Probability and Expectation

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- Preliminaries
 - Sample Space and Events
 - Defining Probabilities on Events

Motivation

- Preliminaries
 - Sample Space and Events
 - Defining Probabilities on Events

Conditional Probability

- Preliminaries
 - Sample Space and Events
 - Defining Probabilities on Events
- Conditional Probability
- Independent Events

Motivation

- Preliminaries
 - Sample Space and Events
 - Defining Probabilities on Events
- Conditional Probability
- Independent Events

5 Bayes' Formula

- Preliminaries
 - Sample Space and Events
 - Defining Probabilities on Events
- Conditional Probability
- Independent Events
- Bayes' Formula
- 6 Random Variables

- Preliminaries
 - Sample Space and Events
 - Defining Probabilities on Events
- Conditional Probability
- Independent Events
- Bayes' Formula
- Random Variables
- Expectation

- Preliminaries
 - Sample Space and Events
 - Defining Probabilities on Events
- Conditional Probability
- Independent Events
- 5 Bayes' Formula
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- Expectation of a function of a random variable

- Preliminaries
 - Sample Space and Events
 - Defining Probabilities on Events
- Conditional Probability
- Independent Events
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- Preliminaries
 - Sample Space and Events
 - Defining Probabilities on Events
- Conditional Probability
- Independent Events
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- Expectation through conditioning

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Why study probability?

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Note

Discrete Probability combines aspects of logic, set theory, combinatorics and inference.



Preliminaries

Sample Space and Events

- Defining Probabilities on Events

Probability Theory

Preliminaries

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Probability Theory Preliminaries

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- (ii) Suppose that the experiment consists of tossing a die.

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- (i) Suppose that the experiment consists of tossing a coin. Then, $S = \{H, T\}$.
- (ii) Suppose that the experiment consists of tossing a die.

Then, $S = \{1, 2, 3, 4, 5, 6\}$.

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(iii) Suppose that the experiment consists of tossing two coins.

Then, $S = \{HH, HT, TH, TT\}$.

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Then, $S = [0, \infty)$.

Definition

Any subset of the sample space S is called an event.

Probability Theory Preliminaries Sample Space and Events

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Sample Space and Events

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Probability Theory Prolimination Sample Space and Events Combining Events

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Sample Space and Events

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Note

Remember that events are sets when reasoning about them.

Sample Space and Events

Mutually Exclusive events

Probability Theory

Preliminaries

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Defining Probabilities on Events

Outline

Motivation

- 2
 - Preliminaries
 - Sample Space and Events
 - Defining Probabilities on Events
- Conditional Probability
- Independent Events
- 6 Bayes' Formula
- 6 Random Variables
- Expectation
- Expectation of a function of a random variable
- Linearity of Expectation
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Probability Axioms

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P(E) is called the probability of event E.

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The above three conditions are called the axioms of probability theory.

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In general, we assign a probability distribution to the outcomes in *S*, i.e., outcome x_i has probability $p(x_i)$.

In this case,

$$\mathsf{P}(\mathsf{E}) = \sum_{x_i \in \mathsf{E}} \mathsf{p}(x_i).$$



Defining Probabilities on Events



Example

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In the die tossing experiment, what is the probability of the event $\{2, 4, 6\}$?

Probability Theory Preliminaries

Two Identities

Probability Theory Preliminaries Defining Probabilities on Events

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Consider the experiment of tossing two coins.

Assume that all 4 outcomes are equally likely.

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Exercise

Consider the experiment of tossing two coins.

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Let E denote the event that the first coin turns up heads and F denote the event that the second coin turns up heads.

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Then, $P(E) + P(E^c) = 1$.

(ii) Let E and F denote two arbitrary events on the sample space S. Then, P(E ∪ F) = P(E) + P(F) - P(EF). What is P(E ∪ F), when E and F are mutually exclusive? Let G be another event on S. What is P(E ∪ F ∪ G)?

Exercise

Consider the experiment of tossing two coins.

Assume that all 4 outcomes are equally likely.

Let E denote the event that the first coin turns up heads and F denote the event that the second coin turns up heads.

What is the probability that either the first or the second coin turns up heads?

Conditional Probability

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Consider the experiment of tossing two fair coins.

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Consider the experiment of tossing two fair coins.

What is the probability that both coins turn up heads?

