_1 E \frac{\sigma^4(x_t)}{f(x_t)} < \infty$, so $(E\phi_{2n}^2(Z_t, Z_i))^{\frac{1}{2}} = O(h_n^{-\frac{1}{2}})$ and $A_{31n} = O_p((n^2 h_n)^{-\frac{1}{2}})$.

$$\begin{aligned}
A_{2n} + A_{31n} &= \frac{1}{n^2 h_n} \sum_{t=1}^n \sum_{\substack{i=1 \\ t < i}}^n \left[(-2) \frac{1}{f(x_t)} K\left(\frac{x_i - x_t}{h_n}\right) \epsilon_i \epsilon_t - 2 \frac{1}{f(x_i)} K\left(\frac{x_t - x_i}{h_n}\right) \epsilon_t \epsilon_i \right. \\
&\quad \left. + \epsilon_i \epsilon_t E\left(\frac{1}{h_n f^2(x_j)} K\left(\frac{x_i - x_j}{h_n}\right) K\left(\frac{x_t - x_j}{h_n}\right) \middle| Z_t, Z_i\right) + \epsilon_t \epsilon_i E\left(\frac{1}{h_n f^2(x_j)} K\left(\frac{x_t - x_j}{h_n}\right) K\left(\frac{x_i - x_j}{h_n}\right) \middle| Z_i, Z_t\right) \right] \\
&= \frac{1}{n^2 h_n} \sum_{t=1}^n \sum_{\substack{i=1 \\ t < i}}^n [\psi_n(Z_t, Z_i) + \psi_n(Z_i, Z_t) + \psi'_n(Z_t, Z_i) + \psi'_n(Z_i, Z_t)] \\
&= \frac{1}{n^2 h_n} \sum_{t=1}^n \sum_{\substack{i=1 \\ t < i}}^n \phi_n(Z_t, Z_i).
\end{aligned}$$

Since $\phi_n(Z_t, Z_i)$ is symmetric and $E(\phi_n(Z_t, Z_i) | Z_t) = 0$, $A_{2n} + A_{31n}$ is a degenerated second order U-statistic. In the following we show

$$(i) \quad \frac{1}{h_n} E\phi_n^2(Z_t, Z_i) \rightarrow 2[2E \frac{\sigma^4(x_t)}{f(x_t)} (4 \int K^2(\psi) d\psi + \int \kappa^2(\psi) d\psi - 4 \int K(\psi) \kappa(\psi) d\psi)] = 2\sigma_\phi^2.$$

$$(ii) \quad E\phi_n^4(Z_t, Z_i) = O(h_n).$$

$$(iii) \quad \text{For } G_n(Z_1, Z_2) = E(\phi_n(Z_t, Z_1)\phi_n(Z_t, Z_2) | Z_1, Z_2), \text{ we have } EG_n^2(Z_1, Z_2) = O(h_n^3).$$

From (i)-(iii), we have $\frac{EG_n^2(Z_1, Z_2)}{(E\phi_n^2(Z_t, Z_i))^2} = \frac{O(h_n^3)}{O(h_n^2)} \rightarrow 0$, and $\frac{\frac{1}{n} E\phi_n^4(Z_t, Z_i)}{(E\phi_n^2(Z_t, Z_i))^2} = \frac{\frac{1}{n} O(h_n)}{O(h_n^2)} \rightarrow 0$, which is condition (2.1) in Hall (1984). So we apply Hall's Central Limit Theorem to have

$$nh_n^{\frac{1}{2}} (A_{2n} + A_{31n}) \xrightarrow{d} N(0, \sigma_\phi^2).$$

$$\begin{aligned}
(i) \quad E\phi_n^2(Z_t, Z_i) &= E\psi_n^2(Z_t, Z_i) + E\psi_n^2(Z_i, Z_t) + E\psi_n'^2(Z_t, Z_i) + E\psi_n'^2(Z_i, Z_t) \\
&\quad + 2E\psi_n(Z_t, Z_i)\psi_n(Z_i, Z_t) + 2E\psi_n(Z_t, Z_i)\psi_n'(Z_t, Z_i) + 2E\psi_n(Z_t, Z_i)\psi_n'(Z_i, Z_t) \\
&\quad + 2E\psi_n(Z_i, Z_t)\psi_n'(Z_t, Z_i) + 2E\psi_n(Z_i, Z_t)\psi_n'(Z_i, Z_t) + 2E\psi_n'(Z_t, Z_i)\psi_n'(Z_i, Z_t).
\end{aligned}$$

With assumptions A2, A3, A7 and A9,

$$\begin{aligned}
& \frac{1}{h_n} E\psi_n^2(Z_t, Z_i) = \frac{1}{h_n} E\psi_n^2(Z_i, Z_t) \\
&= \frac{4}{h_n} E \frac{1}{f^2(x_t)} K^2\left(\frac{x_i - x_t}{h_n}\right) \sigma^2(x_i) \sigma^2(x_t) \rightarrow 4 \int K^2(\psi) d\psi E \frac{\sigma^4(x_t)}{f(x_t)} < \infty.
\end{aligned}$$

$$\begin{aligned} \frac{1}{h_n} E\psi_n'^2(Z_t, Z_i) &= \frac{1}{h_n} E\psi_n'^2(Z_i, Z_t) \\ &= \frac{1}{h_n} E\sigma^2(x_i)\sigma^2(x_t) [E \frac{1}{h_n f^2(x_j)} K(\frac{x_i - x_j}{h_n}) K(\frac{x_t - x_j}{h_n}) | Z_t, Z_i)]^2 \rightarrow \int \kappa^2(\psi) d\psi E \frac{\sigma^4(x_t)}{f(x_t)} < \infty. \end{aligned}$$

With the assumption A7 that $K(\cdot)$ is a symmetric function, we obtain

$$\begin{aligned} \frac{2}{h_n} E\psi_n(Z_t, Z_i)\psi_n(Z_i, Z_t) &\rightarrow 8 \int K^2(\psi) d\psi E \frac{\sigma^4(x_t)}{f(x_t)}, \quad \frac{2}{h_n} E\psi_n(Z_t, Z_i)\psi_n'(Z_t, Z_i) \rightarrow -4 \int K(\psi)\kappa(\psi) d\psi E \frac{\sigma^4(x_t)}{f(x_t)}, \\ \frac{2}{h_n} E\psi_n(Z_i, Z_t)\psi_n'(Z_i, Z_t) &\rightarrow -4 \int K(\psi)\kappa(\psi) d\psi E \frac{\sigma^4(x_t)}{f(x_t)}, \quad \frac{2}{h_n} E\psi_n'(Z_t, Z_i)\psi_n'(Z_i, Z_t) \rightarrow 2 \int \kappa^2(\psi) d\psi E \frac{\sigma^4(x_t)}{f(x_t)}, \\ \frac{2}{h_n} E\psi_n(Z_t, Z_i)\psi_n'(Z_i, Z_t) &\rightarrow -4 \int K(\psi)\kappa(\psi) d\psi E \frac{\sigma^4(x_t)}{f(x_t)}, \\ \frac{2}{h_n} E\psi_n(Z_i, Z_t)\psi_n'(Z_t, Z_i) &\rightarrow -4 \int K(\psi)\kappa(\psi) d\psi E \frac{\sigma^4(x_t)}{f(x_t)}, \end{aligned}$$

so $\frac{1}{h_n} E\phi_n^2(Z_t, Z_i) \rightarrow 2\sigma_\phi^2$.

$$\begin{aligned} \text{(ii)} \quad E\phi_n^4(Z_t, Z_i) &\leq c[E\psi_n^4(Z_t, Z_i) + E\psi_n^4(Z_i, Z_t) + E\psi_n'^4(Z_t, Z_i) + E\psi_n'^4(Z_t, Z_i)] \text{ by } c_r \text{ inequality.} \\ \frac{1}{h_n} E\psi_n^4(Z_t, Z_i) &= \frac{1}{h_n} E\psi_n^4(Z_i, Z_t) = \frac{16}{h_n} E \frac{1}{f^4(x_t)} K^4(\frac{x_i - x_t}{h_n}) E(\epsilon_i^4 | x_i) E(\epsilon_t^4 | x_t) \\ &< \frac{16c}{h_n} E \frac{1}{f^4(x_t)} K^4(\frac{x_i - x_t}{h_n}) \text{ by assumption A9,} \\ &\rightarrow 16c \int K^4(\psi) d\psi \int \frac{1}{f^2(x_t)} dx_t. \\ \frac{1}{h_n} E\psi_n'^4(Z_t, Z_i) &= \frac{1}{h_n} E\psi_n'^4(Z_i, Z_t) \\ &= \frac{1}{h_n} E[E(\epsilon_i^4 | x_i) E(\epsilon_t^4 | x_t) [E \frac{1}{h_n f^2(x_j)} K(\frac{x_i - x_j}{h_n}) K(\frac{x_t - x_j}{h_n}) | Z_t, Z_i)]^4] \\ &< \frac{c}{h_n} E[[E \frac{1}{h_n f^2(x_j)} K(\frac{x_i - x_j}{h_n}) K(\frac{x_t - x_j}{h_n}) | Z_t, Z_i]]^4 \text{ by assumption A9,} \\ &\rightarrow c \int \kappa^4(\psi) d\psi \int \frac{1}{f^2(x_t)} dx_t. \end{aligned}$$

So we have $E\phi_n^4(Z_t, Z_i) = O(h_n)$.

$$\begin{aligned} \text{(iii)} \quad \text{Since } G_n(Z_1, Z_2) &= E(\phi_n(Z_t, Z_1)\phi_n(Z_t, Z_2) | Z_1, Z_2) \\ &= E\{[\psi_n(Z_t, Z_1) + \psi_n(Z_1, Z_t) + \psi_n'(Z_t, Z_1) + \psi_n'(Z_1, Z_t)] \\ &\quad \times [\psi_n(Z_t, Z_2) + \psi_n(Z_2, Z_t) + \psi_n'(Z_t, Z_2) + \psi_n'(Z_2, Z_t)] | Z_1, Z_2\} \\ &= G_1 + G_2 + \dots + G_{16}. \end{aligned}$$

$EG_n^2(Z_1, Z_2) = E[G_1 + \dots + G_{16}]^2 \leq c\{EG_1^2 + EG_2^2 + \dots + EG_{16}^2\}$. Let's consider

$$\begin{aligned} \frac{1}{h_n^3} EG_1^2 &= \frac{1}{h_n^3} E\{E[\psi_n(Z_t, Z_1)\psi_n(Z_t, Z_2) | Z_1, Z_2]\}^2 \\ &= \frac{16}{h_n} E\{\sigma^2(x_1)\sigma^2(x_2) [E \frac{1}{f^2(x_t)} K(\frac{x_1 - x_t}{h_n}) K(\frac{x_2 - x_t}{h_n}) \sigma^2(x_t) | x_1, x_2]\}^2 \\ &\rightarrow 16 \int \sigma^8(x_1) dx_1 \int \kappa^2(\psi_1) d\psi_1. \end{aligned}$$

We follow similar arguments to show the other terms in $EG_n^2(Z_1, Z_2) = O(h_n^3)$. So we have the claim in (iii).

(4) Combining results in (1)-(3), we conclude

$$nh_n^{\frac{1}{2}} \{ \frac{1}{n} \sum_{t=1}^n (y_t - \bar{y})^2 - \frac{1}{n} \sum_{t=1}^n (y_t - \hat{m}(x_t))^2 + (A_n + A_{1n})(1 + o_{a.s.}(1)) \} \xrightarrow{d} N(0, \sigma_\phi^2).$$

Since $\frac{1}{n} \sum_{t=1}^n (y_t - \bar{y})^2 \xrightarrow{p} E\sigma^2(x_t) > 0$, we have

$$nh_n^{\frac{1}{2}} \{ \hat{R}^2 + [\frac{1}{n} \sum_{t=1}^n (y_t - \bar{y})^2]^{-1} ((A_n + A_{1n})(1 + o_{a.s.}(1))) \} \xrightarrow{d} N(0, (E\sigma^2(x_t))^{-2} \sigma_\phi^2).$$

We follow the proof in Theorem 3 to show $I(\cdot) = I(\frac{1}{n} \sum_{t=1}^n (y_t - \bar{y})^2 \geq \frac{1}{n} \sum_{t=1}^n (y_t - \hat{m}(x_t))^2) \xrightarrow{p} 1$. So

$nh_n^{\frac{1}{2}} \{ \hat{R}^2 + I(\cdot) [\frac{1}{n} \sum_{t=1}^n (y_t - \bar{y})^2]^{-1} ((A_n + A_{1n})(1 + o_{a.s.}(1))) \} \xrightarrow{d} N(0, (E\sigma^2(x_t))^{-2} \sigma_\phi^2)$, and $T_n \xrightarrow{d} N(0, 1)$ as desired.

For Theorem 1(b), it is sufficient to prove the following results:

(i) $\hat{A}_n - A_n = o_p((nh_n^{\frac{1}{2}})^{-1})$. (ii) $\hat{A}_{1n} - A_{1n} = o_p((nh_n^{\frac{1}{2}})^{-1})$. (iii) $\hat{V}_T - V_T = o_p(1)$.
(i) Note under H_0 , $y_t = \mu + \epsilon_t$. We estimate ϵ_t by $\tilde{\epsilon}_t = y_t - \bar{y}$. So $\tilde{\epsilon}_t - \epsilon_t = (y_t - \bar{y}) - (y_t - \mu) = u - \bar{y} = O_p(n^{-\frac{1}{2}})$. Also note: $\sup_{x_t \in G} |\hat{f}(x_t) - f(x_t)| = \sup_{x_t \in G} |\hat{f}(x_t) - E\hat{f}(x_t) + E\hat{f}(x_t) - f(x_t)|$. Since by Lemma 2, $\sup_{x_t \in G} |\hat{f}(x_t) - E\hat{f}(x_t)| = O_{a.s.}((\frac{\ln n}{nh_n})^{\frac{1}{2}})$. From Theorem 3, we have $E\hat{f}(x_t) - f(x_t) = O(h_n)$ uniformly over $x_t \in G$. Let $L_n = (\frac{\ln n}{nh_n})^{\frac{1}{2}} + h_n$. Then $\sup_{x_t \in G} |\hat{f}(x_t) - f(x_t)| = O_p(L_n) = o_p(h_n^{\frac{1}{2}})$.

