Chernoff Bounds (Applications)

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Confidence Interval

O Chernoff Bounds - Coin Flips

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Better Bounds for Special Cases

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Better Bounds for Special Cases

Application: Set Balancing

Chernoff Bounds - Coin Flips

Example

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$$P(|X - \mu| \ge \delta \mu) \le 2e^{-\frac{\mu\delta^2}{3}}$$
 for $\delta = \sqrt{\frac{6\ln n}{n}}$.

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Most of the times, the deviations from the mean are on the order of $O(\sqrt{n \ln n})$.

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Among *n* samples, we find the specific value we are interested in exactly in $X = \overline{p}n$ samples.

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Setting $\gamma = e^{-\frac{n\delta^2}{2}} + e^{-\frac{n\delta^2}{3}}$, we obtain a tradeoff between δ, γ and n.

Better Bounds for Special Cases

Special Case

Consider the case where each random variable takes its value from the set $\{-1,1\}$ with the exact same probability (i.e., $p = \frac{1}{2}$).

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Proof.

For any t > 0,

$$E[\mathbf{e}^{tX_i}] = \frac{1}{2}\mathbf{e}^t + \frac{1}{2}\mathbf{e}^{-t}$$

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$$e^t = 1 + \frac{t}{1!} + \frac{t^2}{2!} + \dots + \frac{t^i}{i!} + \dots$$

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Better Bounds for Special Cases

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$$E[e^{tX_i}] = \frac{1}{2}e^t + \frac{1}{2}e^{-t}$$
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$$\le \sum_{i \ge 0} \frac{(t^2/2)^i}{i!}$$

Better Bounds for Special Cases

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Better Bounds for Special Cases

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Using this estimation, we can evaluate $E[e^{tX}]$:

$$E[e^{tX}] = \prod_{i=1}^{n} E[e^{e^{tX_i}}] \le e^{\frac{t^2n}{2}}$$

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$$P(X \ge a) = P(e^{tX} \ge e^{ta})$$

F

Better Bounds for Special Cases

Proof.

Thus:

$$\begin{aligned} \mathbf{F}[\mathbf{e}^{tX_j}] &= \frac{1}{2}\mathbf{e}^t + \frac{1}{2}\mathbf{e}^- \\ &= \sum_{i\geq 0} \frac{t^{2i}}{(2i)!} \\ &\leq \sum_{i\geq 0} \frac{(t^2/2)^i}{i!} \\ &= \mathbf{e}^{\frac{t^2}{2}} \end{aligned}$$

Using this estimation, we can evaluate $E[e^{tX}]$:

$$E[e^{tX}] = \prod_{i=1}^{n} E[e^{e^{tX_i}}] \le e^{\frac{t^2n}{2}}$$

$$\mathsf{P}(\mathsf{X} \geq \mathsf{a}) = \mathsf{P}(\mathsf{e}^{t\mathsf{X}} \geq \mathsf{e}^{t\mathsf{a}}) \leq rac{\mathsf{E}[\mathsf{e}^{t\mathsf{X}}]}{\mathsf{e}^{t\mathsf{a}}}$$

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Better Bounds for Special Cases

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Thus:

$$\begin{aligned} [e^{tX_i}] &= \frac{1}{2}e^t + \frac{1}{2}e^{-t} \\ &= \sum_{i \ge 0} \frac{t^{2i}}{(2i)!} \\ &\le \sum_{i \ge 0} \frac{(t^2/2)^i}{i!} \\ &= e^{\frac{t^2}{2}} \end{aligned}$$

Using this estimation, we can evaluate $E[e^{tX}]$:

$$E[e^{tX}] = \prod_{i=1}^{n} E[e^{e^{tX_i}}] \le e^{\frac{t^2n}{2}}$$

$$P(X \ge a) = P(e^{tX} \ge e^{ta}) \le rac{E[e^{tX}]}{e^{ta}} \le e^{rac{t^2n}{2} - ta}$$