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Let *E* and *F* denote two events on a sample space *S*.

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Let *E* and *F* denote two events on a sample space *S*.

The conditional probability of *E*, given that the event *F* has occurred is denoted by $P(E \mid F)$.

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Let *E* and *F* denote two events on a sample space *S*.

The conditional probability of *E*, given that the event *F* has occurred is denoted by P(E | F). P(E | F) is equal to $\frac{P(EF)}{P(E)}$, assuming P(F) > 0.





Example



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Observe that P(F) =



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Hence, $P(E | F) = \frac{\frac{1}{4}}{\frac{1}{2}} = \frac{1}{2}$.

Notice that $P(E) = \frac{1}{4} \neq P(E \mid F)$.

Some more examples

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Assume that the sample space is $S = \{(b, g), (b, b), (g, b), (g, g)\}$

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A family has two children.

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Assume that the sample space is $S = \{(b, g), (b, b), (g, b), (g, g)\}$ and that all outcomes are equally likely.







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Now, $P(F | E) = \frac{P(EF)}{P(E)}$, and hence, $P(EF) = P(F | E) \cdot P(E) = \frac{6}{11} \cdot \frac{7}{12} = \frac{42}{132}$.

One more example

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$$P(Y \mid X) = \frac{P(YX)}{P(X)} = \frac{\frac{8}{100}}{\frac{1}{100}} \approx 0.47$$

Probability Theory Independent Events

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Alternatively,

 $P(EF) = P(E) \cdot P(F)$

Probability Theory Independent Events



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Exercise



Consider the experiment of tossing two fair dice.



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Let F denote the event that the first die turns up 4.



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Consider the experiment of tossing two fair dice.

Let F denote the event that the first die turns up 4.

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Are E₁ and F independent?



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Let E_1 denote the event that the sum of the faces of the two dice is 6.

Let E_2 denote the event that the sum of the faces of the two dice is 7.

Are E₁ and F independent?

How about E_2 and F?

Derivation

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Clearly, $E = EF \cup EF^c$, where the events *EF* and *EF^c* are mutually exclusive.

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Derivation

Let *E* and *F* denote two arbitrary events on a sample space *S*. Clearly, $E = EF \cup EF^c$, where the events *EF* and *EF^c* are mutually exclusive. Now, observe that.

$$P(E) = P(EF) + P(EF^{c}) = P(E | F)P(F) + P(E | F^{c})P(F^{c}) = P(E | F)P(F) + P(E | F^{c})(1 - P(F))$$

Thus, the probability of an event E occurring can be computed as the weighted average of its conditional probability on an arbitrary event F.





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Given that the ball drawn was white, what is the conditional probability that it was drawn from Urn 1?

Solution

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Let W denote the event that a white ball was drawn and let H denote the event that the coin turned up heads.

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Note that H is precisely the event that the ball was drawn from Urn 1.

Accordingly, we have, $P(H) = \frac{1}{2}$ and $P(W | H) = \frac{2}{9}$.

We are interested in the quantity P(H | W).

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Now, $P(HW) = P(W \mid H) \cdot P(H) = \frac{2}{9} \cdot \frac{1}{2} = \frac{1}{9}$. As per Bayes' formula,

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Now, $P(HW) = P(W | H) \cdot P(H) = \frac{2}{9} \cdot \frac{1}{2} = \frac{1}{9}$. As per Bayes' formula,

$$P(W) = P(W \mid H) \cdot P(H) + P(W \mid H^{c})(1 - P(H))$$

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Now, $P(HW) = P(W \mid H) \cdot P(H) = \frac{2}{9} \cdot \frac{1}{2} = \frac{1}{9}$. As per Bayes' formula,

$$P(W) = P(W | H) \cdot P(H) + P(W | H^{c})(1 - P(H))$$

= $\frac{2}{9} \cdot \frac{1}{2} + \frac{5}{11} \cdot \frac{1}{2}$
= $\frac{67}{198}$.

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From conditional probability, we know that, $P(H | W) = \frac{P(HW)}{P(W)}$.

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= $\frac{2}{9} \cdot \frac{1}{2} + \frac{5}{11} \cdot \frac{1}{2}$
= $\frac{67}{198}$.