Furthermore, $\sup_{x_t \in G} |\hat{f}^{-1}(x_t) - f^{-1}(x_t)| \leq \frac{\sup_{x_t \in G} |\hat{f}(x_t) - f(x_t)|}{\inf_{x_t \in G} \hat{f}(x_t) \inf_{x_t \in G} f(x_t)}$. With assumption A4(1), $\inf_{x_t \in G} \hat{f}(x_t) \geq$

$\inf_{x_t \in G} (\hat{f}(x_t) - f(x_t)) + \inf_{x_t \in G} f(x_t) > 0$, since $\inf_{x_t \in G} (\hat{f}(x_t) - f(x_t)) \leq \sup_{x_t \in G} |\hat{f}(x_t) - f(x_t)| = o_p(1)$. So we also have $\sup_{x_t \in G} |\hat{f}^{-1}(x_t) - f^{-1}(x_t)| = O_p(L_n)$.

$$\begin{aligned} \hat{A}_n - A_n &= \frac{1}{n^3 h_n^2} \sum_{t=1}^n \sum_{\substack{i=1 \\ t \neq i}}^n K^2\left(\frac{x_i - x_t}{h_n}\right) \left[\frac{1}{\hat{f}^2(x_t)} - \frac{1}{f^2(x_t)} \right] [\tilde{\epsilon}_i^2 - \epsilon_i^2 + \epsilon_i^2] + \frac{1}{n^3 h_n^2} \sum_{t=1}^n \sum_{\substack{i=1 \\ t \neq i}}^n K^2\left(\frac{x_i - x_t}{h_n}\right) \frac{1}{f^2(x_t)} [\tilde{\epsilon}_i^2 - \epsilon_i^2] \\ &= I_1 + I_2 \\ I_1 &\leq O_p(L_n) \left\{ \frac{1}{n^3 h_n^2} \sum_{t=1}^n \sum_{\substack{i=1 \\ t \neq i}}^n K^2\left(\frac{x_i - x_t}{h_n}\right) \epsilon_i^2 + \frac{1}{n^3 h_n^2} \sum_{t=1}^n \sum_{\substack{i=1 \\ t \neq i}}^n K^2\left(\frac{x_i - x_t}{h_n}\right) |(\tilde{\epsilon}_i + \epsilon_i)(\tilde{\epsilon}_i - \epsilon_i)| \right\}. \end{aligned}$$

With Lemma 1, we easily obtain $\frac{1}{n^3 h_n^2} \sum_{t=1}^n \sum_{\substack{i=1 \\ t \neq i}}^n K^2\left(\frac{x_i - x_t}{h_n}\right) \epsilon_i^2 = O_p((nh_n)^{-1})$.

$$\begin{aligned} &\frac{1}{n^3 h_n^2} \sum_{t=1}^n \sum_{\substack{i=1 \\ t \neq i}}^n K^2\left(\frac{x_i - x_t}{h_n}\right) |(\tilde{\epsilon}_i - \epsilon_i + 2\epsilon_i)(\tilde{\epsilon}_i - \epsilon_i)| \\ &\leq \frac{1}{n^3 h_n^2} \sum_{t=1}^n \sum_{\substack{i=1 \\ t \neq i}}^n K^2\left(\frac{x_i - x_t}{h_n}\right) (\tilde{\epsilon}_i - \epsilon_i)^2 + \frac{2}{n^3 h_n^2} \sum_{t=1}^n \sum_{\substack{i=1 \\ t \neq i}}^n K^2\left(\frac{x_i - x_t}{h_n}\right) |(\tilde{\epsilon}_i - \epsilon_i)\epsilon_i| \\ &= I_{11} + I_{12} = O_p((n^{\frac{3}{2}} h_n)^{-1}), \end{aligned}$$

since $I_{11} = O_p(n^{-1}) O_p((nh_n)^{-1}) = O_p((n^2 h_n)^{-1})$ and $I_{12} = O_p(n^{-\frac{1}{2}}) O_p((nh_n)^{-1}) = O_p((n^{\frac{3}{2}} h_n)^{-1})$. Given the note above, we have $I_1 \leq O_p(L_n) O_p((nh_n)^{-1}) = o_p((nh_n^{\frac{1}{2}})^{-1})$.

$$\begin{aligned} I_2 &= \frac{1}{n^3 h_n^2} \sum_{t=1}^n \sum_{\substack{i=1 \\ t \neq i}}^n K^2\left(\frac{x_i - x_t}{h_n}\right) \frac{1}{f^2(x_t)} (\tilde{\epsilon}_i - \epsilon_i)(\tilde{\epsilon}_i - \epsilon_i + 2\epsilon_i) \\ &= \frac{1}{n^3 h_n^2} \sum_{t=1}^n \sum_{\substack{i=1 \\ t \neq i}}^n K^2\left(\frac{x_i - x_t}{h_n}\right) \frac{1}{f^2(x_t)} (\tilde{\epsilon}_i - \epsilon_i)^2 + \frac{1}{n^3 h_n^2} \sum_{t=1}^n \sum_{\substack{i=1 \\ t \neq i}}^n K^2\left(\frac{x_i - x_t}{h_n}\right) \frac{1}{f^2(x_t)} (\tilde{\epsilon}_i - \epsilon_i) 2\epsilon_i \\ &= I_{21} + I_{22} = O_p((n^{\frac{3}{2}} h_n)^{-1}) = o_p((nh_n^{\frac{1}{2}})^{-1}), \end{aligned}$$

since $I_{21} = O_p(n^{-1}) O_p((nh_n)^{-1}) = O_p((n^2 h_n)^{-1})$ and $I_{22} = O_p(n^{-\frac{1}{2}}) O_p((nh_n)^{-1}) = O_p((n^{\frac{3}{2}} h_n)^{-1})$. So we conclude $\hat{A}_n - A_n = o_p((nh_n^{\frac{1}{2}})^{-1})$.

$$\begin{aligned} \text{(ii)} \quad \hat{A}_{1n} - A_{1n} &= -\frac{2}{n^2 h_n} \sum_{t=1}^n \left[\frac{1}{\hat{f}(x_t)} - \frac{1}{f(x_t)} \right] K(0) (\tilde{\epsilon}_t^2 - \epsilon_t^2 + \epsilon_t^2) - \frac{2}{n^2 h_n} \sum_{t=1}^n \frac{K(0)}{f(x_t)} (\tilde{\epsilon}_t^2 - \epsilon_t^2) \\ &= o_p((nh_n^{\frac{1}{2}})^{-1}) \text{ with similar arguments.} \end{aligned}$$

(iii) Given assumption A2 that $V(y) = E\sigma^2(x) > 0$, we easily obtain $\frac{1}{n} \sum_{t=1}^n (y_t - \bar{y})^2 \xrightarrow{P} E\sigma^2(x)$. Thus $[\frac{1}{n} \sum_{t=1}^n (y_t - \bar{y})^2]^{-2} \xrightarrow{P} (E\sigma^2(x))^{-2}$. So to prove $\hat{V}_T - V_T = o_p(1)$, we only need to show

$$\begin{aligned} (1) \quad &\frac{1}{n^2} \sum_{t=1}^n \sum_{\substack{i=1 \\ t \neq i}}^n K\left(\frac{x_i - x_t}{h_n}\right) \frac{\epsilon_i^2 \epsilon_t^2}{h_n f^2(x_t)} \xrightarrow{P} E \frac{\sigma^4(x_t)}{f(x_t)} = \int \sigma^4(x_t) dx_t \text{ and} \\ (2) \quad &\frac{1}{n^2} \sum_{t=1}^n \sum_{\substack{i=1 \\ t \neq i}}^n K\left(\frac{x_i - x_t}{h_n}\right) \frac{\epsilon_i^2 \epsilon_t^2}{h_n f^2(x_t)} - \frac{1}{n^2} \sum_{t=1}^n \sum_{\substack{i=1 \\ t \neq i}}^n K\left(\frac{x_i - x_t}{h_n}\right) \frac{\epsilon_i^2 \epsilon_t^2}{h_n f^2(x_t)} = o_p(1). \\ (1) \quad &\frac{1}{n^2} \sum_{t=1}^n \sum_{\substack{i=1 \\ t \neq i}}^n K\left(\frac{x_i - x_t}{h_n}\right) \frac{\epsilon_i^2 \epsilon_t^2}{h_n f^2(x_t)} = \frac{1}{n^2} \sum_{t=1}^n \sum_{\substack{i=1 \\ t \neq i}}^n \psi_n(Z_t, Z_i) = \frac{1}{2n^2} \sum_{t=1}^n \sum_{\substack{i=1 \\ t \neq i}}^n (\psi_n(Z_t, Z_i) + \psi_n(Z_i, Z_t)) \\ &= \frac{1}{2n^2} \sum_{t=1}^n \sum_{\substack{i=1 \\ t \neq i}}^n \phi_n(Z_t, Z_i), \text{ where } Z_t = (x_t, \epsilon_t) \text{ and } \phi_n(Z_t, Z_i) \text{ is symmetric.} \\ &= \frac{1}{n} \sum_{t=1}^n \int \phi_n(Z_t, Z_i) dP(Z_i) - \frac{1}{2} E \phi_n(Z_t, Z_i) + O_p(n^{-1} (E \phi_n^2(Z_t, Z_i))^{\frac{1}{2}}) \text{ by Lemma 1.} \\ \frac{1}{2} E \phi_n(Z_t, Z_i) &= E \psi_n(Z_t, Z_i) = \frac{1}{h_n} \int K\left(\frac{x_i - x_t}{h_n}\right) \sigma^2(x_i) \sigma^2(x_t) \frac{f(x_i) f(x_t)}{f^2(x_t)} dx_i dx_t \\ &\rightarrow \int \sigma^4(x_t) dx_t. \end{aligned}$$

$$\begin{aligned} h_n E \phi_n^2(Z_t, Z_i) &\leq 4 h_n E \psi_n^2(Z_t, Z_i) = \frac{4}{h_n} \int K^2\left(\frac{x_i - x_t}{h_n}\right) E(\epsilon_i^4 | x_i) E(\epsilon_t^4 | x_t) \frac{f(x_t) f(x_i)}{f^4(x_t)} dx_i dx_t \\ &\rightarrow 4 \int K^2(\psi) d\psi E\left(\frac{E^2(\epsilon_i^4 | x_t)}{f^3(x_t)}\right) < \infty \text{ by assumption A9.} \end{aligned}$$

Since $E(\frac{1}{n} \sum_{t=1}^n \int \phi_n(Z_t, Z_i) dP(Z_i)) = E(\phi_n(Z_t, Z_i)) = 2E\psi_n(Z_t, Z_i) \rightarrow \int 2\sigma^4(x_t) dx_t$,

$$\begin{aligned} V(\frac{1}{n} \sum_{t=1}^n \int \phi_n(Z_t, Z_i) dP(Z_i)) &= \frac{1}{n} V(\int \psi_n(Z_t, Z_i) dP(Z_i) + \int \psi_n(Z_i, Z_t) dP(Z_i)) \\ &\leq \frac{C}{n} E \psi_n^2(Z_t, Z_i) = O(\frac{1}{nh_n}), \text{ so } V(\frac{1}{n} \sum_{t=1}^n \int \phi_n(Z_t, Z_i) dP(Z_i)) \rightarrow 0. \end{aligned}$$

Since $O_p(n^{-1}(E\phi_n^2(Z_t, Z_i))^{\frac{1}{2}}) = O_p(n^{-1}h_n^{-\frac{1}{2}})$, so $\frac{1}{n^2} \sum_{t=1}^n \sum_{i=1}^n K(\frac{x_i - x_t}{h_n}) \frac{\epsilon_i^2 \epsilon_t^2}{h_n f^2(x_t)} \xrightarrow{P} \int \sigma^4(x_t) dx_t$.

$$\begin{aligned} (2) \quad &\frac{1}{n^2} \sum_{t=1}^n \sum_{i=1}^n K(\frac{x_i - x_t}{h_n}) \frac{\tilde{\epsilon}_i^2 \tilde{\epsilon}_t^2}{h_n f^2(x_t)} - \frac{1}{n^2} \sum_{t=1}^n \sum_{i=1}^n K(\frac{x_i - x_t}{h_n}) \frac{\epsilon_i^2 \epsilon_t^2}{h_n f^2(x_t)} \\ &= \frac{1}{n^2} \sum_{t=1}^n \sum_{i=1}^n \frac{1}{h_n} \left[\frac{1}{f^2(x_t)} - \frac{1}{f^2(x_t)} \right] K(\frac{x_i - x_t}{h_n}) \tilde{\epsilon}_i^2 \tilde{\epsilon}_t^2 + \frac{1}{n^2} \sum_{t=1}^n \sum_{i=1}^n \frac{1}{h_n} \frac{1}{f^2(x_t)} K(\frac{x_i - x_t}{h_n}) (\tilde{\epsilon}_i^2 \tilde{\epsilon}_t^2 - \epsilon_i^2 \epsilon_t^2) \\ &= o_p(1). \end{aligned}$$

Given assumption A8, we have $\frac{1}{n^2} \sum_{t=1}^n \sum_{i=1}^n \frac{1}{h_n} \frac{1}{f^2(x_t)} K(\frac{x_i - x_t}{h_n}) |\epsilon_i|^l |\epsilon_t|^m = O_p(1)$ with the application of

Lemma 1 for $l, m = \{0, 1, 2\}$. Furthermore, $\sup_{x_t \in G} \left| \frac{1}{\tilde{f}(x_t)} - \frac{1}{f(x_t)} \right| = O_p(L_n)$.

$$\begin{aligned} &\tilde{\epsilon}_i^2 \tilde{\epsilon}_t^2 - \epsilon_i^2 \epsilon_t^2 \\ &= (\tilde{\epsilon}_i - \epsilon_i + \epsilon_i)^2 (\tilde{\epsilon}_t - \epsilon_t + \epsilon_t)^2 - \epsilon_i^2 \epsilon_t^2 \\ &= (\tilde{\epsilon}_i - \epsilon_i)^2 (\tilde{\epsilon}_t - \epsilon_t)^2 + (\tilde{\epsilon}_i - \epsilon_i)^2 \epsilon_t^2 + 2\epsilon_t (\tilde{\epsilon}_i - \epsilon_i)^2 (\tilde{\epsilon}_t - \epsilon_t) + \epsilon_i^2 (\tilde{\epsilon}_t - \epsilon_t)^2 + 2\epsilon_t (\tilde{\epsilon}_t - \epsilon_t) \epsilon_i^2 \\ &\quad + 2\epsilon_i (\tilde{\epsilon}_i - \epsilon_i) (\tilde{\epsilon}_t - \epsilon_t)^2 + 2\epsilon_i \epsilon_t^2 (\tilde{\epsilon}_i - \epsilon_i) + 4\epsilon_i \epsilon_t (\tilde{\epsilon}_i - \epsilon_i) (\tilde{\epsilon}_t - \epsilon_t). \end{aligned}$$

Since $\tilde{\epsilon}_i - \epsilon_i = O_p(n^{-\frac{1}{2}})$, we obtain the claim in (2).

Theorem 2: *Proof.*

1. The asymptotic local power of the test is examined with the sequence of Pitman local alternative: $H_1(l_n) : m(x) = \mu + l_n D(x)$. Since

$$\begin{aligned} &\frac{1}{n} \sum_{t=1}^n (y_t - \hat{m}(x_t))^2 = \frac{1}{n} \sum_{t=1}^n (\mu + l_n D(x_t) + \epsilon_t - \hat{m}(x_t))^2 \\ &= \frac{1}{n} \sum_{t=1}^n \epsilon_t^2 + \frac{2}{n} \sum_{t=1}^n (\mu + l_n D(x_t) - \hat{m}(x_t)) \epsilon_t + \frac{1}{n} \sum_{t=1}^n (\mu + l_n D(x_t) - \hat{m}(x_t))^2, \end{aligned}$$

we establish the following results under $H_1(l_n)$ with $l_n = \frac{1}{\sqrt{nh_n^{\frac{1}{2}}}}$:

$$\begin{aligned} (1) \quad &\frac{1}{n} \sum_{t=1}^n (y_t - \bar{y})^2 = \frac{1}{n} \sum_{t=1}^n \epsilon_t^2 + \frac{l_n^2}{n} \sum_{t=1}^n D(x_t) (D(x_t) - E(D(x_t))) + o_p((nh_n^{\frac{1}{2}})^{-1}). \\ (2) \quad &\frac{2}{n} \sum_{t=1}^n (\mu + l_n D(x_t) - \hat{m}(x_t)) \epsilon_t = o_p((nh_n^{\frac{1}{2}})^{-1}) + \frac{2}{n} \sum_{t=1}^n (\mu - \hat{m}(x_t)) \epsilon_t. \\ (3) \quad &\frac{1}{n} \sum_{t=1}^n (\mu + l_n D(x_t) - \hat{m}(x_t))^2 = o_p((nh_n^{\frac{1}{2}})^{-1}) + \frac{1}{n} \sum_{t=1}^n (\mu - \hat{m}(x_t))^2. \end{aligned}$$

Given the results in Theorem 1(a), and $\frac{1}{n} \sum_{t=1}^n D(x_t) (D(x_t) - E(D(x_t))) \xrightarrow{P} V(D(x_t))$, we obtain the claim in part 1 of Theorem 2.

(1) With result in (1) Theorem 3, we have $\bar{y}^2 = \frac{2}{n} \sum_{t=1}^n (\mu + l_n D(x_t) + \epsilon_t) (\mu + l_n E D(x_t)) - (\mu + l_n E D(x_t))^2 + O_p(n^{-1})$. So

$$\begin{aligned} &\frac{1}{n} \sum_{t=1}^n (y_t - \bar{y})^2 = \frac{1}{n} \sum_{t=1}^n (\mu + l_n D(x_t) + \epsilon_t)^2 - \bar{y}^2 \\ &= \frac{1}{n} \sum_{t=1}^n \epsilon_t^2 + \frac{2l_n}{n} \sum_{t=1}^n \epsilon_t (D(x_t) - E(D(x_t))) + \frac{l_n}{n} \sum_{t=1}^n (\mu + l_n D(x_t)) (D(x_t) - E(D(x_t))) \\ &\quad + (\mu + l_n E D(x_t)) \frac{l_n}{n} \sum_{t=1}^n (E D(x_t) - D(x_t)) + O_p(n^{-1}). \end{aligned}$$

Since $E \frac{1}{n} \sum_{t=1}^n \epsilon_t (D(x_t) - E(D(x_t))) = 0$ and $nV(\frac{1}{n} \sum_{t=1}^n \epsilon_t (D(x_t) - E(D(x_t)))) = E(\sigma^2(x_t) (D(x_t) - E(D(x_t)))^2) <$

∞ by assumption A10, $\frac{2l_n}{n} \sum_{t=1}^n \epsilon_t (D(x_t) - E(D(x_t))) = O_p((nh_n^{\frac{1}{2}})^{-\frac{1}{2}} * n^{-\frac{1}{2}}) = o_p((nh_n^{\frac{1}{2}})^{-1})$.