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Using this estimation, we can evaluate $E[e^{tX}]$:

$$E[e^{tX}] = \prod_{i=1}^{n} E[e^{e^{tX_i}}] \le e^{\frac{t^2n}{2}}$$

and, using Markov's bound (recall that t > 0):

$$P(X \ge a) = P(e^{tX} \ge e^{ta}) \le rac{E[e^{tX}]}{e^{ta}} \le e^{rac{t^2n}{2} - ta}$$

Better Bounds for Special Cases

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Proof. Recall that for any a > 0 we want to prove that

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Note

By symmetry, we also have:

$$P(X \le a) \le e^{-\frac{a^2}{2n}}$$

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Corollary

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$$P(|X| \ge a) \le$$

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Let X_1, \ldots, X_n be independent random variables with

$$P(X_i = 1) = P(X_i = -1) = \frac{1}{2}$$

Let $X = \sum_{i=1}^{n} X_i$. For any a > 0,

$$P(|X| \ge a) \le 2e^{-\frac{a^2}{2n}}$$

Better Bounds for Special Cases

Transformation

We apply the transformation $Y_i = \frac{(X_i+1)}{2}$.

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We apply the transformation $Y_i = \frac{(X_i+1)}{2}$. Why?

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Let Y_1, \ldots, Y_n be independent random variables with

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Let $Y = \sum_{i=1}^{n} Y_i$ and $\mu = E[Y] = \frac{n}{2}$. • For any a > 0, $P(Y \ge \mu + a) \le e^{-\frac{2a^2}{n}}$

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2

Let $Y = \sum_{i=1}^{n} Y_i$ and $\mu = E[Y] = \frac{n}{2}$.

• For any
$$a > 0$$
, $P(Y \ge \mu + a) \le e^{-\frac{2a^2}{n}}$

② For any
$$\delta >$$
 0, P(Y \geq (1 + $\delta)\mu$) \leq e $^{-\delta^{2}\mu}$

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Question

Why is the bound given by point (1) above special?

Better Bounds for Special Cases

Proof.

For point (1):

Better Bounds for Special Cases

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For point (1): Using the fact that $Y_i = (X_i + 1)/2$, we have that

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Better Bounds for Special Cases

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Better Bounds for Special Cases

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Better Bounds for Special Cases

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Better Bounds for Special Cases

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$$P(Y \ge \mu + a) = P(\frac{1}{2}X + \mu \ge \mu + a) = P(X \ge 2a)$$

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$$P(Y \ge \mu + a) = P(\frac{1}{2}X + \mu \ge \mu + a) = P(X \ge 2a)$$

Since, for any a > 0 (by the previous theorem), $P(X \ge a) \le e^{-\frac{a^2}{2n}}$, we have that:

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which completes the proof for point (1).

Better Bounds for Special Cases

Proof.

For point (2):

Better Bounds for Special Cases

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For point (2): We will utilize our previous result:

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Better Bounds for Special Cases

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$$P(Y \ge (1+\delta)\mu) =$$

Better Bounds for Special Cases

Proof.

For point (2): We will utilize our previous result:

$$P(Y \ge \mu + a) \le e^{-rac{2a^2}{n}}$$

$$\mathsf{P}(\mathsf{Y} \ge (\mathsf{1} + \delta)\mu) \quad = \quad \mathsf{P}(\mathsf{Y} \ge \mu + \delta\mu)$$

Better Bounds for Special Cases

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For point (2): We will utilize our previous result:

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$$egin{array}{rl} {\sf P}({\sf Y}\geq(1+\delta)\mu)&=&{\sf P}({\sf Y}\geq\mu+\delta\mu)\ &<&e^{-rac{2(\delta\mu)^2}{n}} \end{array}$$

Better Bounds for Special Cases

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For point (2): We will utilize our previous result:

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$$= e^{-\frac{(\delta\mu)^2}{2}}$$