Therefore, $P(H \mid W) = \frac{\frac{1}{9}}{\frac{67}{100}}$

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= $\frac{67}{198}$.
Therefore, $P(H | W) = \frac{\frac{1}{9}}{\frac{67}{198}} = \frac{22}{67}$,

Solution

From conditional probability, we know that, $P(H | W) = \frac{P(HW)}{P(W)}$.

Now, $P(HW) = P(W \mid H) \cdot P(H) = \frac{2}{9} \cdot \frac{1}{2} = \frac{1}{9}$. As per Bayes' formula,

$$P(W) = P(W | H) \cdot P(H) + P(W | H^{c})(1 - P(H))$$

= $\frac{2}{9} \cdot \frac{1}{2} + \frac{5}{11} \cdot \frac{1}{2}$
= $\frac{67}{198}$.

Therefore, $P(H | W) = \frac{\frac{1}{67}}{\frac{67}{198}} = \frac{22}{67}$, i.e., the conditional probability that the ball was drawn from Urn 1, given that it is white, is $\frac{22}{67}$.

Random Variables

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We may not care whether the actual outcome is (1, 6), (6, 1),or

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In other words, a random variable is a function $X : S \rightarrow E$, where S is the sample space and E is a measurable space.

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$$P\{X = 1\} = 0$$

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$$\vdots$$

$$P\{X = 12\} = \frac{1}{36}$$

Probability Theory Random Variables







Consider the experiment of tossing two fair coins.



Consider the experiment of tossing two fair coins.

Let *Y* denote the random variable that counts the number of heads.



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Discrete Random Variable

Discrete Random Variable

Definition

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A random variable that can take on only a countable number of possible values is said to be *discrete*.

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For a discrete random variable X, the probability mass function (pmf) p(a) is defined as:

$$p(a)=P\{X=a\}.$$

The Bernoulli Random Variable

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Main idea

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If we let the random variable X assume the value 1, if the experiment was a success and 0, if the experiment was a failure, then X is said to be a Bernoulli random variable.

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The probability mass function of *X* is given by:

 $p(1) = P\{X = 1\} = p.$ $p(0) = P\{X = 0\} = 1 - p.$

The Binomial Random Variable

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Consider an experiment which consists of n independent Bernoulli trials, with the probability of success in each trial being p.

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The probability mass function of *X* is given by:

$$p(i) = P\{X = i\} = C(n, i) \cdot p^i \cdot (1 - p)^{n-i}, i = 0, 1, 2, \dots n.$$

Probability Theory Random Variables







Consider the experiment of tossing four fair coins.



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What is the probability that you will get two heads and two tails?



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Solution

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Let the event of heads turning up denote a "success."

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$$p(2) = C(4,2) \cdot (\frac{1}{2})^2 \cdot (1-\frac{1}{2})^2$$
$$= \frac{3}{8}.$$

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The probability mass function of *X* is given by:

$$p(i) = P\{X = i\} = (1 - p)^{i-1} \cdot p, i = 1, 2, \dots$$

Expectation

Definition

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Let X denote a discrete random variable with probability mass function p(x).

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The expected value of X, denoted by E[X] is defined by:

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Let X denote the random variable that records the outcome of tossing a fair die. What is E[X]?

Expectation of a Bernoulli Random Variable

Expectation of a Bernoulli Random Variable

Bernoulli Random Variable

Expectation of a Bernoulli Random Variable

Bernoulli Random Variable

Let X denote a Bernoulli Random Variable with p denoting the probability of success.

Expectation of a Bernoulli Random Variable

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Solution

E[X] =

Expectation of a Bernoulli Random Variable

Bernoulli Random Variable

Let X denote a Bernoulli Random Variable with p denoting the probability of success. Compute E[X].

$$E[X] = 1 \cdot p$$

Expectation of a Bernoulli Random Variable

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Let X denote a Bernoulli Random Variable with p denoting the probability of success. Compute E[X].

$$E[X] = 1 \cdot p + 0 \cdot (1 - p)$$

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Let X denote a Bernoulli Random Variable with p denoting the probability of success. Compute E[X].

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= p .

Expectation of a Binomial Random Variable

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Let *X* denote a Binomial Random Variable with parameters *n* and *p*.

Expectation of a Binomial Random Variable

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Let X denote a Binomial Random Variable with parameters n and p. Compute E[X].