Similarly, $(\mu + l_n E(D(x_t))) \frac{l_n}{n} \sum_{t=1}^n (ED(x_t) - D(x_t)) = o_p((nh_n^{\frac{1}{2}})^{-1})$. So we have the claim in (1).

(2) Since $m(x_t) = \mu + l_n D(x_t)$, we follow Theorem 1(a) to have

$$\begin{aligned} \hat{m}(x_t) - m(x_t) &= \frac{1}{nh_n} \sum_{i=1}^n W_n\left(\frac{x_i - x_t}{h_n}, x_t\right) (y_i - \mu - l_n D(x_t)) \\ &= e' S_n^{-1}(x_t) \begin{bmatrix} \frac{1}{nh_n} \sum_{i=1}^n K\left(\frac{x_i - x_t}{h_n}\right) ((D(x_i) - D(x_t))l_n + \epsilon_i) \\ \frac{1}{nh_n} \sum_{i=1}^n K\left(\frac{x_i - x_t}{h_n}\right) \left(\frac{x_i - x_t}{h_n}\right) ((D(x_i) - D(x_t))l_n + \epsilon_i) \end{bmatrix} \\ &\quad - \frac{2}{n} \sum_{t=1}^n (\mu + l_n D(x_t) - \hat{m}(x_t)) \epsilon_t \\ &= -\frac{2l_n}{n^2 h_n} \sum_{t=1}^n \sum_{i=1}^n e' S_n^{-1}(x_t) K\left(\frac{x_i - x_t}{h_n}\right) (D(x_i) - D(x_t)) \epsilon_t \begin{bmatrix} 1 \\ \frac{x_i - x_t}{h_n} \end{bmatrix} \\ &\quad - \frac{2}{n^2 h_n} \sum_{t=1}^n \sum_{i=1}^n e' S_n^{-1}(x_t) K\left(\frac{x_i - x_t}{h_n}\right) \epsilon_i \epsilon_t \begin{bmatrix} 1 \\ \frac{x_i - x_t}{h_n} \end{bmatrix} \\ &= I_2 + \frac{2}{n} \sum_{t=1}^n (\mu - \hat{m}(x_t)) \epsilon_t. \end{aligned}$$

To obtain the claim in (2), we only need to show $I_2 = o_p((nh_n^{\frac{1}{2}})^{-1})$.

Since $\sup_{x_t \in G} |S_{jn}^{-1}(x_t) - S_j^{-1}(x_t)| = o_{a.s.}(1)$, we follow Theorem 1 (a) (2)(i) to have

$$I_2 = -\frac{2l_n}{n^2 h_n} \sum_{t=1}^n \sum_{i=1}^n \frac{1}{f(x_t)} K\left(\frac{x_i - x_t}{h_n}\right) (D(x_i) - D(x_t)) \epsilon_t (1 + o_{a.s.}(1)) = I_{21} (1 + o_{a.s.}(1)).$$

When $t = i$, $I_{21} = 0$. So we have

$$\begin{aligned} \frac{I_{21}}{l_n} &= -\frac{2}{n^2} \sum_{t=1}^n \sum_{i=1}^n \frac{1}{h_n f(x_t)} K\left(\frac{x_i - x_t}{h_n}\right) (D(x_i) - D(x_t)) \epsilon_t = -\frac{2}{n^2} \sum_{t=1}^n \sum_{i=1, i \neq t}^n \psi_n(Z_t, Z_i) \\ &= -\frac{1}{n^2} \sum_{t=1}^n \sum_{i=1, i \neq t}^n (\psi_n(Z_t, Z_i) + \psi_n(Z_i, Z_t)) = -\frac{1}{n^2} \sum_{t=1}^n \sum_{i=1, i \neq t}^n \phi_n(Z_t, Z_i) \\ &= -\frac{2}{n} \sum_{t=1}^n \int \phi_n(Z_t, Z_i) dP(Z_i) + O_p(n^{-1} (E\phi_n^2(Z_t, Z_i))^{\frac{1}{2}}) \end{aligned}$$

by Lemma 1, since $E\phi_n(Z_t, Z_i) = 0$.

$$\int \phi_n(Z_t, Z_i) dP(Z_i) = \frac{\epsilon_t}{f(x_t)} \frac{1}{h_n} \int K\left(\frac{x_i - x_t}{h_n}\right) (D(x_i) - D(x_t)) f(x_i) dx_i,$$

$$E\left(\frac{1}{n} \sum_{t=1}^n \phi_n(Z_t, Z_i) dP(Z_i)\right) = 0,$$

$$V\left(\frac{1}{n} \sum_{t=1}^n \int \phi_n(Z_t, Z_i) dP(Z_i)\right) = \frac{1}{n} \int \frac{\sigma^2(x_t)}{f^2(x_t)} [\int K(\psi) (D(x_t + h_n \psi) - D(x_t)) f(x_t + h_n \psi) d\psi]^2 dx_t = o(n^{-1}).$$

$$h_n E\phi_n^2(Z_t, Z_i) \leq 4h_n E\psi_n^2(Z_t, Z_i) = \int \frac{K^2(\psi)}{f^2(x_t)} (D(x_t + h_n \psi) - D(x_t))^2 \sigma^2(x_t) f(x_t) f(x_t + h_n \psi) d\psi dx_t \rightarrow 0.$$

So $I_{21} = o_p(n^{-\frac{1}{2}} \frac{1}{\sqrt{nh_n^{\frac{1}{2}}}}) = o_p((nh_n^{\frac{1}{2}})^{-1})$ and $I_2 = o_p((nh_n^{\frac{1}{2}})^{-1})$.

$$\begin{aligned} (3) \quad &\frac{1}{n} \sum_{t=1}^n (\mu + l_n D(x_t) - \hat{m}(x_t))^2 \\ &= \frac{1}{n} \sum_{t=1}^n \left[e' S_n^{-1}(x_t) \frac{1}{nh_n} \sum_{i=1}^n K\left(\frac{x_i - x_t}{h_n}\right) (l_n (D(x_i) - D(x_t)) + \epsilon_i) \begin{pmatrix} 1 \\ \frac{x_i - x_t}{h_n} \end{pmatrix} \right]^2 \\ &= \frac{1}{n} \sum_{t=1}^n (\mu - \hat{m}(x_t))^2 + \left\{ \frac{1}{n^3 h_n^2} \sum_{t=1}^n \sum_{i=1}^n \sum_{j=1}^n \frac{1}{f^2(x_t)} K\left(\frac{x_i - x_t}{h_n}\right) K\left(\frac{x_j - x_t}{h_n}\right) [l_n^2 (D(x_i) - D(x_t)) (D(x_j) - D(x_t))] \right. \\ &\quad + \frac{1}{n^3 h_n^2} \sum_{t=1}^n \sum_{i=1}^n \sum_{j=1}^n \frac{1}{f^2(x_t)} K\left(\frac{x_i - x_t}{h_n}\right) K\left(\frac{x_j - x_t}{h_n}\right) l_n (D(x_i) - D(x_t)) \epsilon_j \\ &\quad \left. + \frac{1}{n^3 h_n^2} \sum_{t=1}^n \sum_{i=1}^n \sum_{j=1}^n \frac{1}{f^2(x_t)} K\left(\frac{x_i - x_t}{h_n}\right) K\left(\frac{x_j - x_t}{h_n}\right) l_n (D(x_j) - D(x_t)) \epsilon_i \right\} (1 + o_{a.s.}(1)) \\ &= \frac{1}{n} \sum_{t=1}^n (\mu - \hat{m}(x_t))^2 + I_{31} + I_{32} + I_{33}. \end{aligned}$$

Since I_{33} is similar to I_{32} , we show (i) $I_{31} = o_p((nh_n^{\frac{1}{2}})^{-1})$, and (ii) $I_{32} = o_p((nh_n^{\frac{1}{2}})^{-1})$, which are sufficient for the claim in (3).

(i) We note when $t = i = j$, $I_{31} = 0$. Similarly, when $t = i$ or $t = j$, $I_{31} = 0$. So we consider the case when $i = j$, but $t \neq i$. Since $\frac{1}{t_n^2} = \frac{1}{nh_n^{\frac{1}{2}}}$, we only need to show $\frac{I_{31}}{t_n^2} = o_p(1)$.

$$\begin{aligned}\frac{I_{31}}{t_n^2} &= \frac{1}{nh_n} \frac{1}{n^2 h_n} \sum_{t=1}^n \sum_{i=1}^n \frac{1}{f^2(x_t)} K^2\left(\frac{x_i - x_t}{h_n}\right) (D(x_i) - D(x_t))^2 = \frac{1}{nh_n} I_{311}. \\ I_{311} &= \frac{1}{n^2} \sum_{t=1}^n \sum_{i=1}^n \underbrace{\frac{1}{h_n f^2(x_t)} K^2\left(\frac{x_i - x_t}{h_n}\right) (D(x_i) - D(x_t))^2}_{\psi_n(x_i, x_t)} \\ &= \frac{1}{2n^2} \sum_{t=1}^n \sum_{i=1}^n (\psi_n(x_t, x_i) + \psi_n(x_i, x_t)) = \frac{1}{2n^2} \sum_{t=1}^n \sum_{i=1}^n \phi_n(x_t, x_i).\end{aligned}$$

Since $\int \phi_n(x_t, x_i) dP(x_i) = \int \frac{1}{f^2(x_t)} K^2(\psi)(D(x_t + h_n \psi) - D(x_t))^2 f(x_t + h_n \psi) d\psi + \int \frac{1}{f(x_t + h_n \psi)} K^2(\psi)(D(x_t) - D(x_t + h_n \psi))^2 d\psi$, $E \int \phi_n(x_t, x_i) dP(x_i) = E \phi_n(x_t, x_i) \rightarrow 0$.

$h_n E \phi_n^2(x_t, x_i) \leq 4h_n E \psi_n^2(x_t, x_i) = 4 \frac{1}{h_n} E \frac{1}{f^4(x_t)} K^4\left(\frac{x_i - x_t}{h_n}\right) (D(x_i) - D(x_t))^4 \rightarrow 0$, $O_p(n^{-1}(E \phi_n^2(x_t, x_i))^{\frac{1}{2}}) = o_p((nh_n^{\frac{1}{2}})^{-1})$, and $V(\phi_n(x_t, x_i)) = o(h_n^{-1})$, so by Lemma 1, we have

$$I_{311} = \frac{1}{n} \sum_{t=1}^n \int \phi_n(x_t, x_i) dP(x_i) - \frac{1}{2} E \phi_n(x_t, x_i) + o_p((nh_n^{\frac{1}{2}})^{-1}) = o_p(1).$$

So we have the claimed result that $I_{31} = o_p((nh_n^{\frac{1}{2}})^{-1})$.

When t, i, j are three distinct indices,

$$\begin{aligned}\frac{I_{31}}{t_n^2} &= \frac{1}{n^3} \sum_{t=1}^n \sum_{i=1}^n \sum_{j=1}^n \underbrace{\frac{1}{h_n^2 f^2(x_t)} K\left(\frac{x_i - x_t}{h_n}\right) K\left(\frac{x_j - x_t}{h_n}\right) (D(x_i) - D(x_t))(D(x_j) - D(x_t))}_{\psi_n(x_t, x_i, x_j)} \\ &= \frac{1}{3n^3} \sum_{t=1}^n \sum_{i=1}^n \sum_{j=1}^n \underbrace{[\psi_n(x_t, x_i, x_j) + \psi_n(x_i, x_t, x_j) + \psi_n(x_j, x_i, x_t)]}_{\phi_n(x_t, x_i, x_j)} = \frac{1}{n^3} \sum_{1=t < i < j=n} 2\phi_n(x_t, x_i, x_j) \\ &= \frac{1}{6} \binom{n}{3} \sum_{1=t < i < j=n} 2\phi_n(x_t, x_i, x_j) (1 + O(n^{-4})) = \frac{1}{6} I_{311} (1 + O(n^{-4})).\end{aligned}$$

I_{311} is a U-statistic. We consider the H-decomposition in Lemma 3 and write

$$I_{311} = \frac{3}{n} \sum_{t=1}^n \int 2\phi_n(x_t, x_i, x_j) dP(x_i) dP(x_j) - 2\theta_n + O_p(H_n^{(2)} + H_n^{(3)}).$$

$\theta_n = 6E\psi_n(x_t, x_i, x_j) = 6 \int \frac{1}{f^4(x_t)} K(\psi_1) K(\psi_2) (D(x_t + h_n \psi_1) - D(x_t))(D(x_t + h_n \psi_2) - D(x_t)) f(x_t + h_n \psi_1) f(x_t + h_n \psi_2) dx_t d\psi_1 d\psi_2 \rightarrow 0$.

$E[\int 2\phi_n(x_t, x_i, x_j) dP(x_i) dP(x_j)] = \theta_n \rightarrow 0$. Furthermore,

$$\begin{aligned}h_n^2 E(\int 2\phi_n(x_t, x_i, x_j) dP(x_i) dP(x_j))^2 &\leq 4h_n^2 E(\phi_n^2(x_t, x_i, x_j)) \leq 4ch_n^2 E(\psi_n^2(x_t, x_i, x_j)) \\ &= 4c \int \frac{1}{f^4(x_t)} K^2(\psi_1) K^2(\psi_2) (D(x_t + h_n \psi_1) - D(x_t))^2 (D(x_t + h_n \psi_2) - D(x_t))^2 \\ &\quad \times f(x_t) f(x_t + h_n \psi_1) f(x_t + h_n \psi_2) dx_t d\psi_1 d\psi_2 \\ &\rightarrow 0.\end{aligned}$$

$V(\int 2\phi_n(x_t, x_i, x_j) dP(x_i) dP(x_j)) = o(h_n^{-2})$. Since $nh_n^2 \rightarrow \infty$, we have

$$\frac{3}{n} \sum_{t=1}^n \int 2\phi_n(x_t, x_i, x_j) dP(x_i) dP(x_j) - 2\theta_n = o_p(1).$$

$V(H_n^{(2)}) = O(n^{-2} \sigma_{2n}^2)$, $V(H_n^{(3)}) = O(n^{-3} \sigma_{3n}^2)$. $\sigma_{2n}^2 = V(\int 2\phi_n(x_t, x_i, x_j) dP(x_j)) \leq cE\phi_n^2(x_t, x_i, x_j)$, $\sigma_{3n}^2 \leq cE\phi_n^2(x_t, x_i, x_j) = o(h_n^{-2})$, $V(H_n^{(2)}) = o(n^{-2} h_n^{-2})$, and $V(H_n^{(3)}) = o(n^{-3} h_n^{-2})$, so $O_p(H_n^{(2)} + H_n^{(3)}) = o_p((nh_n)^{-1})$. So we conclude $I_{311} = o_p(1)$ and $I_{31} = o_p((nh_n^{\frac{1}{2}})^{-1})$.

(ii) With similar arguments as (i), we apply Lemma 1 and 3 to obtain $I_{32} = o_p((nh_n^{\frac{1}{2}})^{-1})$.

With results in (i) and (ii) above, we obtain the claim in (3) and the claim in part 1 of Theorem 2.

2. We follow the proof in Theorem 1(b) closely. It is sufficient to prove the following results:

(i) $\hat{A}_n - A_n = o_p((nh_n^{\frac{1}{2}})^{-1})$. (ii) $\hat{A}_{1n} - A_{1n} = o_p((nh_n^{\frac{1}{2}})^{-1})$. (iii) $\hat{V}_T - V_T = o_p(1)$.