Better Bounds for Special Cases

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For $a = \delta \mu$ (Note that since $\delta >$ 0, also $\delta \mu >$ 0):

$$P(Y \ge (1+\delta)\mu) = P(Y \ge \mu + \delta\mu)$$
$$\leq e^{-\frac{2(\delta\mu)^2}{n}}$$
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For $a = \delta \mu$ (Note that since $\delta >$ 0, also $\delta \mu >$ 0):

$$\begin{array}{rcl} \mathsf{P}(\mathsf{Y} \ge (\mathsf{1} + \delta)\mu) & = & \mathsf{P}(\mathsf{Y} \ge \mu + \delta\mu) \\ & \leq & \mathsf{e}^{-\frac{2(\delta\mu)^2}{n}} \\ & = & \mathsf{e}^{-\frac{(\delta\mu)^2}{\frac{\mu}{2}}} \\ & = & \mathsf{e}^{-\frac{(\delta\mu)^2}{\mu}} \\ & = & \mathsf{e}^{-\delta^2\mu} \end{array}$$

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Corollary

Let Y_1, \ldots, Y_n be independent random variables with

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 and $\mu = E[Y] = \frac{n}{2}$.
• For any $0 < a < \mu$, $P(Y \ge \mu - a) \le e^{-\frac{2a^2}{n}}$
• For any $0 < \delta < 1$, $P(Y \ge (1 - \delta)\mu) \le e^{-\delta^2\mu}$

Application: Set Balancing

Set balancing

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Set balancing

Consider:

• A *n*×*m* matrix **A** with entries in {0,1}, where *a_{ij}*, (*i* = 1,...*n*, *j* = 1,...*m*) corresponds to the element of the *i*th row and the *j*th column.

Application: Set Balancing

Set balancing

Consider:

- A n×m matrix A with entries in {0,1}, where a_{ij}, (i = 1,...n, j = 1,...m) corresponds to the element of the ith row and the jth column.
- A $m \times 1$ vector **b** with entries in $\{-1, 1\}$, where $b_j, (j = 1, ..., m)$ corresponds to the j^{th} element of **b**.

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- A $n \times 1$ vector **c**, where $c_i, (i = 1, ..., n)$ corresponds to the *i*th element of **c**.

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- A $n \times 1$ vector **c**, where $c_i, (i = 1, ..., n)$ corresponds to the *i*th element of **c**.

Given A, we want to find the entries of vector b that minimize

$$||\mathbf{A} \cdot \mathbf{b}||_{\infty} = \max_{i=1,\dots,n} |c_i|$$

Application: Set Balancing

Set balancing

Consider:

- A n×m matrix A with entries in {0,1}, where a_{ij}, (i = 1,...n, j = 1,...m) corresponds to the element of the ith row and the jth column.
- A m × 1 vector b with entries in {−1,1}, where b_j, (j = 1,...m) corresponds to the jth element of b.
- A $n \times 1$ vector **c**, where c_i , (i = 1, ..., n) corresponds to the i^{th} element of **c**.

Given A, we want to find the entries of vector b that minimize

$$\|\mathbf{A}\cdot\mathbf{b}\|_{\infty} = \max_{i=1,\dots,n} |c_i|$$

Motivation

This problem rises in designing statistical experiments.

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This problem rises in designing statistical experiments. Each column of **A** represents a subject in the experiment and each row a feature. The vector **b** partitions the subjects into two disjoint groups (through multiplying either by 1 or by -1). So we are looking a way to separate the participants into two groups so that each feature is roughly as balanced as possible between the two groups.

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We will show that using this approach it is likely that we obtain a rather tight bound for $||\mathbf{A} \cdot \mathbf{b}||_{\infty}$ (i.e., $O(\sqrt{m \ln n})$).

Application: Set Balancing

Theorem

For a random vector ${\bm b}$ with entries chosen independently and with equal probability from the set $\{-1,1\},$

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Therefore, we can use the previously obtained results, namely that for a > 0:

$$P(|X| \ge a) \le 2e^{-\frac{a^2}{2n}}$$

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Since we are interested for all Z_i , i = 1, ..., n, the probability for the bound $\sqrt{4m \ln n}$ to fail is $\frac{2}{n}$.