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Expectation of a Binomial Random Variable

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Let X denote a Binomial Random Variable with parameters n and p. Compute E[X].

$$E[X] = \sum_{i=0}^{n} i \cdot p(i)$$
, by definition

Expectation of a Binomial Random Variable

Binomial Random Variable

Ε

Let X denote a Binomial Random Variable with parameters n and p. Compute E[X].

$$[X] = \sum_{i=0}^{n} i \cdot p(i), \text{ by definition}$$
$$= \sum_{i=0}^{n} i \cdot C(n, i) \cdot p^{i} \cdot (1-p)^{n-1}$$

Expectation of a Binomial Random Variable

Binomial Random Variable

F

Let X denote a Binomial Random Variable with parameters n and p. Compute E[X].

$$F[X] = \sum_{i=0}^{n} i \cdot p(i), \text{ by definition}$$
$$= \sum_{i=0}^{n} i \cdot C(n, i) \cdot p^{i} \cdot (1-p)^{n-i}$$
$$= \sum_{i=0}^{n} i \cdot \frac{n!}{i! \cdot (n-i)!} \cdot p^{i} \cdot (1-p)^{n-i}$$

Expectation of a Binomial Random Variable

Binomial Random Variable

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Let X denote a Binomial Random Variable with parameters n and p. Compute E[X].

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Expectation of a Binomial Random Variable (contd.)

$$E[X] = \sum_{i=1}^{n} i \cdot \frac{n!}{i! \cdot (n-i)!} \cdot p^{i} \cdot (1-p)^{n-i}$$

Expectation of a Binomial Random Variable (contd.)

$$E[X] = \sum_{i=1}^{n} i \cdot \frac{n!}{i! \cdot (n-i)!} \cdot p^{i} \cdot (1-p)^{n-i}$$
$$=$$

Expectation of a Binomial Random Variable (contd.)

$$E[X] = \sum_{i=1}^{n} i \cdot \frac{n!}{i! \cdot (n-i)!} \cdot p^{i} \cdot (1-p)^{n-i}$$
$$= \sum_{i=1}^{n} \frac{n!}{(i-1)! \cdot (n-i)!} \cdot p^{i} \cdot (1-p)^{n-i}$$

Expectation of a Binomial Random Variable (contd.)

$$E[X] = \sum_{i=1}^{n} i \cdot \frac{n!}{i! \cdot (n-i)!} \cdot p^{i} \cdot (1-p)^{n-i}$$

=
$$\sum_{i=1}^{n} \frac{n!}{(i-1)! \cdot (n-i)!} \cdot p^{i} \cdot (1-p)^{n-i}$$

=
$$n \cdot p \cdot \sum_{i=1}^{n} \frac{(n-1)!}{(i-1)! \cdot (n-i)!} \cdot p^{i-1} \cdot (1-p)^{n-i}$$

Expectation of a Binomial Random Variable (contd.)

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Expectation of a Binomial Random Variable (contd.)

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Expectation of a Binomial Random Variable (contd.)

Solution

Substituting k = (i - 1), we get,

E[X] =

Expectation of a Binomial Random Variable (contd.)

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$$E[X] = n \cdot p \cdot \sum_{k=0}^{n-1} \frac{(n-1)!}{k! \cdot (n-k-1)!} \cdot p^k \cdot (1-p)^{n-k-1}$$

Expectation of a Binomial Random Variable (contd.)

Solution

Substituting k = (i - 1), we get,

$$E[X] = n \cdot p \cdot \sum_{k=0}^{n-1} \frac{(n-1)!}{k! \cdot (n-k-1)!} \cdot p^k \cdot (1-p)^{n-k-1}$$
$$= n \cdot p \cdot \sum_{k=0}^{n-1} \frac{(n-1)!}{k! \cdot ((n-1)-k)!} \cdot p^k \cdot (1-p)^{(n-1)-k}$$

Expectation of a Binomial Random Variable (contd.)

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=

$$E[X] = n \cdot p \cdot \sum_{k=0}^{n-1} \frac{(n-1)!}{k! \cdot (n-k-1)!} \cdot p^k \cdot (1-p)^{n-k-1}$$

= $n \cdot p \cdot \sum_{k=0}^{n-1} \frac{(n-1)!}{k! \cdot ((n-1)-k)!} \cdot p^k \cdot (1-p)^{(n-1)-k}$

Expectation of a Binomial Random Variable (contd.)