Note under $H_1(l_n) : y_t = \mu + l_n D(x_t) + \epsilon_t$, so

$$\begin{aligned}\tilde{\epsilon}_t - \epsilon_t &= (y_t - \bar{y}) - (y_t - \mu - l_n D(x_t)) \\ &= \mu + l_n ED(x_t) - \bar{y} + l_n (D(x_t) - ED(x_t)) = O_p(n^{-\frac{1}{2}}) + l_n (D(x_t) - ED(x_t)).\end{aligned}$$

(i) As show in Theorem 1(b),

$$\begin{aligned}\hat{A}_n - A_n &= \frac{1}{n^3 h_n^2} \sum_{t=1}^n \sum_{\substack{i=1 \\ t \neq i}}^n K^2\left(\frac{x_i - x_t}{h_n}\right) \left[\frac{1}{f^2(x_t)} - \frac{1}{f^2(x_i)} \right] [\tilde{\epsilon}_i^2 - \epsilon_i^2 + \epsilon_i^2] + \frac{1}{n^3 h_n^2} \sum_{t=1}^n \sum_{\substack{i=1 \\ t \neq i}}^n K^2\left(\frac{x_i - x_t}{h_n}\right) \frac{1}{f^2(x_t)} [\tilde{\epsilon}_i^2 - \epsilon_i^2] \\ &= I_1 + I_2 \\ I_1 &\leq O_p(L_n) \left\{ \frac{1}{n^3 h_n^2} \sum_{t=1}^n \sum_{\substack{i=1 \\ t \neq i}}^n K^2\left(\frac{x_i - x_t}{h_n}\right) \epsilon_i^2 + \frac{1}{n^3 h_n^2} \sum_{t=1}^n \sum_{\substack{i=1 \\ t \neq i}}^n K^2\left(\frac{x_i - x_t}{h_n}\right) |(\tilde{\epsilon}_i + \epsilon_i)(\tilde{\epsilon}_i - \epsilon_i)| \right\}.\end{aligned}$$

With Lemma 1, we easily obtain $\frac{1}{n^3 h_n^2} \sum_{t=1}^n \sum_{\substack{i=1 \\ t \neq i}}^n K^2\left(\frac{x_i - x_t}{h_n}\right) \epsilon_i^2 = O_p((nh_n)^{-1})$.

$$\begin{aligned}& \frac{1}{n^3 h_n^2} \sum_{t=1}^n \sum_{\substack{i=1 \\ t \neq i}}^n K^2\left(\frac{x_i - x_t}{h_n}\right) |(\tilde{\epsilon}_i - \epsilon_i + 2\epsilon_i)(\tilde{\epsilon}_i - \epsilon_i)| \\ & \leq \frac{1}{n^3 h_n^2} \sum_{t=1}^n \sum_{\substack{i=1 \\ t \neq i}}^n K^2\left(\frac{x_i - x_t}{h_n}\right) (\tilde{\epsilon}_i - \epsilon_i)^2 + \frac{2}{n^3 h_n^2} \sum_{t=1}^n \sum_{\substack{i=1 \\ t \neq i}}^n K^2\left(\frac{x_i - x_t}{h_n}\right) |(\tilde{\epsilon}_i - \epsilon_i)\epsilon_i| \\ & = I_{11} + I_{12}. \\ I_{11} &= O_p(n^{-1}) \frac{1}{n^3 h_n^2} \sum_{t=1}^n \sum_{\substack{i=1 \\ t \neq i}}^n K^2\left(\frac{x_i - x_t}{h_n}\right) + \frac{l_n^2}{n^3 h_n^2} \sum_{t=1}^n \sum_{\substack{i=1 \\ t \neq i}}^n K^2\left(\frac{x_i - x_t}{h_n}\right) (D(x_i) - ED(x_i))^2 \\ & \quad + O_p(n^{-\frac{1}{2}}) \frac{2l_n}{n^3 h_n^2} \sum_{t=1}^n \sum_{\substack{i=1 \\ t \neq i}}^n K^2\left(\frac{x_i - x_t}{h_n}\right) (D(x_i) - ED(x_i)) \\ & = I_{111} + I_{112} + I_{113}\end{aligned}$$

We can show that $I_{111} = O_p(n^{-2}h_n^{-1})$, $I_{112} = O_p(n^{-2}h_n^{-\frac{3}{2}})$, and $I_{113} = O_p(n^{-2}h_n^{-\frac{5}{4}})$. So $I_{11} = O_p(n^{-2}h_n^{-\frac{5}{4}}) = o_p((nh_n^{\frac{1}{2}})^{-1})$.

$$\begin{aligned}I_{12} &= O_p(n^{-\frac{1}{2}}) \frac{2}{n^3 h_n^2} \sum_{t=1}^n \sum_{\substack{i=1 \\ t \neq i}}^n K^2\left(\frac{x_i - x_t}{h_n}\right) |\epsilon_i| + \frac{2l_n}{n^3 h_n^2} \sum_{t=1}^n \sum_{\substack{i=1 \\ t \neq i}}^n K^2\left(\frac{x_i - x_t}{h_n}\right) |(D(x_i) - ED(x_i))\epsilon_i| \\ & = I_{121} + I_{122}.\end{aligned}$$

We can show that $I_{121} = O_p(n^{-\frac{3}{2}}h_n^{-1})$ and $I_{122} = O_p(n^{-\frac{3}{2}}h_n^{-\frac{5}{4}})$. So $I_{12} = O_p(n^{-\frac{3}{2}}h_n^{-\frac{5}{4}})$.

So $I_1 = O_p(L_n)O_p((nh_n)^{-1}) = o_p((nh_n^{\frac{1}{2}})^{-1})$.

With similar arguments, we have $I_2 = o_p((nh_n^{\frac{1}{2}})^{-1})$. Thus, $\hat{A}_n - A_n = o_p((nh_n^{\frac{1}{2}})^{-1})$.

The proof of (ii) and (iii) follows from Theorem 1(b) in an analogous fashion.

Theorem 4: Proof. Recall the definition of $\hat{T}_n$ in section 2,

$$\hat{T}_n = \frac{n h_n^{\frac{1}{2}} \{ \hat{R}^2 + I(\frac{1}{n} \sum_{t=1}^n (y_t - \bar{y})^2) \geq \frac{1}{n} \sum_{t=1}^n (y_t - \hat{m}(x_t))^2 \} \frac{\hat{A}_{1n} + \hat{A}_n}{\frac{1}{n} \sum_{t=1}^n (y_t - \bar{y})^2}}{\sqrt{\hat{V}_T}}.$$

From Theorem 3, $\hat{R}^2 \xrightarrow{P} R^2 > 0$ under H_1 . From Theorem 3 proof (4), we know $I(\frac{1}{n} \sum_{t=1}^n (y_t - \bar{y})^2) \geq \frac{1}{n} \sum_{t=1}^n (y_t - \hat{m}(x_t))^2 \xrightarrow{P} 1$. It is also shown that $\frac{1}{n} \sum_{t=1}^n (y_t - \bar{y})^2 \xrightarrow{P} E\sigma^2(x_t) > 0$.

Under H_1 , $y_t = m(x_t) + \epsilon_t$, so

$$\begin{aligned}\tilde{\epsilon}_t - \epsilon_t &= (y_t - \bar{y}) - (y_t - m(x_t)) = m(x_t) - E(m(x_t)) + (E(m(x_t)) - \bar{y}) \\ &= m(x_t) - E(m(x_t)) + O_p(n^{-\frac{1}{2}}).\end{aligned}$$

Following the proof in Theorem 1(b), it is easy to see that

$$\hat{A}_n = \frac{1}{n^3 h_n^2} \sum_{t=1}^n \sum_{\substack{i=1 \\ t \neq i}}^n \frac{1}{f^2(x_t)} K^2\left(\frac{x_i - x_t}{h_n}\right) \tilde{\epsilon}_i^2 = O_p((nh_n)^{-1}), \quad \hat{A}_{1n} = -\frac{2}{n^2 h_n} \sum_{t=1}^n \frac{K(0)}{f(x_t)} \tilde{\epsilon}_t^2 = O_p((nh_n)^{-1}).$$

So to prove Theorem 4, we only need to show $\hat{V}_T$'s probability limit is positive, since

$\hat{T}_n = \{nh_n^{\frac{1}{2}}[R^2 + O_p((nh_n)^{-1})]\}O_p(1) \rightarrow \infty$ at rate $nh_n^{\frac{1}{2}}$, as $R^2 > 0$ under H_1 . So $P(\hat{T}_n > c_n) \rightarrow 1$ for any positive constant $c_n = o(nh_n^{\frac{1}{2}})$. Thus the $\hat{T}_n$ test is consistent.

It is sufficient to show

$$\begin{aligned} & \frac{1}{n^2} \sum_{t=1}^n \sum_{i=1}^n \frac{1}{h_n f^2(x_t)} K\left(\frac{x_i - x_t}{h_n}\right) \tilde{\epsilon}_i^2 \tilde{\epsilon}_t^2 \\ & \xrightarrow{p} E\left(\frac{\sigma^4(x_t)}{f(x_t)}\right) + E\left(\frac{(m(x_t) - E(m(x_t)))^4}{f(x_t)}\right) + 2E\left(\frac{(m(x_t) - E(m(x_t)))^2 \sigma^2(x_t)}{f(x_t)}\right), \text{ which is positive.} \end{aligned}$$

Theorem 1(b) (iii)(1) shows $\frac{1}{n^2} \sum_{t=1}^n \sum_{i=1}^n \frac{1}{h_n f^2(x_t)} K\left(\frac{x_i - x_t}{h_n}\right) \epsilon_i^2 \epsilon_t^2 \xrightarrow{p} E\left(\frac{\sigma^4(x_t)}{f(x_t)}\right)$. So we only need to show

$$\begin{aligned} & \frac{1}{n^2} \sum_{t=1}^n \sum_{i=1}^n \frac{1}{h_n f^2(x_t)} K\left(\frac{x_i - x_t}{h_n}\right) \tilde{\epsilon}_i^2 \tilde{\epsilon}_t^2 - \frac{1}{n^2} \sum_{t=1}^n \sum_{i=1}^n \frac{1}{h_n f^2(x_t)} K\left(\frac{x_i - x_t}{h_n}\right) \epsilon_i^2 \epsilon_t^2 \\ & = o_p(1) + E\left(\frac{(m(x_t) - E(m(x_t)))^4}{f(x_t)}\right) + 2E\left(\frac{(m(x_t) - E(m(x_t)))^2 \sigma^2(x_t)}{f(x_t)}\right). \\ & \frac{1}{n^2} \sum_{t=1}^n \sum_{i=1}^n K\left(\frac{x_i - x_t}{h_n}\right) \frac{\tilde{\epsilon}_i^2 \tilde{\epsilon}_t^2}{h_n f^2(x_t)} - \frac{1}{n^2} \sum_{t=1}^n \sum_{i=1}^n K\left(\frac{x_i - x_t}{h_n}\right) \frac{\epsilon_i^2 \epsilon_t^2}{h_n f^2(x_t)} \\ & = \frac{1}{n^2} \sum_{t=1}^n \sum_{i=1}^n \frac{1}{h_n} \left[\frac{1}{f^2(x_t)} - \frac{1}{f^2(x_t)}\right] K\left(\frac{x_i - x_t}{h_n}\right) \tilde{\epsilon}_i^2 \tilde{\epsilon}_t^2 + \frac{1}{n^2} \sum_{t=1}^n \sum_{i=1}^n \frac{1}{h_n} \frac{1}{f^2(x_t)} K\left(\frac{x_i - x_t}{h_n}\right) (\tilde{\epsilon}_i^2 \tilde{\epsilon}_t^2 - \epsilon_i^2 \epsilon_t^2). \end{aligned}$$

Note $\sup_{x_t \in G} \left| \frac{1}{f(x_t)} - \frac{1}{f(x_t)} \right| = O_p(L_n)$, which implies

$\frac{1}{n^2} \sum_{t=1}^n \sum_{i=1}^n \frac{1}{h_n} \left[\frac{1}{f^2(x_t)} - \frac{1}{f^2(x_t)}\right] K\left(\frac{x_i - x_t}{h_n}\right) \tilde{\epsilon}_i^2 \tilde{\epsilon}_t^2 = o_p(1)$. Following Theorem 1(b)(iii)(2), we have

$$\begin{aligned} & \frac{1}{n^2} \sum_{t=1}^n \sum_{i=1}^n \frac{1}{h_n} \frac{1}{f^2(x_t)} K\left(\frac{x_i - x_t}{h_n}\right) (\tilde{\epsilon}_i^2 \tilde{\epsilon}_t^2 - \epsilon_i^2 \epsilon_t^2) \\ & = \frac{1}{n^2} \sum_{t=1}^n \sum_{i=1}^n \frac{1}{h_n} \frac{1}{f^2(x_t)} K\left(\frac{x_i - x_t}{h_n}\right) [(\tilde{\epsilon}_i - \epsilon_i)^2 (\tilde{\epsilon}_t - \epsilon_t)^2 + (\tilde{\epsilon}_i - \epsilon_i)^2 \epsilon_t^2 + 2\epsilon_t (\tilde{\epsilon}_i - \epsilon_i)^2 (\tilde{\epsilon}_t - \epsilon_t) \\ & \quad + \epsilon_i^2 (\tilde{\epsilon}_t - \epsilon_t)^2 + 2\epsilon_t (\tilde{\epsilon}_t - \epsilon_t) \epsilon_i^2 + 2\epsilon_i (\tilde{\epsilon}_i - \epsilon_i) (\tilde{\epsilon}_t - \epsilon_t)^2 + 2\epsilon_i \epsilon_t^2 (\tilde{\epsilon}_i - \epsilon_i) + 4\epsilon_i \epsilon_t (\tilde{\epsilon}_i - \epsilon_i) (\tilde{\epsilon}_t - \epsilon_t)] \\ & = \frac{1}{n^2} \sum_{t=1}^n \sum_{i=1}^n \frac{1}{h_n} \frac{1}{f^2(x_t)} K\left(\frac{x_i - x_t}{h_n}\right) (\tilde{\epsilon}_i - \epsilon_i)^2 (\tilde{\epsilon}_t - \epsilon_t)^2 + \frac{1}{n^2} \sum_{t=1}^n \sum_{i=1}^n \frac{1}{h_n} \frac{1}{f^2(x_t)} K\left(\frac{x_i - x_t}{h_n}\right) \epsilon_t^2 (\tilde{\epsilon}_i - \epsilon_i)^2 \\ & \quad + \frac{1}{n^2} \sum_{t=1}^n \sum_{i=1}^n \frac{1}{h_n} \frac{1}{f^2(x_t)} K\left(\frac{x_i - x_t}{h_n}\right) \epsilon_i^2 (\tilde{\epsilon}_t - \epsilon_t)^2 + o_p(1), \end{aligned}$$

with Lemma 1 and $\tilde{\epsilon}_t - \epsilon_t = m(x_t) - E(m(x_t)) + O_p(n^{-\frac{1}{2}})$. Consider

$$\begin{aligned} & \frac{1}{n^2} \sum_{t=1}^n \sum_{i=1}^n \frac{1}{h_n} \frac{1}{f^2(x_t)} K\left(\frac{x_i - x_t}{h_n}\right) (\tilde{\epsilon}_i - \epsilon_i)^2 (\tilde{\epsilon}_t - \epsilon_t)^2 \\ & = O_p(n^{-\frac{1}{2}}) + \frac{1}{n^2} \sum_{t=1}^n \sum_{i=1}^n \frac{1}{h_n} \frac{1}{f^2(x_t)} K\left(\frac{x_i - x_t}{h_n}\right) (m(x_i) - E(m(x_i)))^2 (m(x_t) - E(m(x_t)))^2 \\ & = O_p(n^{-\frac{1}{2}}) + \frac{1}{n^2} \sum_{t=1}^n \sum_{i=1}^n \psi_n(x_t, x_i) = O_p(n^{-\frac{1}{2}}) + I. \end{aligned}$$

By Lemma 1, we have

$$\begin{aligned} I & = \frac{1}{2n^2} \sum_{t=1}^n \sum_{i=1}^n (\psi_n(x_t, x_i) + \psi_n(x_i, x_t)) = \frac{1}{2n^2} \sum_{t=1}^n \sum_{i=1}^n \phi_n(x_t, x_i) \\ & = \frac{1}{n} \sum_{t=1}^n \int \phi_n(x_t, x_i) dP(x_i) - \frac{1}{2} E\phi_n(x_t, x_i) + O_p(n^{-1} (E\phi_n^2(x_t, x_i))^{\frac{1}{2}}) \end{aligned}$$

$$\begin{aligned} E(\int \phi_n(x_t, x_i) dP(x_i)) & = E(\phi_n(x_t, x_i)) \rightarrow 2E\left(\frac{(m(x_t) - E(m(x_t)))^4}{f(x_t)}\right), \\ h_n E(\phi_n^2(x_t, x_i)) & \leq 4h_n E(\psi_n^2(x_t, x_i)) \rightarrow 4 \int K^2(\psi) d\psi E\left(\frac{(m(x_t) - E(m(x_t)))^8}{f^3(x_t)}\right) < \infty, \text{ so} \end{aligned}$$

$O_p(n^{-1}(E\phi_n^2(x_t, x_i))^{\frac{1}{2}}) = O_p(n^{-1}h_n^{-\frac{1}{2}})$. Since $V(\frac{1}{n}\sum_{t=1}^n \int \phi_n(x_t, x_i)dP(x_i)) = O((nh_n)^{-1})$, we conclude $I \xrightarrow{p} E\left(\frac{(m(x_t)-E(m(x_t)))^4}{f(x_t)}\right)$. So

$$\frac{1}{n^2} \sum_{t=1}^n \sum_{\substack{i=1 \\ t \neq i}}^n \frac{1}{h_n} \frac{1}{f^2(x_t)} K\left(\frac{x_i - x_t}{h_n}\right) (\tilde{\epsilon}_i - \epsilon_i)^2 (\tilde{\epsilon}_t - \epsilon_t)^2 = o_p(1) + E\left(\frac{(m(x_t)-E(m(x_t)))^4}{f(x_t)}\right).$$

Similarly, we show $\frac{1}{n^2} \sum_{t=1}^n \sum_{\substack{i=1 \\ t \neq i}}^n \frac{1}{h_n} \frac{1}{f^2(x_t)} K\left(\frac{x_i - x_t}{h_n}\right) \epsilon_t^2 (\tilde{\epsilon}_i - \epsilon_i)^2 = o_p(1) + E\left(\frac{(m(x_t)-E(m(x_t)))^2 \sigma^2(x_t)}{f(x_t)}\right)$,

$$\frac{1}{n^2} \sum_{t=1}^n \sum_{\substack{i=1 \\ t \neq i}}^n \frac{1}{h_n} \frac{1}{f^2(x_t)} K\left(\frac{x_i - x_t}{h_n}\right) \epsilon_i^2 (\tilde{\epsilon}_t - \epsilon_t)^2 = o_p(1) + E\left(\frac{(m(x_t)-E(m(x_t)))^2 \sigma^2(x_t)}{f(x_t)}\right).$$

Combing above results, we obtain the claim in Theorem 4.

Theorem 5: *Proof.* Recall the definition of

$$\hat{T}_n^* = \frac{n h_n^{\frac{1}{2}} \{ \hat{R}^{2*} - I(\frac{1}{n} \sum_{t=1}^n (\epsilon_{t,0}^*)^2 \geq \frac{1}{n} \sum_{t=1}^n (\epsilon_{t,b}^*)^2) - \frac{A_{1n}^* + A_{2n}^*}{\frac{1}{n} \sum_{t=1}^n (\epsilon_{t,0}^*)^2} \}}{\sqrt{\hat{V}_T^*}}, \text{ for } \hat{R}^{2*} = \underbrace{\left[1 - \frac{\frac{1}{n} \sum_{t=1}^n (\epsilon_{t,b}^*)^2}{\frac{1}{n} \sum_{t=1}^n (\epsilon_{t,0}^*)^2} \right]}_{\hat{R}^{2*}} I(\frac{1}{n} \sum_{t=1}^n (\epsilon_{t,0}^*)^2 \geq \frac{1}{n} \sum_{t=1}^n (\epsilon_{t,b}^*)^2).$$

To prove the Theorem, it is sufficient to show that given $W = \{x_i, y_i\}_{i=1}^n$,

$$(1) \frac{1}{n} \sum_{t=1}^n (\epsilon_{t,0}^*)^2 = \frac{1}{n} \sum_{t=1}^n (\epsilon_t^*)^2 + O_p(n^{-1}).$$

$$\text{Since } \frac{1}{n} \sum_{t=1}^n (\epsilon_{t,b}^*)^2 = \frac{1}{n} \sum_{t=1}^n (y_t^* - \hat{m}^*(x_t))^2 = \frac{1}{n} \sum_{t=1}^n (\epsilon_t^*)^2 + \frac{2}{n} \sum_{t=1}^n (\bar{y} - \hat{m}^*(x_t)) \epsilon_t^* + \frac{1}{n} \sum_{t=1}^n (\bar{y} - \hat{m}^*(x_t))^2,$$

$$(2) \text{ we show } \frac{2}{n} \sum_{t=1}^n (\bar{y} - \hat{m}^*(x_t)) \epsilon_t^* = (A_{1n}^* + A_{2n}^*)(1 + o_p(1)), \text{ where } A_{1n}^* = -\frac{2}{n^2 h_n} \sum_{t=1}^n \frac{K(0)}{f(x_t)} (\epsilon_t^*)^2, \text{ and}$$

$$A_{2n}^* = -\frac{2}{n^2 h_n} \sum_{t=1}^n \sum_{\substack{i=1 \\ t \neq i}}^n \frac{1}{f(x_t)} K\left(\frac{x_i - x_t}{h_n}\right) \epsilon_i^* \epsilon_t^*.$$

$$(3) \frac{1}{n} \sum_{t=1}^n (\bar{y} - \hat{m}^*(x_t))^2 - A_n^*(1 + o_p(1)) = A_{3n}^*(1 + o_p(1)) + o_p(n^{-1}), A_n^* = \frac{1}{n^3 h_n^2} \sum_{t=1}^n \sum_{\substack{i=1 \\ t \neq i}}^n \frac{1}{f^2(x_t)} K^2\left(\frac{x_i - x_t}{h_n}\right) (\epsilon_i^*)^2,$$

$$A_{3n}^* = \frac{1}{n^3 h_n^2} \sum_{t=1}^n \sum_{\substack{i=1 \\ t \neq i}}^n \sum_{\substack{j=1 \\ t < i}}^n \frac{1}{f^2(x_t)} K\left(\frac{x_i - x_t}{h_n}\right) K\left(\frac{x_j - x_t}{h_n}\right) \epsilon_i^* \epsilon_j^* = \frac{1}{n(n-1)} \sum_{t=1}^n \sum_{\substack{i=1 \\ t < i}}^n \left[\frac{\epsilon_i^* \epsilon_t^*}{h_n^2} E(K\left(\frac{x_i - x_i}{h_n}\right) K\left(\frac{x_t - x_i}{h_n}\right) \frac{1}{f^2(x_j)} | x_t, x_i) \right. \\ \left. + \frac{\epsilon_t^* \epsilon_i^*}{h_n^2} E(K\left(\frac{x_t - x_j}{h_n}\right) K\left(\frac{x_i - x_j}{h_n}\right) \frac{1}{f^2(x_j)} | x_i, x_t) \right] + o_p(n^{-1}).$$

Results (2) and (3) above imply

$$\frac{1}{n} \sum_{t=1}^n (\epsilon_{t,b}^*)^2 - (A_n^* + A_{1n}^*)(1 + o_p(1)) = \frac{1}{n} \sum_{t=1}^n (\epsilon_t^*)^2 + (A_{2n}^* + A_{3n}^*)(1 + o_p(1)) + o_p(n^{-1}).$$

$$(4) (A_{2n}^* + A_{3n}^*)(S_n^*)^{-1} \xrightarrow{d} N(0, 1), \text{ where}$$

$$(S_n^*)^2 = E\left(\left(\sum_{t=1}^n \sum_{\substack{i=1 \\ t < i}}^n \phi_n(Z_t^*, Z_i^*)\right)^2 | W\right) = n^{-2} h_n^{-1} \sigma_\phi^2 + o_p(n^{-2} h_n^{-1}), \text{ and}$$

$$\begin{aligned} & \phi_n(Z_t^*, Z_i^*) \\ &= \frac{1}{n^2 h_n} \left[-\frac{2}{f(x_t)} K\left(\frac{x_i - x_t}{h_n}\right) \epsilon_i^* \epsilon_t^* - \frac{2}{f(x_i)} K\left(\frac{x_t - x_i}{h_n}\right) \epsilon_t^* \epsilon_i^* \right. \\ & \quad \left. + \epsilon_i^* \epsilon_t^* E\left(\frac{1}{h_n f^2(x_j)} K\left(\frac{x_i - x_j}{h_n}\right) K\left(\frac{x_t - x_j}{h_n}\right) | x_t, x_i\right) + \epsilon_t^* \epsilon_i^* E\left(\frac{1}{h_n f^2(x_j)} K\left(\frac{x_t - x_j}{h_n}\right) K\left(\frac{x_i - x_j}{h_n}\right) | x_t, x_i\right) \right]. \end{aligned}$$

Combining results (1)-(4), we conclude

$$\frac{n h_n^{\frac{1}{2}}}{n h_n^{\frac{1}{2}} S_n^*} \left\{ \frac{1}{n} \sum_{t=1}^n (\epsilon_{t,0}^*)^2 - \frac{1}{n} \sum_{t=1}^n (\epsilon_{t,b}^*)^2 + (A_n^* + A_{1n}^*)(1 + o_p(1)) \right\} \xrightarrow{d} N(0, 1).$$

Since $\frac{1}{n} \sum_{t=1}^n (\epsilon_{t,0}^*)^2 \xrightarrow{p} E\sigma^2(x_t) > 0$, we have

$$\frac{n h_n^{\frac{1}{2}}}{n h_n^{\frac{1}{2}} \frac{S_n^*}{E(\sigma^2(x_t))}} \left\{ \hat{R}^{2*} + \left[\frac{1}{n} \sum_{t=1}^n (\epsilon_{t,0}^*)^2 \right]^{-1} ((A_n^* + A_{1n}^*)(1 + o_p(1))) \right\} \xrightarrow{d} N(0, 1).$$

We follow the proof in Theorem 1(a) to show $I(\cdot) = I(\frac{1}{n} \sum_{t=1}^n (\epsilon_{t,0}^*)^2 \geq \frac{1}{n} \sum_{t=1}^n (\epsilon_{t,b}^*)^2) \xrightarrow{p} 1$. So

$$\frac{nh_n^{\frac{1}{2}}}{nh_n^{\frac{1}{2}} \frac{S_n^*}{E(\sigma^2(x_t))}} \{ \hat{R}^{2*} + I(\cdot) [\frac{1}{n} \sum_{t=1}^n (\epsilon_{t,0}^*)^2]^{-1} ((A_n^* + A_{1n}^*)(1 + o_p(1))) \} \xrightarrow{d} N(0, 1).$$

(5) If we further have $A_n^* - \hat{A}_n^* = o_p((nh_n^{\frac{1}{2}})^{-1})$, $A_{1n}^* - \hat{A}_{1n}^* = o_p((nh_n^{\frac{1}{2}})^{-1})$, and $\hat{V}_T^* - [nh_n^{\frac{1}{2}} \frac{S_n^*}{E(\sigma^2(x_t))}]^2 = o_p(1)$, the claim in Theorem 5 follows.

$$(1) \epsilon_{t,0}^* = y_t^* - \hat{\mu}^* = y_t^* - \frac{1}{n} \sum_{t=1}^n y_t^* = \bar{y} + \epsilon_t^* - \bar{y} - \frac{1}{n} \sum_{t=1}^n \epsilon_t^* = \epsilon_t^* - \frac{1}{n} \sum_{t=1}^n \epsilon_t^*,$$

So $\frac{1}{n} \sum_{t=1}^n (\epsilon_{t,0}^*)^2 = \frac{1}{n} \sum_{t=1}^n (\epsilon_t^* - \frac{1}{n} \sum_{t=1}^n \epsilon_t^*)^2 = \frac{1}{n} \sum_{t=1}^n (\epsilon_t^*)^2 - (\frac{1}{n} \sum_{t=1}^n \epsilon_t^*)^2$. Since $E_{\hat{F}_t}(\epsilon_t^*) = 0$, $E_{\hat{F}_t}(\epsilon_t^*)^2 = \hat{\epsilon}_t^2$, so for given W , $\sqrt{n} \frac{\frac{1}{n} \sum_{t=1}^n \epsilon_t^*}{(\frac{1}{n} \sum_{t=1}^n \epsilon_t^2)^{\frac{1}{2}}} \xrightarrow{d} N(0, 1)$. In Theorem 3, we show that $\frac{1}{n} \sum_{t=1}^n \hat{\epsilon}_t^2 = \frac{1}{n} \sum_{t=1}^n \epsilon_t^2 + o_p(n^{-\frac{1}{2}}) = O_p(1)$, so

$\frac{1}{n} \sum_{t=1}^n \epsilon_t^* = O_p(n^{-\frac{1}{2}})$ and $(\frac{1}{n} \sum_{t=1}^n \epsilon_t^*)^2 = O_p(n^{-1})$. Thus, we conclude

$$\begin{aligned} \frac{1}{n} \sum_{t=1}^n (\epsilon_{t,0}^*)^2 &= \frac{1}{n} \sum_{t=1}^n (\epsilon_t^*)^2 + O_p(n^{-1}). \\ (2) \hat{m}^*(x) - \bar{y} &= \frac{1}{nh_n} \sum_{i=1}^n W_n(\frac{x_i - x}{h_n}, x) (y_i^* - \bar{y}) = \frac{1}{nh_n} \sum_{i=1}^n W_n(\frac{x_i - x}{h_n}, x) \epsilon_i^* \\ &= e' S_n^{-1}(x) \begin{bmatrix} \frac{1}{nh_n} \sum_{i=1}^n K(\frac{x_i - x}{h_n}) \epsilon_i^* \\ \frac{1}{nh_n} \sum_{i=1}^n K(\frac{x_i - x}{h_n}) \frac{x_i - x}{h_n} \epsilon_i^* \end{bmatrix} \end{aligned}$$

We follow the proof in Theorem 1(a)(2)(i) to have $\frac{2}{n} \sum_{t=1}^n (\bar{y} - \hat{m}^*(x_t)) \epsilon_t^* = (A_{1n}^* + A_{2n}^*)(1 + o_p(1))$.

(3) With similar arguments as in Theorem 1(a)(2)(ii), we obtain

$$\frac{1}{n} \sum_{t=1}^n (\bar{y} - \hat{m}^*(x_t))^2 = \underbrace{\frac{1}{n^3 h_n^2} \sum_{t=1}^n \sum_{i=1}^n \sum_{j=1}^n \frac{1}{f^2(x_t)} K(\frac{x_i - x_t}{h_n}) K(\frac{x_j - x_t}{h_n}) \epsilon_i^* \epsilon_j^*}_{I} (1 + o_p(1)).$$

(a) When $t = i = j$, $I = \frac{1}{n^3 h_n^2} \sum_{t=1}^n \frac{K^2(0)}{f^2(x_t)} (\epsilon_t^*)^2$. Since $E_{\hat{F}_t}(\epsilon_t^*)^2 = \hat{\epsilon}_t^2$,

$$\begin{aligned} E(|I||W) &= \frac{1}{n^3 h_n^2} \sum_{t=1}^n \frac{K^2(0)}{f^2(x_t)} E_{\hat{F}_t}(\epsilon_t^*)^2 = \frac{1}{n^3 h_n^2} \sum_{t=1}^n \frac{K^2(0)}{f^2(x_t)} \hat{\epsilon}_t^2 \\ &= \frac{1}{n^3 h_n^2} \sum_{t=1}^n \frac{K^2(0)}{f^2(x_t)} [\epsilon_t^2 + 2(m(x_t) - \hat{m}(x_t)) \epsilon_t + (m(x_t) - \hat{m}(x_t))^2] \\ &= O_p((n^2 h_n^2)^{-1}) = o_p(n^{-1}), \text{ since } \sup_{x_t \in G} |m(x_t) - \hat{m}(x_t)| = o_p(1). \end{aligned}$$

(b) When $t = i$ (or $t = j$),

$$\begin{aligned} nh_n^{\frac{3}{2}} I &= \frac{2}{n^2} \sum_{t=1}^n \sum_{j=1}^n \frac{K(0)}{2\sqrt{h_n} f^2(x_t)} K(\frac{x_j - x_t}{h_n}) \epsilon_t^* \epsilon_j^* = \frac{2}{n^2} \sum_{t=1}^n \sum_{j=1}^n \psi_n(Z_t, Z_j) \\ &= \frac{1}{n^2} \sum_{t=1}^n \sum_{j=1}^n (\psi_n(Z_t, Z_j) + \psi_n(Z_j, Z_t)) = \frac{1}{n^2} \sum_{t=1}^n \sum_{j=1}^n \phi_n(Z_t, Z_j) \end{aligned}$$

where $Z_t \equiv Z_t^* = (x_t, \epsilon_t^*)$. Since $E(\phi_n(Z_t, Z_j)|W, Z_t) = 0$, we apply Lemma 1 to obtain

$$nh_n^{\frac{3}{2}} I = O_p((\frac{1}{n^4} \sum_{t=1}^n \sum_{j=1}^n E(\phi_n^2(Z_t, Z_j)|W))^{\frac{1}{2}}).$$

$$\begin{aligned} &\frac{1}{n^4} \sum_{t=1}^n \sum_{j=1}^n E(\phi_n^2(Z_t, Z_j)|W) \\ &= \frac{1}{4n^4} \sum_{t=1}^n \sum_{j=1}^n [\frac{K^2(0)}{h_n f^4(x_t)} K^2(\frac{x_j - x_t}{h_n}) \hat{\epsilon}_t^2 \hat{\epsilon}_j^2 + \frac{K^2(0)}{h_n f^4(x_j)} K^2(\frac{x_t - x_j}{h_n}) \hat{\epsilon}_j^2 \hat{\epsilon}_t^2 + \frac{2K^2(0)}{h_n f^2(x_t) f^2(x_j)} K(\frac{x_j - x_t}{h_n}) K(\frac{x_t - x_j}{h_n}) \hat{\epsilon}_t^2 \hat{\epsilon}_j^2] \\ &= I_1 + I_2 + I_3. \end{aligned}$$

Let's take a look at I_1 .