Solution

$$E[X] = n \cdot p \cdot \sum_{k=0}^{n-1} \frac{(n-1)!}{k! \cdot (n-k-1)!} \cdot p^k \cdot (1-p)^{n-k-1}$$

= $n \cdot p \cdot \sum_{k=0}^{n-1} \frac{(n-1)!}{k! \cdot ((n-1)-k)!} \cdot p^k \cdot (1-p)^{(n-1)-k}$
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Expectation of a Binomial Random Variable (contd.)

Solution

$$E[X] = n \cdot p \cdot \sum_{k=0}^{n-1} \frac{(n-1)!}{k! \cdot (n-k-1)!} \cdot p^k \cdot (1-p)^{n-k-1}$$

= $n \cdot p \cdot \sum_{k=0}^{n-1} \frac{(n-1)!}{k! \cdot ((n-1)-k)!} \cdot p^k \cdot (1-p)^{(n-1)-k}$
= $n \cdot p \cdot \sum_{k=0}^{n-1} C(n-1,k) \cdot p^k \cdot (1-p)^{(n-1)-k}$
= $n \cdot p \cdot [p+(1-p)]^{n-1}$, recall the Binomial theorem

Expectation of a Binomial Random Variable (contd.)

Solution

$$F[X] = n \cdot p \cdot \sum_{k=0}^{n-1} \frac{(n-1)!}{k! \cdot (n-k-1)!} \cdot p^k \cdot (1-p)^{n-k-1}$$

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= $n \cdot p \cdot [p+(1-p)]^{n-1}$, recall the Binomial theorem
= $n \cdot p \cdot 1$

Expectation of a Binomial Random Variable (contd.)

Solution

$$E[X] = n \cdot p \cdot \sum_{k=0}^{n-1} \frac{(n-1)!}{k! \cdot (n-k-1)!} \cdot p^k \cdot (1-p)^{n-k-1}$$

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$$E[X] = \sum_{i=1}^{\infty} i \cdot p(i)$$
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Expectation of a Geometric Random Variable

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Expectation of a Geometric Random Variable

Geometric Random Variable

Е

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$$= \sum_{i=1}^{\infty} i \cdot q^{i-1} \cdot p, \text{ where } q = 1 - 1$$

р

Expectation of a Geometric Random Variable

Geometric Random Variable

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Expectation of a Geometric Random Variable (contd.)

Expectation of a Geometric Random Variable (contd.)

Expectation of a Geometric Random Variable (contd.)

$$E[X] = p \cdot \sum_{i=1}^{\infty} i \cdot q^{i-1}$$

Expectation of a Geometric Random Variable (contd.)

Solution

$$E[X] = p \cdot \sum_{i=1}^{\infty} i \cdot q^{i-1}$$

=

Expectation of a Geometric Random Variable (contd.)

$$E[X] = p \cdot \sum_{i=1}^{\infty} i \cdot q^{i-1}$$
$$= p \cdot \sum_{i=1}^{\infty} \frac{d}{dq}[q^{i}]$$

Expectation of a Geometric Random Variable (contd.)

Solution

$$\mathbf{E}[X] = \mathbf{p} \cdot \sum_{i=1}^{\infty} i \cdot q^{i-1}$$
$$= \mathbf{p} \cdot \sum_{i=1}^{\infty} \frac{d}{dq} [q^i]$$

=

Expectation of a Geometric Random Variable (contd.)

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E

Expectation of a Geometric Random Variable (contd.)

Solution

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$$= p \cdot \frac{d}{dq} [\sum_{i=1}^{\infty} q^i]$$
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Е

Expectation of a Geometric Random Variable (contd.)

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$$= p \cdot \frac{d}{dq} [\sum_{i=1}^{\infty} q^i]$$

$$= p \cdot \frac{d}{dq} [\frac{q}{1-q}]$$

$$= p \cdot \frac{(1-q) \cdot \frac{d}{dq} [q] - q \cdot \frac{d}{dq} [1-q]}{(1-q)^2}$$

q]

Expectation of a Geometric Random Variable (contd.)