$$\begin{aligned}
I_1 &= \frac{1}{4n^4} \sum_{t=1}^n \sum_{j=1}^n \frac{K^2(0)}{h_n f^4(x_t)} K^2\left(\frac{x_j - x_t}{h_n}\right) (m(x_t) - \hat{m}(x_t) + \epsilon_t)^2 (m(x_j) - \hat{m}(x_j) + \epsilon_j)^2 \\
&= \frac{1}{4n^4} \sum_{t=1}^n \sum_{\substack{j=1 \\ t \neq j}}^n \frac{K^2(0)}{h_n f^4(x_t)} K^2\left(\frac{x_j - x_t}{h_n}\right) [(m(x_t) - \hat{m}(x_t))^2 (m(x_j) - \hat{m}(x_j))^2 + (m(x_t) - \hat{m}(x_t))^2 \epsilon_j^2 \\
&\quad + 2(m(x_t) - \hat{m}(x_t))^2 (m(x_j) - \hat{m}(x_j)) \epsilon_j + \epsilon_t^2 (m(x_j) - \hat{m}(x_j))^2 + \epsilon_t^2 \epsilon_j^2 + 2\epsilon_t^2 (m(x_j) - \hat{m}(x_j)) \epsilon_j \\
&\quad + 2(m(x_t) - \hat{m}(x_t)) \epsilon_t (m(x_j) - \hat{m}(x_j))^2 + 2(m(x_t) - \hat{m}(x_t)) \epsilon_t \epsilon_j^2 \\
&\quad + 4(m(x_t) - \hat{m}(x_t)) (m(x_j) - \hat{m}(x_j)) \epsilon_t \epsilon_j].
\end{aligned}$$

Let's consider a generic term showing up repeatedly as

$$\begin{aligned}
I_{110} &= \frac{1}{n^2} \sum_{t=1}^n \sum_{\substack{j=1 \\ t \neq j}}^n \frac{K^2(0)}{h_n f^4(x_t)} K^2\left(\frac{x_j - x_t}{h_n}\right) |\epsilon_t|^k |\epsilon_j|^\tau = \frac{1}{n^2} \sum_{t=1}^n \sum_{\substack{j=1 \\ t \neq j}}^n \psi_n(Z_t, Z_j) \\
&= \frac{1}{2n^2} \sum_{t=1}^n \sum_{\substack{j=1 \\ t \neq j}}^n (\psi_n(Z_t, Z_j) + \psi_n(Z_j, Z_t)) = \frac{1}{2n^2} \sum_{t=1}^n \sum_{\substack{j=1 \\ t \neq j}}^n \phi_n(Z_t, Z_j) \\
&= \frac{1}{n} \sum_{t=1}^n E(\phi_n(Z_t, Z_j) | Z_t) - \frac{1}{2} E\phi_n(Z_t, Z_j) + O_p\left(\frac{1}{n} (E\phi_n^2(Z_t, Z_j))^{1/2}\right)
\end{aligned}$$

where $Z_t = (x_t, \epsilon_t)$, $k, \tau = \{0, 1, 2\}$, and we apply Lemma 1 in the last line.

$$E(\phi_n(Z_t, Z_j) | Z_t) = E(\psi_n(Z_t, Z_j) | Z_t) + E(\psi_n(Z_j, Z_t) | Z_t),$$

$$\begin{aligned}
E(\psi_n(Z_t, Z_j) | Z_t) &= \int \frac{K^2(0)}{h_n f^4(x_t)} K^2\left(\frac{x_j - x_t}{h_n}\right) |\epsilon_t|^k E(|\epsilon_j|^\tau | x_j) f(x_j) dx_j \rightarrow \frac{K^2(0)}{f^3(x_t)} |\epsilon_t|^k E(|\epsilon_t|^\tau | x_t) \int K^2(\psi) d\psi, \\
E(\psi_n(Z_j, Z_t) | Z_t) &= \int \frac{K^2(0)}{h_n f^4(x_j)} K^2\left(\frac{x_t - x_j}{h_n}\right) |\epsilon_t|^\tau E(|\epsilon_j|^k | x_j) f(x_j) dx_j \rightarrow \frac{K^2(0)}{f^3(x_t)} |\epsilon_t|^\tau E(|\epsilon_t|^k | x_t) \int K^2(\psi) d\psi, \\
E(\phi_n(Z_t, Z_j)) &\rightarrow 2K^2(0) \int K^2(\psi) d\psi E[E(|\epsilon_t|^k | x_t) E(|\epsilon_t|^\tau | x_t) \frac{1}{f^3(x_t)}], \text{ and} \\
h_n E\phi_n^2(Z_t, Z_j) &\leq C h_n^2 E\psi_n^2(Z_t, Z_j) \rightarrow K^2(0) \int K^4(\psi) d\psi E[E(|\epsilon_t|^{2k} | x_t) E(|\epsilon_t|^{2\tau} | x_t) \frac{1}{f^{\tau}(x_t)}] < \infty
\end{aligned}$$

with assumption A9 and A4(1).

So we conclude $E\phi_n^2(Z_t, Z_j) = O(\frac{1}{h_n})$, $\frac{1}{n} \sum_{t=1}^n E(\phi_n(Z_t, Z_j) | Z_t) = O_p(1)$ since $nh_n \rightarrow \infty$, $E(\phi_n(Z_t, Z_j)) = O(1)$, $O_p(\frac{1}{n} (E\phi_n^2(Z_t, Z_j))^{1/2}) = O_p((nh_n^{-1/2})^{-1})$, and thus $I_{110} = O_p(1)$. Furthermore, since $\sup_{x_t \in G} |\hat{m}(x_t) - m(x_t)| = o_p(1)$, we have $I_1 = O_p(n^{-2})$.

With similar arguments, we obtain $I_2 = O_p(n^{-2})$ and $I_3 = O_p(n^{-2})$. So in all, we conclude $nh_n^{\frac{3}{2}} I = O_p(n^{-1})$, and thus $I = O_p((n^2 h_n^{\frac{3}{2}})^{-1}) = o_p(n^{-1})$.

$$(c) \text{ When } i = j \text{ and } t \neq i, I = \frac{1}{n^3 h_n^2} \sum_{t=1}^n \sum_{\substack{i=1 \\ t \neq i}}^n \frac{1}{f^2(x_t)} K^2\left(\frac{x_i - x_t}{h_n}\right) (\epsilon_i^*)^2 = A_n^*.$$

(d) When indices t, i , and j are distinct, we let $\psi_n(Z_t, Z_i, Z_j) = \frac{1}{h_n^2 f^2(x_t)} K\left(\frac{x_i - x_t}{h_n}\right) K\left(\frac{x_j - x_t}{h_n}\right) \epsilon_i^* \epsilon_j^*$, and we follow Theorem 1(a)(2)(ii)(d) to have

$$I = \frac{1}{3} \left[O(n^{-4}) + \binom{n}{3}^{-1} \right] \sum_{1=t < i < j=n} \phi_n(Z_t, Z_i, Z_j),$$

where $\phi_n(Z_t, Z_i, Z_j) = \psi_n(Z_t, Z_i, Z_j) + \psi_n(Z_i, Z_t, Z_j) + \psi_n(Z_j, Z_i, Z_t)$ and $Z_t = (x_t, \epsilon_t^*)$. We focus on the third order U-statistic $U_n = \binom{n}{3}^{-1} \sum_{1=t < i < j=n} \phi_n(Z_t, Z_i, Z_j)$.

We note $\theta_n = E(\phi_n(Z_t, Z_i, Z_j) | W) = 0$ and $\phi_{1n}(Z_t) = E(\phi_n(Z_t, Z_i, Z_j) | W, Z_t) = 0$ because $E_{\hat{F}_t} \epsilon_t^* = 0$. So we apply the H-decomposition in Lemma 3 to have

$$\begin{aligned}
U_n &= \binom{n}{3}^{-1} \binom{n-2}{1} \sum_{t=1}^n \sum_{\substack{i=1 \\ t < i}}^n \phi_{2n}(Z_t, Z_i) \\
&\quad + \binom{n}{3}^{-1} \underbrace{\left\{ \sum_{1=t < i < j=n} \phi_n(Z_t, Z_i, Z_j) - (n-2) \sum_{t=1}^n \sum_{\substack{i=1 \\ t < i}}^n \phi_{2n}(Z_t, Z_i) \right\}}_{U_{1n}},
\end{aligned}$$

where $\phi_{2n}(Z_t, Z_i) = \frac{\epsilon_t^* \epsilon_i^*}{h_n^2} E(K(\frac{x_i - x_j}{h_n}) K(\frac{x_t - x_j}{h_n}) \frac{1}{f^2(x_j)} | x_t, x_i)$.

Since $(n-2) \sum_{t=1}^n \sum_{i=1}^n \phi_{2n}(Z_t, Z_i) = \sum_{t=1}^n \sum_{i=1}^n \sum_{j=n}^n [\phi_{2n}(Z_t, Z_i) + \phi_{2n}(Z_t, Z_j) + \phi_{2n}(Z_i, Z_j)]$, we have

$$\begin{aligned} U_{1n} &= \left(\begin{matrix} n \\ 3 \end{matrix} \right)^{-1} \sum_{t=1}^n \sum_{i=1}^n \sum_{j=n}^n \{\phi_n(Z_t, Z_i, Z_j) - [\phi_{2n}(Z_t, Z_i) + \phi_{2n}(Z_t, Z_j) + \phi_{2n}(Z_i, Z_j)]\} \\ &= \left(\begin{matrix} n \\ 3 \end{matrix} \right)^{-1} \sum_{t=1}^n \sum_{i=1}^n \sum_{j=n}^n \Phi_n(Z_t, Z_i, Z_j). \end{aligned}$$

By construction for given W , conditioning on Z_t, Z_i, Z_j , or $\{Z_t, Z_i\}, \{Z_t, Z_j\}, \{Z_i, Z_j\}$, the expectation of $\Phi_n(Z_t, Z_i, Z_j)$ is zero. So $E(U_{1n}|W) = 0$. Furthermore, $E(\Phi_n(Z_t, Z_i, Z_j)\Phi_n(Z_{t'}, Z_{i'}, Z_{j'})|W) = 0$ if $\{t, i, j\} \neq \{t', i', j'\}$. So

$$\begin{aligned} V(U_{1n}|W) &= \left(\begin{matrix} n \\ 3 \end{matrix} \right)^{-2} \sum_{t=1}^n \sum_{i=1}^n \sum_{j=n}^n E(\Phi_n^2(Z_t, Z_i, Z_j)|W) \\ &\leq c \left(\begin{matrix} n \\ 3 \end{matrix} \right)^{-2} \sum_{t=1}^n \sum_{i=1}^n \sum_{j=n}^n E(\phi_n^2(Z_t, Z_i, Z_j)|W) \leq c \left(\begin{matrix} n \\ 3 \end{matrix} \right)^{-2} \sum_{t=1}^n \sum_{i=1}^n \sum_{j=n}^n E(\psi_n^2(Z_t, Z_i, Z_j)|W) \\ &= c \left(\begin{matrix} n \\ 3 \end{matrix} \right)^{-2} \sum_{t=1}^n \sum_{i=1}^n \sum_{j=n}^n \frac{1}{h_n^4 f^4(x_t)} K^2\left(\frac{x_i - x_t}{h_n}\right) K^2\left(\frac{x_j - x_t}{h_n}\right) \hat{\epsilon}_i^2 \hat{\epsilon}_j^2 \\ &= c \left(\begin{matrix} n \\ 3 \end{matrix} \right)^{-2} \sum_{t=1}^n \sum_{i=1}^n \sum_{j=n}^n \frac{1}{h_n^4 f^4(x_t)} K^2\left(\frac{x_i - x_t}{h_n}\right) K^2\left(\frac{x_j - x_t}{h_n}\right) [\epsilon_i^2 \epsilon_j^2 + \epsilon_i^2 (m(x_j) - \hat{m}(x_j))^2 \\ &\quad + 2\epsilon_i^2 \epsilon_j (m(x_j) - \hat{m}(x_j)) + (m(x_i) - \hat{m}(x_i))^2 \epsilon_j^2 + (m(x_i) - \hat{m}(x_i))^2 (m(x_j) - \hat{m}(x_j))^2 \\ &\quad + 2(m(x_i) - \hat{m}(x_i))^2 \epsilon_j (m(x_j) - \hat{m}(x_j)) + 2(m(x_i) - \hat{m}(x_i)) \epsilon_i \epsilon_j^2 \\ &\quad + 2(m(x_i) - \hat{m}(x_i)) \epsilon_i (m(x_j) - \hat{m}(x_j))^2 + 4(m(x_i) - \hat{m}(x_i)) (m(x_j) - \hat{m}(x_j)) \epsilon_i \epsilon_j] \\ &= I_1 + I_2 + \dots + I_9 \end{aligned}$$

$I_1 = O_p(n^{-3}h_n^{-2}) = o_p(n^{-2})$ as we have show in Theorem 1(a)(2)(ii)(d) for term $H_n^{(3)}$. With assumption A9, and $\sup_{x_t \in G} |\hat{m}(x_t) - m(x_t)| = o_p(1)$, we have $I_i = o_p(n^{-3}h_n^{-2}) = o_p(n^{-2})$ for $i = 2, \dots, 9$ with similar arguments as in (b) above. So we conclude $U_{1n} = o_p(n^{-1})$. So

$$\begin{aligned} U_n &= \left(\begin{matrix} n \\ 3 \end{matrix} \right)^{-1} \left(\begin{matrix} n-2 \\ 1 \end{matrix} \right) \sum_{t=1}^n \sum_{i=1}^n \phi_{2n}(Z_t, Z_i) + o_p(n^{-1}) = \frac{6}{n(n-1)} \sum_{t=1}^n \sum_{i=1}^n \phi_{2n}(Z_t, Z_i) + o_p(n^{-1}). \\ I &= \frac{2}{n(n-1)} \sum_{t=1}^n \sum_{i=1}^n \phi_{2n}(Z_t, Z_i) + o_p(n^{-1}) \\ &= \frac{1}{n(n-1)} \sum_{t=1}^n \sum_{i=1}^n [\frac{\epsilon_i^* \epsilon_t^*}{h_n^2} E(K(\frac{x_i - x_j}{h_n}) K(\frac{x_t - x_j}{h_n}) \frac{1}{f^2(x_j)} | x_t, x_i) + \frac{\epsilon_t^* \epsilon_i^*}{h_n^2} E(K(\frac{x_t - x_j}{h_n}) K(\frac{x_i - x_j}{h_n}) \frac{1}{f^2(x_j)} | x_i, x_t)] + o_p(n^{-1}). \end{aligned}$$

Combining results in (a)-(d) above, we obtain the claim in (3).

$$(4) A_{3n}^* = [O(n^{-3}) + \frac{1}{n^2}] \underbrace{\sum_{t=1}^n \sum_{i=1}^n \sum_{t < i} [\phi_{2n}(Z_t, Z_i) + \phi_{2n}(Z_i, Z_t)]}_{A_{31n}^*} + o_p(n^{-1}),$$

$$\begin{aligned} \text{where } A_{31n}^* &= \frac{1}{n^2} \sum_{t=1}^n \sum_{i=1}^n [\frac{\epsilon_i^* \epsilon_t^*}{h_n^2} E(K(\frac{x_i - x_j}{h_n}) K(\frac{x_t - x_j}{h_n}) \frac{1}{f^2(x_j)} | x_t, x_i) \\ &\quad + \frac{\epsilon_t^* \epsilon_i^*}{h_n^2} E(K(\frac{x_t - x_j}{h_n}) K(\frac{x_i - x_j}{h_n}) \frac{1}{f^2(x_j)} | x_i, x_t)] + o_p(n^{-1}). \end{aligned}$$

Since $E(A_{31n}^*|W) = 0$,

$$\begin{aligned} V(A_{31n}^*|W) &= \frac{1}{n^4} \sum_{t=1}^n \sum_{i=1}^n \{ \frac{1}{h_n^4} [E(K(\frac{x_i - x_j}{h_n}) K(\frac{x_t - x_j}{h_n}) \frac{1}{f^2(x_j)} | x_t, x_i)]^2 \hat{\epsilon}_i^2 \hat{\epsilon}_t^2 \\ &\quad + \frac{1}{h_n^4} [E(K(\frac{x_t - x_j}{h_n}) K(\frac{x_i - x_j}{h_n}) \frac{1}{f^2(x_j)} | x_i, x_t)]^2 \hat{\epsilon}_t^2 \hat{\epsilon}_i^2 \\ &\quad + \frac{2}{h_n^4} [E(K(\frac{x_i - x_j}{h_n}) K(\frac{x_t - x_j}{h_n}) \frac{1}{f^2(x_j)} | x_t, x_i)]^2 \hat{\epsilon}_i^2 \hat{\epsilon}_t^2 \} \\ &= O_p(n^{-2}h_n^{-1}), \end{aligned}$$

we conclude $A_{31n}^* = O_p((nh_n^{\frac{1}{2}})^{-1})$. It is sufficient to show $(A_{2n}^* + A_{31n}^*)(S_n^*)^{-1} \xrightarrow{d} N(0, 1)$.