Solution

$$E[X] = p \cdot \sum_{i=1}^{\infty} i \cdot q^{i-1}$$

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$$= p \cdot \frac{(1-q) \cdot \frac{d}{dq} [q] - q \cdot \frac{d}{dq} [1-q]}{(1-q)^2}$$

$$= p \cdot \frac{(1-q) \cdot 1 - q \cdot (-1)}{(1-q)^2}$$

q]

Expectation of a Geometric Random Variable (contd.)

Expectation of a Geometric Random Variable (contd.)

Expectation of a Geometric Random Variable (contd.)



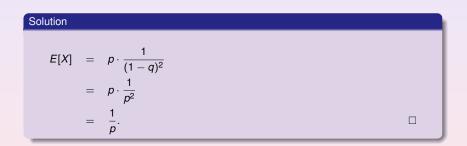
Expectation of a Geometric Random Variable (contd.)

$$E[X] = p \cdot \frac{1}{(1-q)^2}$$

Expectation of a Geometric Random Variable (contd.)

$$E[X] = p \cdot \frac{1}{(1-q)^2}$$
$$= p \cdot \frac{1}{p^2}$$

Expectation of a Geometric Random Variable (contd.)



Probability Theory Expectation of a function of a random variable

Motivation

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Often times, we are interested in a function of the random variable, rather than the random variable itself.

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Compute $E[X^2]$.

Expectation of functions of random variables (contd.)

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Expectation of functions - The Direct Approach

Probability Theory Expectation of a function of a random variable

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Note that linearity of expectation holds even when the random variables are **not** independent.

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Example

What is the expected value of the sum of the upturned faces, when two fair dice are tossed?

Another Application

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Compute the expected value of the Binomial random variable.

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$$X = X_1 + X_2 + \ldots X_n$$

Example

Compute the expected value of the Binomial random variable.

Solution

Define

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$$\begin{array}{rcl} X & = & X_1 + X_2 + \dots X_n \\ \Rightarrow E[X] & = \end{array}$$

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Accordingly, the Binomial random variable (say X) can be expressed as:

=

$$X = X_1 + X_2 + \dots X_n$$

$$\Rightarrow E[X] = E[X_1 + X_2 + \dots X_n]$$

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$$= E[\sum_{i=1}^n X_i]$$

Expectation of the binomial random variable

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= $\sum_{i=1}^{n} p$, since each X_i is a Bermoulli random variable,
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= $n \cdot p$.

Probability Theory

Expectation through conditioning

The Basics

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Expectations by conditioning

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Note that E[X | Y] is itself a random variable.

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If X and Y are discrete, then

$$E[X] = \sum_{y} E[X \mid Y = y] \cdot P\{Y = y\}.$$







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$$\Rightarrow E[X] = E[X | Y = 1] \cdot P\{Y = 1\} + E[X | Y = 0] \cdot P\{Y = 0\}$$

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(a) $P{Y = 1} = p$.

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(b)
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(c)
$$E[X | Y = 1] = 1.$$

(d) E[X | Y = 0] = (1 + E[X]).

Probability Theory Expectation through conditioning

Expectation of the geometric random variable

Computing the expectation

$$\Rightarrow E[X] = E[X | Y = 1] \cdot P\{Y = 1\} + E[X | Y = 0] \cdot P\{Y = 0\}$$

Observe that,

- (a) $P{Y = 1} = p$.
- (b) $P{Y = 0} = (1 p)$.
- (c) E[X | Y = 1] = 1.
- (d) E[X | Y = 0] = (1 + E[X]).

Probability Theory Expectation through conditioning

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$$E[X] = 1 \cdot p + (1 + E[X]) \cdot (1 - p)$$

Expectation of the geometric random variable

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$$\Rightarrow E[X] =$$

Expectation of the geometric random variable

$$\Rightarrow E[X] = p + (1-p) + (1-p) \cdot E[X]$$

Expectation of the geometric random variable

$$\Rightarrow E[X] = p + (1 - p) + (1 - p) \cdot E[X]$$

$$\Rightarrow p \cdot E[X] = 1$$

Expectation of the geometric random variable

$$\begin{array}{rcl} \Rightarrow E[X] &=& p + (1 - p) + (1 - p) \cdot E[X] \\ \Rightarrow p \cdot E[X] &=& 1 \\ \Rightarrow E[X] &=& \frac{1}{p}. \end{array}$$