$$\begin{aligned}
& A_{2n}^* + A_{31n}^* \\
&= \sum_{t=1}^n \sum_{i=1}^n \frac{1}{n^2 h_n} \left[-\frac{2}{f(x_t)} K\left(\frac{x_i - x_t}{h_n}\right) \epsilon_i^* \epsilon_t^* - \frac{2}{f(x_i)} K\left(\frac{x_t - x_i}{h_n}\right) \epsilon_t^* \epsilon_i^* \right. \\
&\quad \left. + \epsilon_i^* \epsilon_t^* E\left(\frac{1}{h_n} K\left(\frac{x_i - x_j}{h_n}\right) K\left(\frac{x_t - x_j}{h_n}\right) \frac{1}{f^2(x_j)} |x_t, x_i\right) + \epsilon_t^* \epsilon_i^* E\left(\frac{1}{h_n} K\left(\frac{x_t - x_j}{h_n}\right) K\left(\frac{x_i - x_j}{h_n}\right) \frac{1}{f^2(x_j)} |x_i, x_t\right) \right] \\
&= \sum_{t=1}^n \sum_{i=1}^n [\psi_n(Z_t^*, Z_i^*) + \psi_n(Z_i^*, Z_t^*) + \psi'_n(Z_t^*, Z_i^*) + \psi'_n(Z_i^*, Z_t^*)] \\
&= \sum_{t=1}^n \sum_{i=1}^n \phi_n(Z_t^*, Z_i^*), \text{ where } Z_t^* = (x_t, \epsilon_t^*).
\end{aligned}$$

Since $\phi_n(Z_t^*, Z_i^*)$ is symmetric and $E(\phi_n(Z_t^*, Z_i^*) | W, Z_t^*) = 0$, $A_{2n}^* + A_{31n}^*$ is a degenerate second order U-statistic. We apply Proposition 3.2 of de Jong (1987) to have $(A_{2n}^* + A_{31n}^*)(S_n^*)^{-1} \xrightarrow{d} N(0, 1)$ if we have G_I, G_{II} and G_{IV} are of order $o_p((S_n^*)^4)$, where

$$\begin{aligned}
(S_n^*)^2 &= E\left(\left(\sum_{t=1}^n \sum_{i=1}^n \phi_n(Z_t^*, Z_i^*)\right)^2 | W\right), G_I = \sum_{t=1}^n \sum_{i=1}^n E(\phi_n^4(Z_t^*, Z_i^*) | W) \\
G_{II} &= \sum_{1=t < i < j=n} \sum_{t < i} \sum_{t < j} \{E(\phi_n^2(Z_t^*, Z_i^*) \phi_n^2(Z_t^*, Z_j^*) | W) + E(\phi_n^2(Z_i^*, Z_t^*) \phi_n^2(Z_i^*, Z_j^*) | W) \\
&\quad + E(\phi_n^2(Z_j^*, Z_t^*) \phi_n^2(Z_j^*, Z_i^*) | W)\} = G_{II1} + G_{II2} + G_{II3}, \\
G_{IV} &= \sum_{1=t < i < j < l=n} \sum_{t < i} \sum_{t < j} \sum_{t < l} \{E(\phi_n(Z_t^*, Z_i^*) \phi_n(Z_t^*, Z_j^*) \phi_n(Z_i^*, Z_j^*) \phi_n(Z_l^*, Z_j^*) | W) \\
&\quad + E(\phi_n(Z_t^*, Z_i^*) \phi_n(Z_t^*, Z_l^*) \phi_n(Z_j^*, Z_i^*) \phi_n(Z_j^*, Z_l^*) | W) \\
&\quad + E(\phi_n(Z_i^*, Z_j^*) \phi_n(Z_i^*, Z_l^*) \phi_n(Z_j^*, Z_i^*) \phi_n(Z_j^*, Z_l^*) | W)\} = G_{IV1} + G_{IV2} + G_{IV3}.
\end{aligned}$$

We can show that

$$\begin{aligned}
(S_n^*)^2 &= n^{-2} h_n^{-1} 2 \int (\sigma^2(x_t))^2 dx_t \int (2K(\psi) - \kappa(\psi))^2 d\psi + o_p(n^{-2} h_n^{-1}) \\
&= n^{-2} h_n^{-1} \sigma_\phi^2 + o_p(n^{-2} h_n^{-1}) = O_p(n^{-2} h_n^{-1}).
\end{aligned}$$

and that G_I, G_{II} and G_{IV} are of order $o_p((S_n^*)^4)$, and details are omitted for brevity.

(5) (i) $A_n^* - \hat{A}_n^* = o_p((nh_n^{\frac{1}{2}})^{-1})$ conditioning on W .

As shown in (1), $\epsilon_{t,0}^* = \epsilon_t^* - \frac{1}{n} \sum_{t=1}^n \epsilon_t^* = \epsilon_t^* + O_p(n^{-\frac{1}{2}})$, and $E(|\epsilon_t^*|^0 | W) = 1$, $E(|\epsilon_t^*|^1 | W) = \frac{2}{\sqrt{5}} |\hat{\epsilon}_t|$, $E(|\epsilon_t^*|^2 | W) = \hat{\epsilon}_t^2$, the claimed result follows with similar arguments as in Theorem 1(b).

(ii) $A_{1n}^* - \hat{A}_{1n}^* = o_p((nh_n^{\frac{1}{2}})^{-1})$ conditioning on W can be shown similarly.

(iii) We claim: $\hat{V}_T^* - \frac{(nh_n^{\frac{1}{2}})^2 (S_n^*)^2}{(E(\sigma^2(x_t)))^2} = o_p(1)$ conditioning on W , where in (4) above, we have shown $(nh_n^{\frac{1}{2}})^2 (S_n^*)^2 = \sigma_\phi^2 + o_p(1)$ with $\sigma_\phi^2 = 2 \int \sigma^4(x_t) dx_t \int (2K(\psi) - \kappa(\psi))^2 d\psi$, and by definition,

$$\hat{V}_T^* = \frac{\sigma_\phi^{2*}}{(\frac{1}{n} \sum_{t=1}^n (\epsilon_{t,0}^*)^2)^2}, \text{ with } \sigma_\phi^{2*} = [\frac{1}{n^2} \sum_{t=1}^n \sum_{i=1}^n K\left(\frac{x_i - x_t}{h_n}\right) \frac{(\epsilon_{i,0}^*)^2 (\epsilon_{t,0}^*)^2}{h_n f^2(x_i)}] 2 \int (2K(\psi) - \kappa(\psi))^2 d\psi.$$

Since $\frac{\hat{a}}{\hat{b}} - \frac{a}{b} = \frac{1}{b} [\hat{a} - a - (\hat{b} - b) \frac{a}{b}]$, it is sufficient to prove the claim by showing

(a) $(\frac{1}{n} \sum_{t=1}^n (\epsilon_{t,0}^*)^2)^2 \xrightarrow{P} (E(\sigma^2(x_t)))^2$ conditioning on W , where $(E(\sigma^2(x_t)))^2 > 0$, and

(b) $\frac{1}{n^2} \sum_{t=1}^n \sum_{i=1}^n K\left(\frac{x_i - x_t}{h_n}\right) \frac{(\epsilon_{i,0}^*)^2 (\epsilon_{t,0}^*)^2}{h_n f^2(x_i)} \xrightarrow{P} \int \sigma^4(x_t) dx_t$.

(a) It is sufficient to show $\frac{1}{n} \sum_{t=1}^n (\epsilon_{t,0}^*)^2 \xrightarrow{P} E(\sigma^2(x_t))$ conditioning on W .

Note $\frac{1}{n} \sum_{t=1}^n (\epsilon_{t,0}^*)^2 = \frac{1}{n} \sum_{t=1}^n (\epsilon_t^*)^2 + O_p(n^{-1})$. Furthermore,

$E(\frac{1}{n} \sum_{t=1}^n (\epsilon_t^*)^2 | W) = E(|\frac{1}{n} \sum_{t=1}^n (\epsilon_t^*)^2| | W) = \frac{1}{n} \sum_{t=1}^n \hat{\epsilon}_t^2 = \frac{1}{n} \sum_{t=1}^n \epsilon_t^2 + o_p(1)$. Since $\frac{1}{n} \sum_{t=1}^n \epsilon_t^2 \xrightarrow{P} E(\sigma^2(x_t))$, we have the claimed result in (a).

$$\begin{aligned}
(b) \frac{1}{n^2} \sum_{t=1}^n \sum_{i=1}^n K\left(\frac{x_i - x_t}{h_n}\right) \frac{(\epsilon_{i,0}^*)^2 (\epsilon_{t,0}^*)^2}{h_n f^2(x_t)} &= \frac{1}{n^2} \sum_{t=1}^n \sum_{i=1}^n K\left(\frac{x_i - x_t}{h_n}\right) \frac{(\epsilon_i^* + O_p(n^{-\frac{1}{2}}))^2 (\epsilon_t^* + O_p(n^{-\frac{1}{2}}))^2}{h_n} \left[\frac{1}{f^2(x_t)} - \frac{1}{f^2(x_t)} \right] \\
&\quad + \frac{1}{n^2} \sum_{t=1}^n \sum_{i=1}^n K\left(\frac{x_i - x_t}{h_n}\right) \frac{(\epsilon_i^* + O_p(n^{-\frac{1}{2}}))^2 (\epsilon_t^* + O_p(n^{-\frac{1}{2}}))^2}{h_n} \frac{1}{f^2(x_t)} \\
&= V_{11} + V_{12}.
\end{aligned}$$

$$\begin{aligned}
V_{11} &\leq \sup_{x_t \in G} \left| \frac{1}{f^2(x_t)} - \frac{1}{f^2(x_t)} \right| \frac{1}{n^2} \sum_{t=1}^n \sum_{i=1}^n \frac{1}{h_n} K\left(\frac{x_i - x_t}{h_n}\right) [(\epsilon_i^*)^2 + 2\epsilon_i^* O_p(n^{-\frac{1}{2}}) + O_p(n^{-1})] \\
&\quad \times [(\epsilon_t^*)^2 + 2\epsilon_t^* O_p(n^{-\frac{1}{2}}) + O_p(n^{-1})] \\
&= O_p(L_n) \frac{1}{n^2} \sum_{t=1}^n \sum_{i=1}^n \frac{1}{h_n} K\left(\frac{x_i - x_t}{h_n}\right) [(\epsilon_i^*)^2 (\epsilon_t^*)^2 + 2(\epsilon_i^*)^2 \epsilon_t^* O_p(n^{-\frac{1}{2}}) + (\epsilon_i^*)^2 O_p(n^{-1}) \\
&\quad + 2\epsilon_i^* (\epsilon_t^*)^2 O_p(n^{-\frac{1}{2}}) + 4\epsilon_i^* \epsilon_t^* O_p(n^{-1}) + 2\epsilon_i^* O_p(n^{-\frac{3}{2}}) + (\epsilon_t^*)^2 O_p(n^{-1}) + 2\epsilon_t^* O_p(n^{-\frac{3}{2}}) + O_p(n^{-2})].
\end{aligned}$$

$$\text{Let's consider } V_{110} = \frac{1}{n^2} \sum_{t=1}^n \sum_{i=1}^n \frac{1}{h_n} K\left(\frac{x_i - x_t}{h_n}\right) (\epsilon_i^*)^2 (\epsilon_t^*)^2.$$

$$E(|V_{110}|W) = \frac{1}{n^2} \sum_{t=1}^n \sum_{i=1}^n \frac{1}{h_n} K\left(\frac{x_i - x_t}{h_n}\right) (\hat{\epsilon}_i)^2 (\hat{\epsilon}_t)^2 = O_p(1), \text{ following similar arguments as in term } I_{110} \text{ in}$$

(3)(b) above. We can show similarly for $(k, \tau) = (2, 1), (2, 0), (1, 2), (1, 1), (1, 0), (0, 2), (0, 1), (0, 0)$ that $\frac{1}{n^2} \sum_{t=1}^n \sum_{i=1}^n \frac{1}{h_n} K\left(\frac{x_i - x_t}{h_n}\right) (\epsilon_i^*)^k (\epsilon_t^*)^\tau = O_p(1)$, and thus, $V_{11} = O_p(L_n) = o_p(1)$.

$$\begin{aligned}
V_{12} &\leq \frac{1}{n^2} \sum_{t=1}^n \sum_{i=1}^n \frac{1}{h_n f^2(x_t)} K\left(\frac{x_i - x_t}{h_n}\right) [(\epsilon_i^*)^2 (\epsilon_t^*)^2 + 2(\epsilon_i^*)^2 \epsilon_t^* O_p(n^{-\frac{1}{2}}) + (\epsilon_i^*)^2 O_p(n^{-1}) \\
&\quad + 2\epsilon_i^* (\epsilon_t^*)^2 O_p(n^{-\frac{1}{2}}) + 4\epsilon_i^* \epsilon_t^* O_p(n^{-1}) + 2\epsilon_i^* O_p(n^{-\frac{3}{2}}) + (\epsilon_t^*)^2 O_p(n^{-1}) + 2\epsilon_t^* O_p(n^{-\frac{3}{2}}) + O_p(n^{-2})].
\end{aligned}$$

$$\text{Let's consider } V_{120} = \frac{1}{n^2} \sum_{t=1}^n \sum_{i=1}^n \frac{1}{h_n f^2(x_t)} K\left(\frac{x_i - x_t}{h_n}\right) (\epsilon_i^*)^2 (\epsilon_t^*)^2.$$

$$E(V_{120}|W) = E(|V_{120}|W) = \frac{1}{n^2} \sum_{t=1}^n \sum_{i=1}^n \frac{1}{h_n f^2(x_t)} K\left(\frac{x_i - x_t}{h_n}\right) (\hat{\epsilon}_i)^2 (\hat{\epsilon}_t)^2 = \int \sigma^4(x_t) dx_t + o_p(1), \text{ as shown in}$$

Theorem 1(b)(iii). Following similar arguments, we can show similarly for (k, τ) considered above,

$$\frac{1}{n^2} \sum_{t=1}^n \sum_{i=1}^n \frac{1}{h_n f^2(x_t)} K\left(\frac{x_i - x_t}{h_n}\right) (\epsilon_i^*)^k (\epsilon_t^*)^\tau = O_p(1), \text{ and thus, } V_{12} = \int \sigma^4(x_t) dx_t + o_p(1). \text{ Results on } V_{11} \text{ and}$$

V_{12} give the claimed result in (b).

Sketch of proof of Remark 1 in section 3: Proof.

Based on our Theorems 1-5, we now establish the asymptotic properties of the alternative tests in section 2. We cite the conditions 1-7 for $d = 1$ in Doksum and Samarov (1995) which are used to establish the asymptotic distribution for the nonparametric R^2 estimators. To obtain the properties of the tests, we need the following additional assumptions.

R1. The weight function $w(x_i)$ in their condition 6 satisfies $\frac{1}{n} \sum_{i=1}^n w(x_i) = 1$.

R2. $f(x)$, the marginal density of x in their condition 3 is a bounded function.

R3. The kernel function $K(\cdot)$ in their condition 5 satisfies the Lipschitz condition.

R4. For $\epsilon = y - m(x)$, $\sigma^2(x) = E(\epsilon^2|x)$ is a continuous function in $x \in S_x$, where S_x is the support of $f(x)$. $E(\epsilon^4|x)$ is a continuous function in $x \in S_x$ and $E(\epsilon^4|x) < \infty$.

R5 The derivative of $m(x)$ up to order k is a bounded function, and uniformly continuous on S_x .

$$\text{For } \hat{\eta}_1^2 = \frac{S_y^2 - \frac{1}{n} \sum_{i=1}^n (y_i - \hat{m}(x_i))^2 w(x_i)}{S_y^2}:$$

(1) $S_y^2 = \frac{1}{n} \sum_{i=1}^n y_i^2 w(x_i) - \bar{y}_w^2 = \frac{1}{n} \sum_{i=1}^n y_i^2 w(x_i) - 2\bar{y}_w \mu_{y,w} + (\mu_{y,w})^2 + O_p(n^{-1})$ by Theorem 3's proof (1).

$$\frac{1}{n} \sum_{i=1}^n (y_i - \tilde{m}(x_i))^2 w(x_i) = \frac{1}{n} \sum_{i=1}^n \epsilon_i^2 w(x_i) - \frac{2}{n} \sum_{i=1}^n (\tilde{m}(x_i) - m(x_i)) \epsilon_i w(x_i) + \frac{1}{n} \sum_{i=1}^n (\tilde{m}(x_i) - m(x_i))^2 w(x_i).$$

We observe that with the conditions, $\sup_{\Phi} |\frac{1}{f(x_i)} - \frac{1}{\bar{f}(x_i)}| = o_p(1)$ where Φ is defined in condition 1 as an open convex set in $\mathfrak{R}$ such that $\inf_{\Phi} f(x) \geq \delta$ for some $\delta > 0$ and $\bar{\Phi}$ denotes its bounded closure. Define $g(x) = m(x)f(x)$, then $\tilde{m}(x_i) - m(x_i) = \frac{\tilde{g}(x_i)}{\tilde{f}(x_i)} - \frac{g(x_i)}{f(x_i)} = \frac{\tilde{g}(x_i) - \tilde{f}(x_i)m(x_i)}{\tilde{f}(x_i)} + (\frac{1}{\tilde{f}(x_i)} - \frac{1}{f(x_i)})(\tilde{g}(x_i) - \tilde{f}(x_i)m(x_i))$
 $= \frac{\tilde{g}(x_i) - \tilde{f}(x_i)m(x_i)}{\tilde{f}(x_i)}(1 + o_p(1))$ uniformly over $x_i \in \Phi$. Then

$$\begin{aligned} & -\frac{2}{n} \sum_{i=1}^n (\tilde{m}(x_i) - m(x_i)) \epsilon_i w(x_i) = -2[\frac{1}{n} \sum_{i=1}^n \frac{1}{\tilde{f}(x_i)} (\tilde{g}(x_i) - \tilde{f}(x_i)m(x_i)) \epsilon_i w(x_i)](1 + o_p(1)) \\ & = -2[\frac{1}{n(n-1)} \sum_{i=1}^n \sum_{j=1, j \neq i}^n K(\frac{x_j - x_i}{h_n})(m(x_j) - m(x_i)) \frac{\epsilon_i w(x_i)}{h_n f(x_i)} + \frac{1}{n(n-1)} \sum_{i=1}^n \sum_{j=1, j \neq i}^n K(\frac{x_j - x_i}{h_n}) \frac{\epsilon_j \epsilon_i w(x_i)}{h_n f(x_i)}](1 + o_p(1)) \\ & = -2[T_{na} + T_{nb}]. \\ & \frac{1}{n} \sum_{i=1}^n (\tilde{m}(x_i) - m(x_i))^2 w(x_i) = [\frac{1}{n} \sum_{i=1}^n \frac{1}{\tilde{f}^2(x_i)} (\tilde{g}(x_i) - \tilde{f}(x_i)m(x_i))^2 w(x_i)](1 + o_p(1)) \\ & = \{ \frac{1}{n(n-1)^2} \sum_{i=1}^n \sum_{j \neq i} \sum_{l \neq i} \frac{w(x_i)}{h_n^2 f^2(x_i)} K(\frac{x_j - x_i}{h_n}) K(\frac{x_l - x_i}{h_n}) [(m(x_j) - m(x_i))(m(x_l) - m(x_i)) \\ & \quad + (m(x_j) - m(x_i))\epsilon_l + (m(x_l) - m(x_i))\epsilon_j + \epsilon_l \epsilon_j] \}(1 + o_p(1)) \\ & = \{T_{nc} + T_{nd} + T_{ne} + T_{nf}\}(1 + o_p(1)). \\ T_{nf} & = [\frac{2}{n(n-1)} \sum_{i=1}^n \sum_{j=1}^n \epsilon_i \epsilon_j E(\frac{w(x_i)}{h_n^2 f^2(x_i)} K(\frac{x_i - x_l}{h_n}) K(\frac{x_j - x_l}{h_n}) | x_i, x_j) \\ & \quad + \frac{1}{n(n-1)^2} \sum_{i=1}^n \sum_{j=1, j \neq i}^n K^2(\frac{x_j - x_i}{h_n}) \epsilon_j^2 \frac{w(x_i)}{h_n^2 f^2(x_i)}](1 + o_p(1)) = [T_{nf1} + T_{n0}](1 + o_p(1)). \end{aligned}$$

Under H_0 , $T_{ni} = 0$ for $i = a, c, d, e$. So $\frac{1}{n} \sum_{i=1}^n (y_i - \tilde{m}(x_i))^2 w(x_i) = \frac{1}{n} \sum_{i=1}^n \epsilon_i^2 w(x_i) + [-2T_{nb} + T_{nf1} + T_{n0}](1 + o_p(1))$. H_0 implies $y_i = \mu + \epsilon_i$, so $S_y^2 = \frac{1}{n} \sum_{i=1}^n \epsilon_i^2 w(x_i) + O_p(n^{-1}) = \sigma_{y,w}^2 + o_p(1)$. The structure of $\hat{T}_{1n}$ is similar to $\hat{T}_n$ and we follow Theorem 1(a) to show $nh_n^{1/2}[2T_{nb} - T_{nf1}] \xrightarrow{d} N(0, \sigma_{\phi 1}^2)$ for $\sigma_{\phi 1}^2 = E[\frac{w^2(x)}{f(x)} \sigma^4(x)] 2 \int (2K(\psi) - \kappa(\psi))^2 d\psi$. It implies $nh_n^{1/2}(\hat{\eta}_1^2 + (S_y^2)^{-1} T_{n0}) \xrightarrow{d} N(0, (\sigma_{y,w}^4)^{-1} \sigma_{\phi 1}^2)$. Following Theorem 1(b), we show further that $\hat{T}_{n0} - T_{n0} = o_p((nh_n^{1/2})^{-1})$ and $\hat{\sigma}_{\phi 1}^2 \xrightarrow{p} \sigma_{\phi 1}^2$, which gives the claimed result in (1).

(2) Under H_1 , $T_{ni} \neq 0$ for $i = a, c, d, e$. We use the additional assumption R5 and follow Theorem 3 to show that $\hat{\eta}_1^2 = (\sigma_{y,w}^2)^{-1}(\sigma_{y,w}^2 - E(\sigma^2(x)w(x))) = \eta_w^2 > 0$, and $\hat{T}_{n0} = O_p((nh_n)^{-1})$, so $\hat{T}_{1n} = nh_n^{1/2} \eta_w^2 + o_p((nh_n^{1/2}))$ and $P(\hat{T}_{1n} > c_n) \rightarrow 1$ for $c_n = o((nh_n^{1/2}))$.

(3) Since $y_i^* = \bar{y}_w + \epsilon_i^*$, $S_y^{*2} = \frac{1}{n} \sum_{i=1}^n (\epsilon_{i,0}^*)^2 w(x_i) = \frac{1}{n} \sum_{i=1}^n \epsilon_i^{*2} w(x_i) + O_p(n^{-1})$.

We note that $\tilde{m}^*(x_t) - \bar{y}_w = \frac{1}{(n-1)h_n f(x_t)} \sum_{i \neq t} K(\frac{x_i - x_t}{h_n}) \epsilon_i^* (1 + o_p(1))$, then

$$\begin{aligned} \frac{1}{n} \sum_{t=1}^n (\epsilon_{t,b}^*)^2 w(x_t) & = \frac{1}{n} \sum_{t=1}^n (\bar{y}_w + \epsilon_t^* - \tilde{m}^*(x_t))^2 w(x_t) \\ & = \frac{1}{n} \sum_{t=1}^n \epsilon_t^{*2} + \frac{2}{n} \sum_{t=1}^n (\bar{y}_w - \tilde{m}^*(x_t)) \epsilon_t^* w(x_t) + \frac{1}{n} \sum_{t=1}^n (\bar{y}_w - \tilde{m}^*(x_t))^2 \\ & = \frac{1}{n} \sum_{t=1}^n \epsilon_t^{*2} - \frac{2}{n(n-1)} \sum_{t=1}^n \sum_{i \neq t} K(\frac{x_i - x_t}{h_n}) \frac{w(x_t)}{h_n f(x_t)} \epsilon_i^* \epsilon_t^* (1 + o_p(1)) \\ & \quad + \frac{1}{n(n-1)^2} \sum_{t=1}^n \sum_{i \neq t} \sum_{j \neq t} K(\frac{x_i - x_t}{h_n}) K(\frac{x_j - x_t}{h_n}) \frac{w(x_t)}{h_n^2 f^2(x_t)} \epsilon_i^* \epsilon_j^* (1 + o_p(1)). \end{aligned}$$

The structure is similar to that in $\hat{T}_n^*$ and we follow Theorem 5 to obtain the claimed result.

$$\text{For } \hat{\eta}_2^2 = \frac{\frac{1}{n} \sum_{i=1}^n (\tilde{m}(x_i) - \bar{m})^2 w(x_i)}{S_y^2}:$$

$$\begin{aligned} \frac{1}{n} \sum_{i=1}^n (\tilde{m}(x_i) - \bar{m})^2 w(x_i) &= \frac{1}{n} \sum_{i=1}^n (\tilde{m}(x_i) - m(x_i))^2 w(x_i) + \frac{1}{n} \sum_{i=1}^n (m(x_i) - \bar{m})^2 w(x_i) \\ &\quad + \frac{2}{n} \sum_{i=1}^n (\tilde{m}(x_i) - m(x_i))(m(x_i) - \bar{m})^2 w(x_i). \end{aligned}$$

Given the similar structure with $\hat{\eta}_1^2$, we only observe that

$$\begin{aligned} \frac{1}{n} \sum_{i=1}^n (m(x_i) - \bar{m})^2 w(x_i) &= \begin{cases} O_p(n^{-1}) + o_p((nh_n^{1/2})^{-1}) & \text{under } H_0. \\ \int m^2(x_i) f(x_i) w(x_i) dx_i - \mu_{y,w}^2 + o_p(1) & \text{under } H_1. \end{cases} \\ \frac{2}{n} \sum_{i=1}^n (\tilde{m}(x_i) - m(x_i))(m(x_i) - \bar{m})^2 w(x_i) &= \begin{cases} O_p(n^{-1}) + o_p((nh_n^{1/2})^{-1}) & \text{under } H_0. \\ O_p(h_n^{k+1}) + O_p((nh_n^{1/2})^{-1} h_n) + O_p(n^{-1/2}) \\ \quad + 2\mu_{y,w}[O_p(h_n^{k+1}) + O_p((nh_n^{1/2})^{-1}) + O_p(n^{1/2})] & \text{under } H_1. \end{cases} \end{aligned}$$

We follow similar arguments in Theorems 1-4 to obtain the desired results in (1) and (2).

For the claim in (3) for the bootstrap test statistic $\hat{T}_{2n}^*$, note

$$\begin{aligned} \tilde{m}^*(x_t) - \bar{y}_w &= \frac{1}{(n-1)h_n f(x_t)} \sum_{i \neq t} K\left(\frac{x_i - x_t}{h_n}\right) \epsilon_i^* (1 + o_p(1)) \text{ and} \\ \bar{m}^* - \bar{y}_w &= \frac{1}{n(n-1)} \sum_{t=1}^n \sum_{i \neq t} K\left(\frac{x_i - x_t}{h_n}\right) \frac{w(x_t) \epsilon_i^*}{h_n f(x_t)} (1 + o_p(1)). \end{aligned}$$

$$\begin{aligned} &\frac{1}{n} \sum_{i=1}^n (\tilde{m}(x_i) - \bar{m}^*)^2 w(x_i) \\ = &\frac{1}{n} \sum_{i=1}^n (\tilde{m}(x_i) - \bar{y}_w)^2 w(x_i) + \frac{1}{n} \sum_{i=1}^n (\bar{y}_w - \bar{m}^*)^2 w(x_i) + \frac{2}{n} \sum_{i=1}^n (\tilde{m}(x_i) - \bar{y}_w)(\bar{y}_w - \bar{m}^*) w(x_i) \\ = &\left[\frac{1}{n(n-1)^2} \sum_{t=1}^n \sum_{i \neq t} K^2\left(\frac{x_i - x_t}{h_n}\right) \frac{w(x_t) \epsilon_i^{*2}}{h_n^2 f^2(x_t)} + \frac{1}{n(n-1)^2} \sum_{t=1}^n \sum_{i \neq t} \sum_{j \neq t} K\left(\frac{x_i - x_t}{h_n}\right) K\left(\frac{x_j - x_t}{h_n}\right) \frac{w(x_t) \epsilon_i^* \epsilon_j^*}{h_n^2 f^2(x_t)} \right] (1 + o_p(1)) \\ &\quad + o_p((nh_n^{1/2})^{-1}). \end{aligned}$$

We follow Theorem 5 to obtain the claimed result in (3) for $\hat{T}_{2n}^*$.

Sketch of proof of Remark 2 in section 3: Proof.

Since the bootstrap test asymptotically has the null limiting distribution of $\hat{T}_{nG}$, we briefly look at the correction terms in $\hat{T}_{nG}$. Under homoskedasticity and bounded support for X denoted by G ,

$$A_{nG} = \frac{1}{n^3 h_n^{2d}} \sum_{t=1}^n \sum_{i \neq t} K_{it}^2 \frac{\epsilon_i^2}{f^2(X_t)} = \frac{1}{n h_n^d} \sigma^2 \int K^2(\psi) d\psi (1 + o_p(1)),$$

$$A_{1nG} = -\frac{2}{n^2 h_n^d} \sum_{t=1}^n K(0) \frac{\epsilon_t^2}{f(X_t)} = -\frac{2}{n h_n^d} \sigma^2 K(0) \int_G dX (1 + o_p(1)),$$

$\frac{1}{n} \sum_{t=1}^n (y_t - \hat{r}(x_{1t}))^2 \xrightarrow{P} \sigma^2$, and $\hat{V}_{TG} \xrightarrow{P} \int_G dX 2 \int (2K(\psi) - \kappa(\psi))^2 d\psi$, so we could focus on $\hat{R}_{nG}^2$ to construct the test, as the other correction terms are independent of the DGP characteristics under H_0 .

The justification of the bootstrap procedure is based on the facts that when we use resample with replacement from the centered $\{\hat{\epsilon}_i\}_{i=1}^n$ to construct $\{\epsilon_i^*\}_{i=1}^n$, we have

$$E(\epsilon_i^* | W) = \frac{1}{n} \sum_{i=1}^n (\hat{\epsilon}_i - \bar{\epsilon}) = 0 \text{ for } \bar{\epsilon} = \frac{1}{n} \sum_{i=1}^n \hat{\epsilon}_i \text{ and } E(\epsilon_i^{*2} | W) = \frac{1}{n} \sum_{i=1}^n (\hat{\epsilon}_i - \bar{\epsilon})^2 = \frac{1}{n} \sum_{i=1}^n \hat{\epsilon}_i^2 - \bar{\epsilon}^2.$$

Since $\bar{\epsilon} = \frac{1}{n} \sum_{i=1}^n (y_i - \hat{m}(x_i)) = \frac{1}{n} \sum_{i=1}^n \epsilon_i - \frac{1}{n} \sum_{i=1}^n (\hat{m}(x_i) - m(x_i)) = O_p(n^{1/2}) + O_p(h_n^v)$, and we can show that $\frac{1}{n} \sum_{i=1}^n \hat{\epsilon}_i^2 = \frac{1}{n} \sum_{i=1}^n (y_i - \hat{m}(x_i))^2 = \frac{1}{n} \sum_{i=1}^n \epsilon_i^2 + o_p(n^{-1/2})$, we obtain $E(\epsilon_i^{*2} | W) = \sigma^2 + O_p(n^{-1/2})$. These observations enable us to follow the proof in Theorem 6 to obtain the claimed result.

References

- de Jong, P., 1987. A central limit theorem for generalized quadratic forms. *Probability Theory and Related Fields* 75, 261–277.
- Doksum, K., Samarov, A., 1995. Nonparametric estimation of global functionals and a measure of the explanatory power of covariates in regression. *Annals of Statistics* 23, 1443–1473.
- Hall, P., 1984. Central Limit Theorem for Integrated Square Error of Multivariate Nonparametric Density Estimators. *Journal of Multivariate Analysis* 14, 1–16.
- Lee, A. J., 1990. *U-statistics Theory and Practice*. Marcel Dekker, New York.

- Martins-Filho, C., Yao, F., 2007. Nonparametric frontier estimation via local linear regression. *Journal of Econometrics* 141, 283–319.
- Newey, W., McFadden, D., 1994. Large sample estimation and hypothesis testing. In: Engle, R. F., McFadden, D. L. (Eds.), *Handbook of Econometrics*. Vol. 4. Elsevier Science B.V., Amsterdam.
- Yao, F., Zhang, J. S., 2010. Efficient semiparametric instrumental variable estimation. Working paper, Economics Department, West Virginia University.
URL <http://community.wvu.edu/~fy006/hp/yao-zhang2011.pdf>